

34th International Workshop on Statistical Genetic Methods for Human Complex Traits - 2020

- **Mike Neale** (director)  
- **Sarah Medland**(co)  
- David Evans 
- Ben Neale  
- Hermine Maes  
- Conor Dolan  
- Lucia Colodro Conde  
- Nathan Gillespie  
- Brad Verhulst  
- Tim Bates  
- **John Hewitt** (host)  
- Jeff Lessem 
- Matt Keller 
- Stacey Cherny   
- Elizabeth Prom-Wormley 
- Nick Martin 
- Dorret Boomsma 
- Eveline de Zeeuw 
- Michel Nivard 
- Rob Kirkpatrick 
- Pete Hatemi 
- Mike Hunter 

The Causes of Variation

2020 International Workshop on Statistical
Genetic Methods for Human Complex Traits
Boulder, CO.

Nick Martin
Queensland Institute of Medical Research
Brisbane, Australia
March 2, 2020

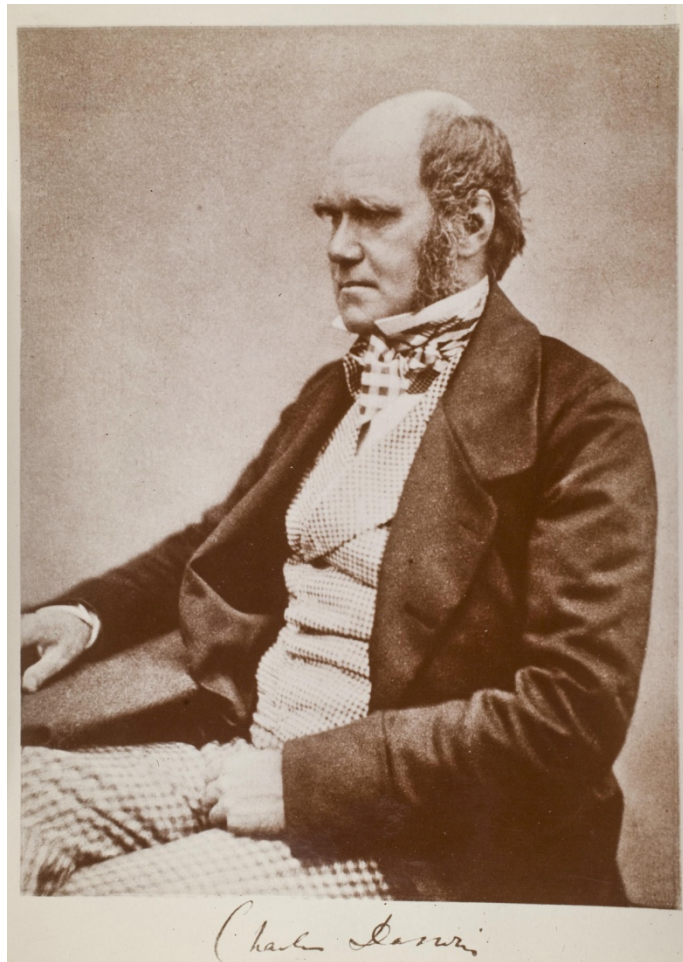
GOALS

- To outline some of the basic issues that we shall address this week
- To set some of the issues in their historical context in the story of genetics and its application to human variation.

Genetics

The Study of
Variation and Heredity

Charles Darwin (1809-1882)



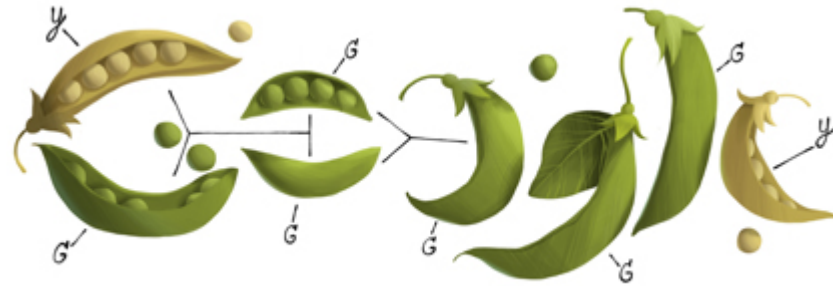
1859: On the Origin of Species

Gregor Mendel (1822-1885)



1865: Experiments in Plant Hybridisation

Mendel 1865 – genetics of discrete traits



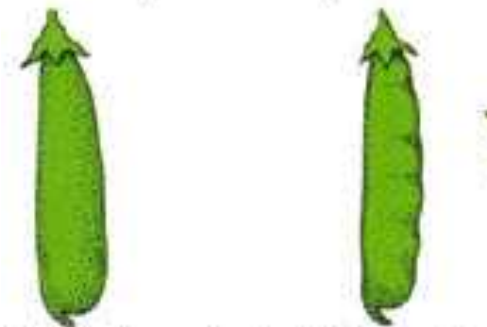
Round or wrinkled ripe seeds



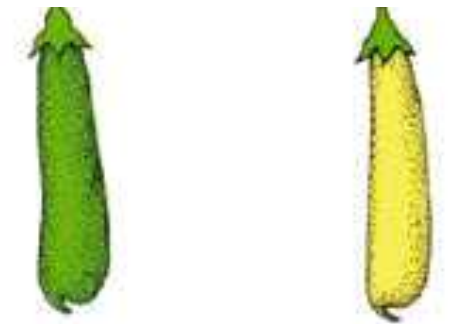
Yellow or green seed interiors



Purple or white petals



Inflated or pinched ripe pods



Green or yellow unripe pods

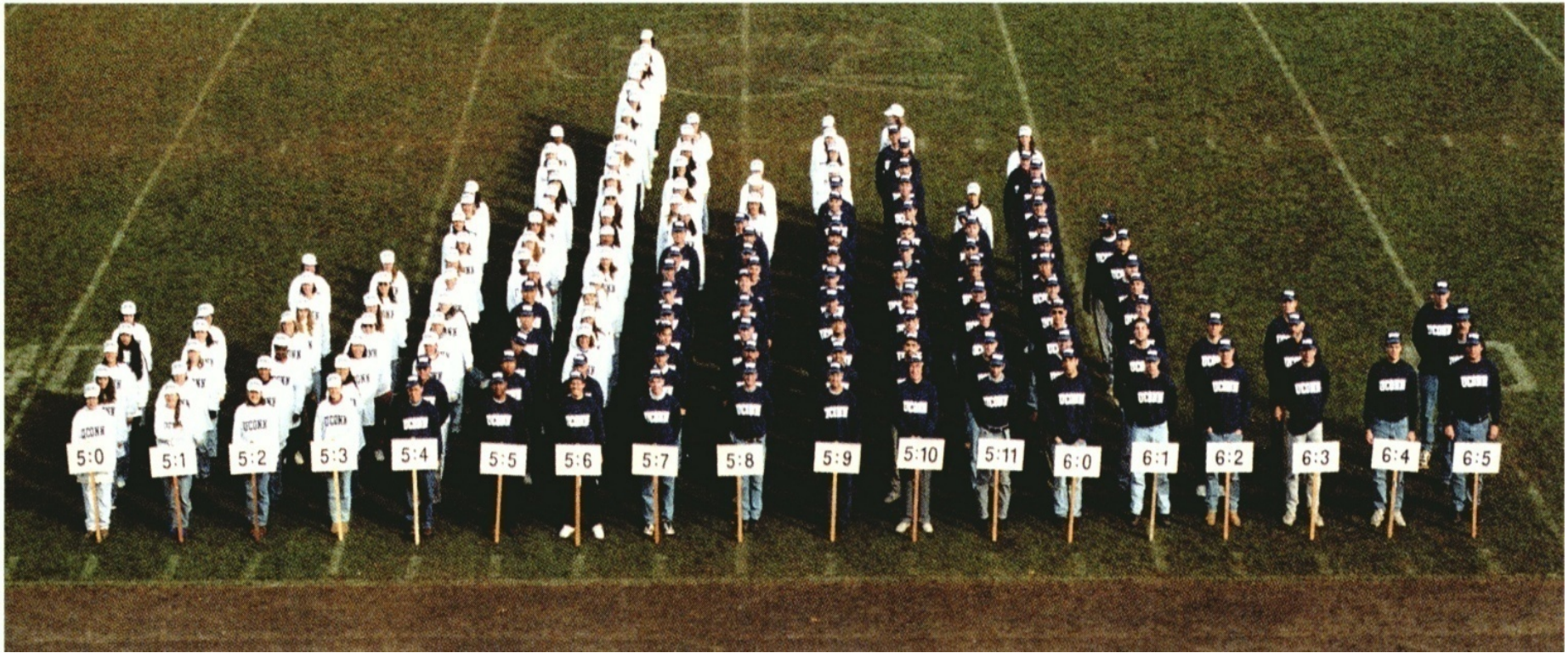


Axial or terminal flowers



Long or short stems

....the distribution of height



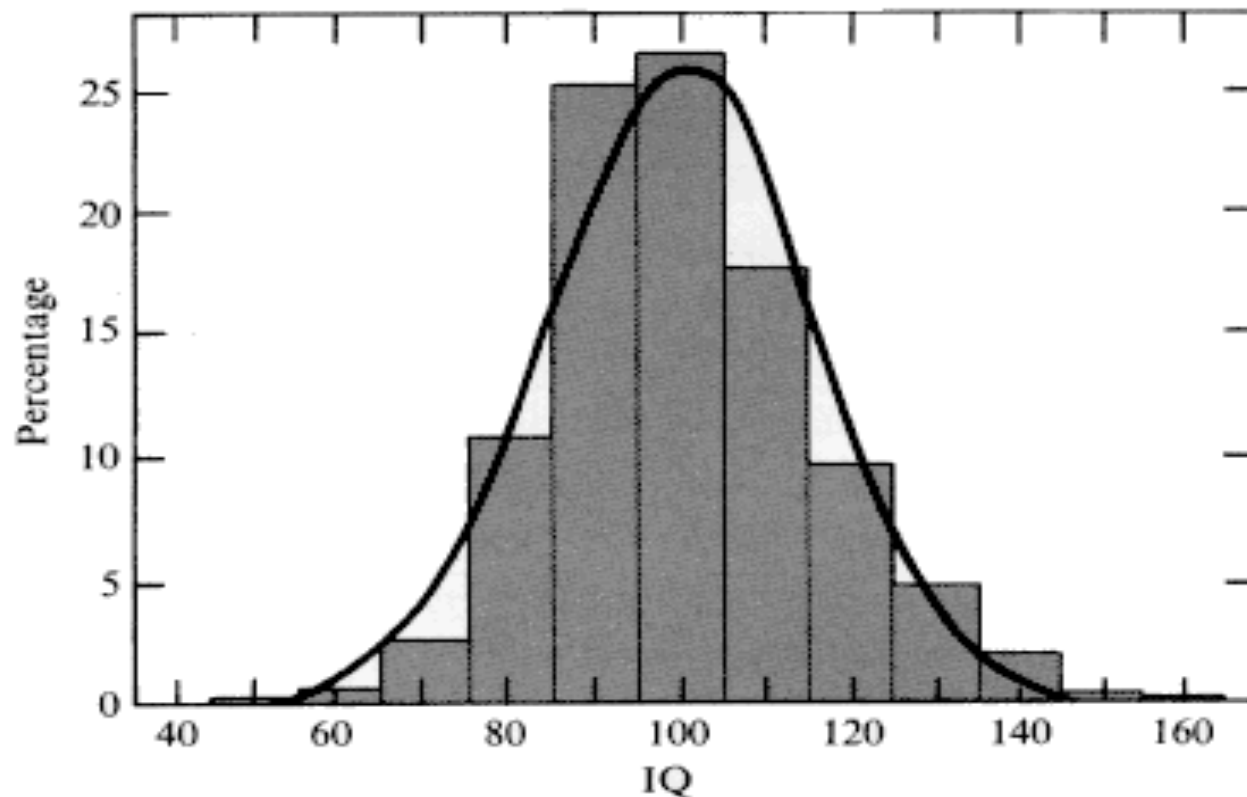
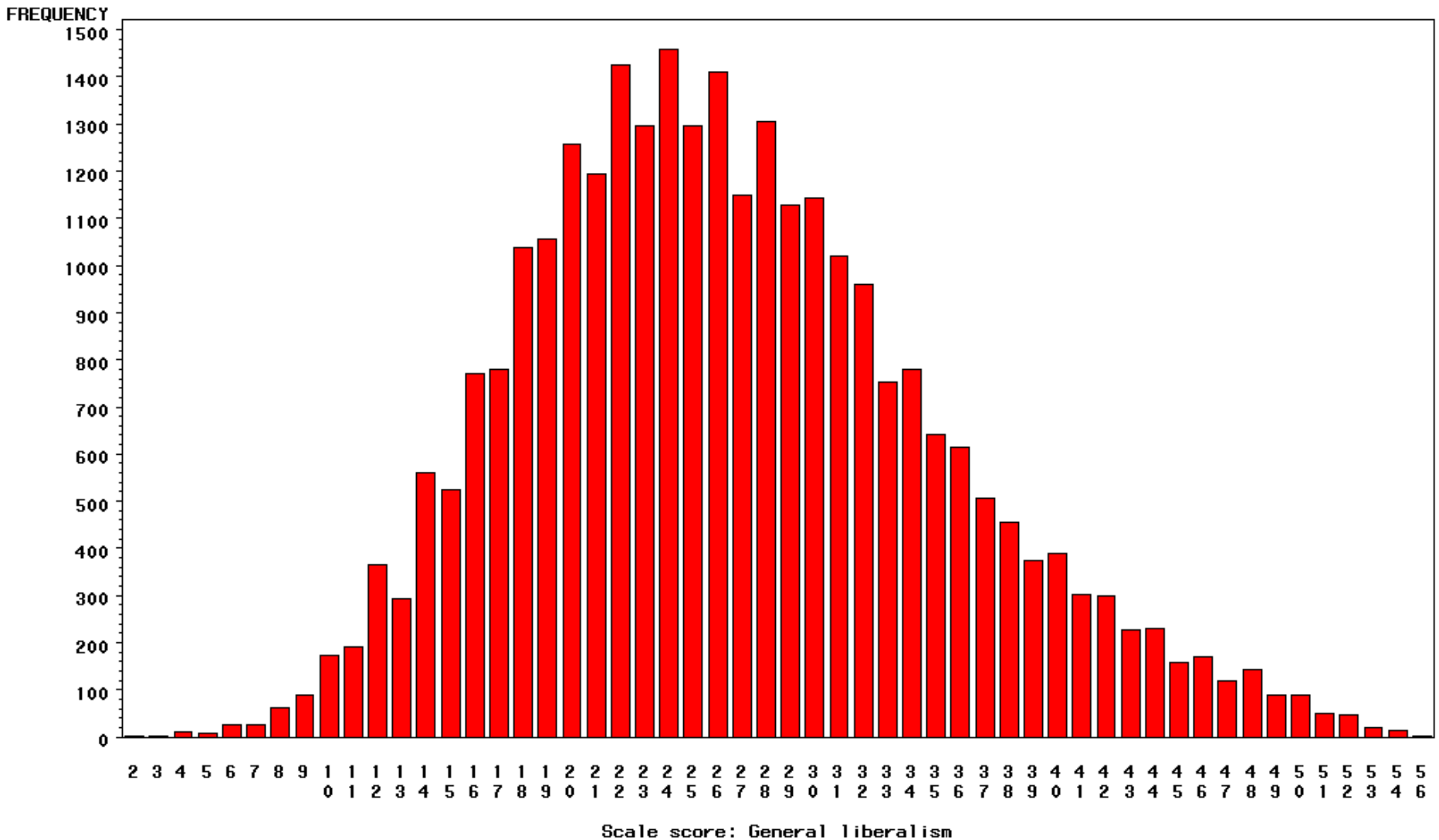


FIGURE 9.4

The distribution of IQ among the 14,963 children born in Scotland on February 1, May 1, August 1, and November 1, 1926. The shaded histogram shows the percentages of the group with IQ's in various ranges of 10 points. This grouping is artificial and is done solely for ease of representation: it does not imply any discontinuity in the values of IQ that children can show. The continuous curve shows the ideal distribution calculated from the observations and representing the statistical population of which the children actually observed are regarded as forming a sample. (Data from MacMeekan; from Mather 1964.)

“Liberalism”



XV.—The Correlation between Relatives on the Supposition of Mendelian Inheritance. By R. A. Fisher, B.A. Communicated by Professor J. ARTHUR THOMSON. (With Four Figures in Text.)

(MS. received June 15, 1918. Read July 8, 1918. Issued separately October 1, 1918.)

CONTENTS.

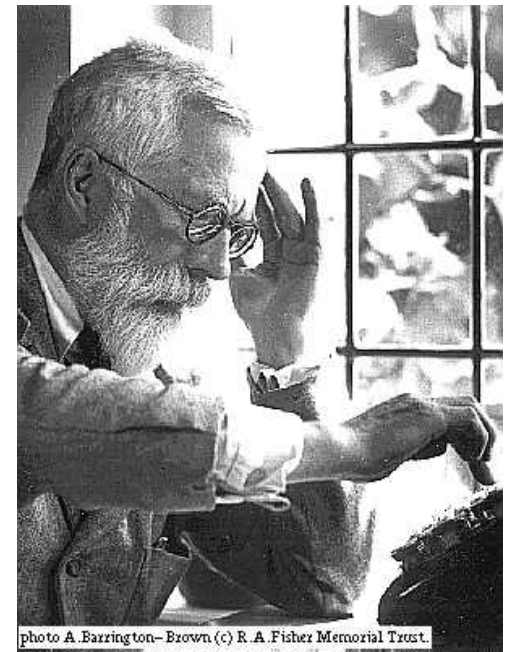
	PAGE		PAGE
1. The superposition of factors distributed independently	402	15. Homogamy and multiple allelomorphism	416
2. Phase frequency in each array	402	16. Coupling	418
3. Parental regression	403	17. Theories of marital correlation; ancestral correlations	419
4. Dominance deviations	403	18. Ancestral correlations (second and third theories)	421
5. Correlation for parent; genetic correlations	404	19. Numerical values of association	421
6. Fraternal correlation	405	20. Fraternal correlation	422
7. Correlations for other relatives	406	21. Numerical values for environment and dominance ratios; analysis of variance	423
8. Epistacy	408	22. Other relatives	424
9. Assortative mating	410	23. Numerical values (third theory)	425
10. Frequency of phases	410	24. Comparison of results	427
11. Association of factors	411	25. Interpretation of dominance ratio (diagrams)	428
12. Conditions of equilibrium	412	26. Summary	432
13. Nature of association	413		
14. Multiple allelomorphism	415		

Several attempts have already been made to interpret the well-established results of biometry in accordance with the Mendelian scheme of inheritance. It is here attempted to ascertain the biometrical properties of a population of a more general type than has hitherto been examined, inheritance in which follows this scheme. It is hoped that in this way it will be possible to make a more exact analysis of the causes of human variability. The great body of available statistics show us that the deviations of a human measurement from its mean follow very closely the Normal Law of Errors, and, therefore, that the variability may be uniformly measured by the standard deviation corresponding to the square root of the mean square error. When there are two independent causes of variability capable of producing in an otherwise uniform population distributions with standard deviations σ_1 and σ_2 , it is found that the distribution, when both causes act together, has a standard deviation $\sqrt{\sigma_1^2 + \sigma_2^2}$. It is therefore desirable in analysing the causes of variability to deal with the square of the standard deviation as the measure of variability. We shall term this quantity the Variance of the normal population to which it refers, and we may now ascribe to the constituent causes fractions or percentages of the total variance which they together produce. It

RA Fisher (1918).
*Transactions of
the Royal Society
of Edinburgh*
52: 399-433.

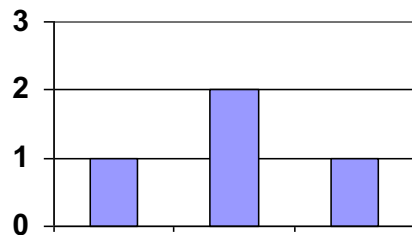
R.A. Fisher, 1918

The explanation of quantitative inheritance in Mendelian terms



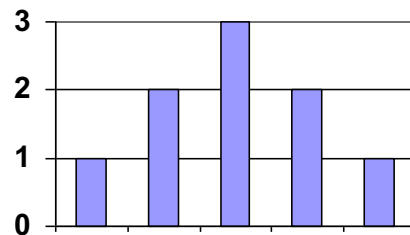
1 Gene

- 3 Genotypes
- 3 Phenotypes



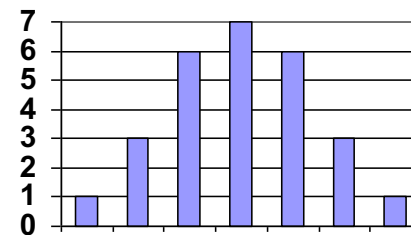
2 Genes

- 9 Genotypes
- 5 Phenotypes



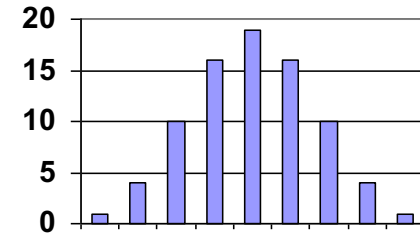
3 Genes

- 27 Genotypes
- 7 Phenotypes



4 Genes

- 81 Genotypes
- 9 Phenotypes

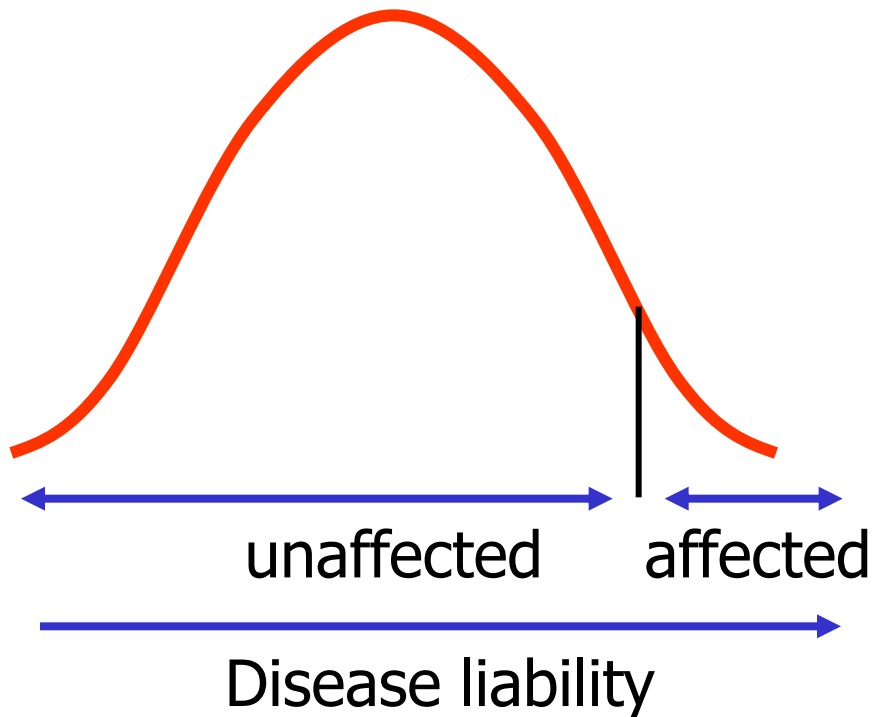


Fisher (1918): Basic Ideas

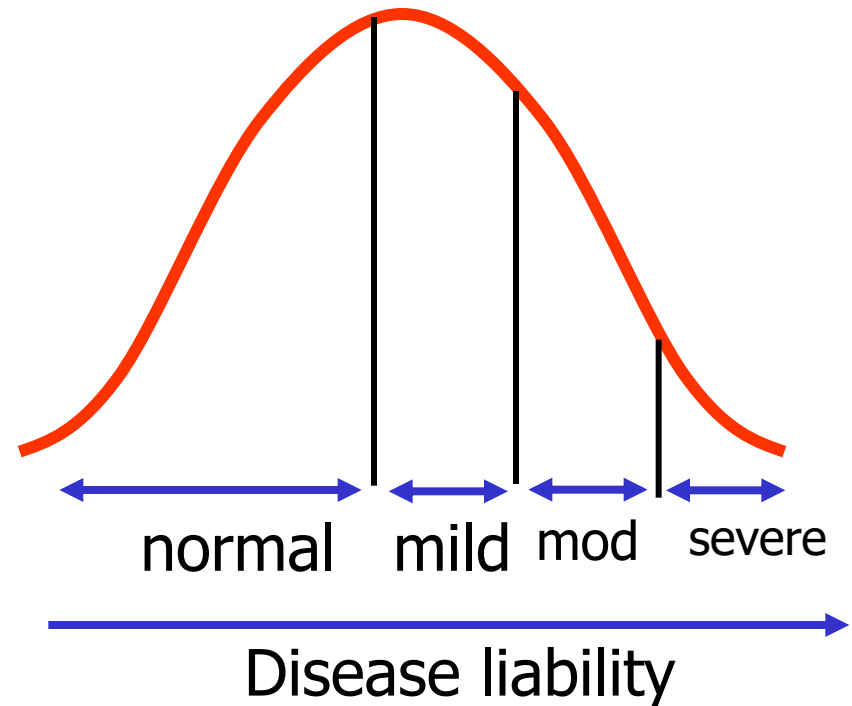
- Continuous variation caused by lots of genes (“polygenic inheritance”)
- Each gene followed Mendel’s laws
- Environment smoothed out genetic differences
- Genes may show different degrees of “dominance”
- Genes may have many forms (“multiple alleles”)
- Mating may not be random (“assortative mating”)
- Showed that correlations obtained by e.g. Pearson and Lee were explained well by polygenic inheritance
- Led to “Biometrical Genetics” (Mather, Jinks etc.)

Multifactorial Threshold Model of Disease – normally distributed “liability”

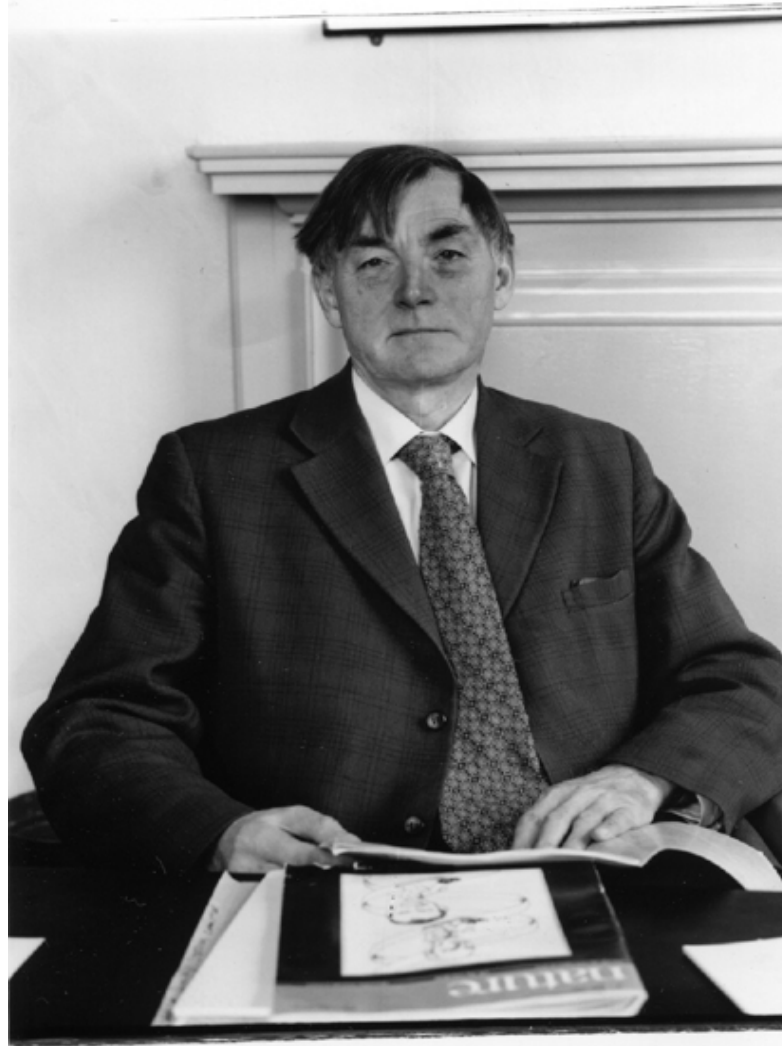
Single threshold



Multiple thresholds



Douglas Scott Falconer, FRS, FRSE (1913-2004)



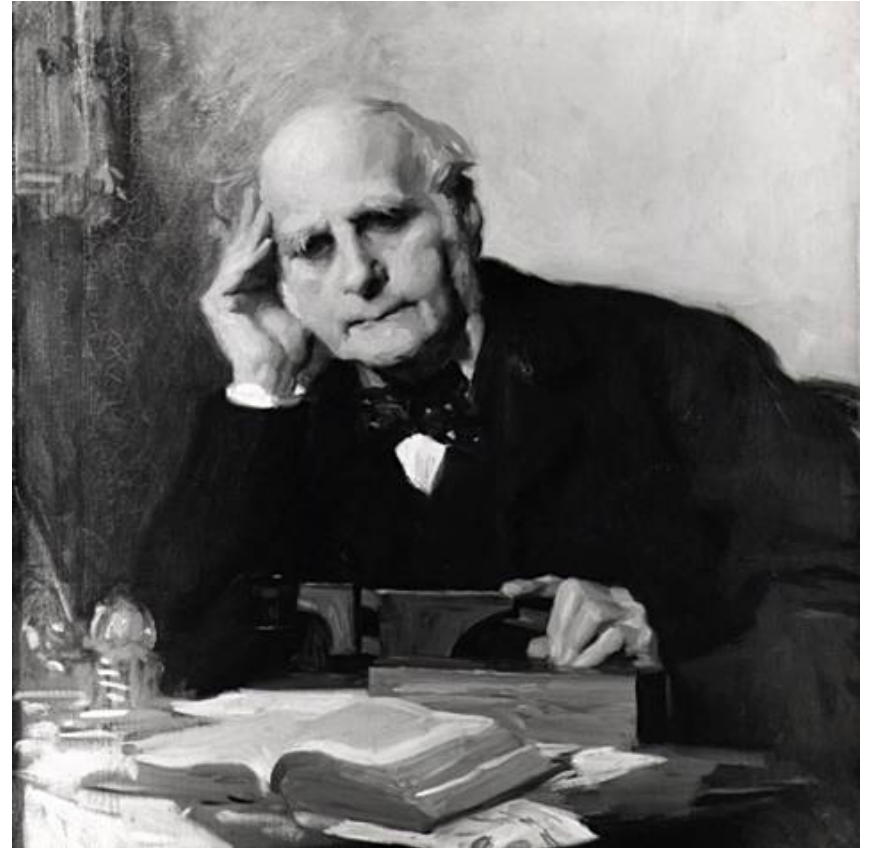
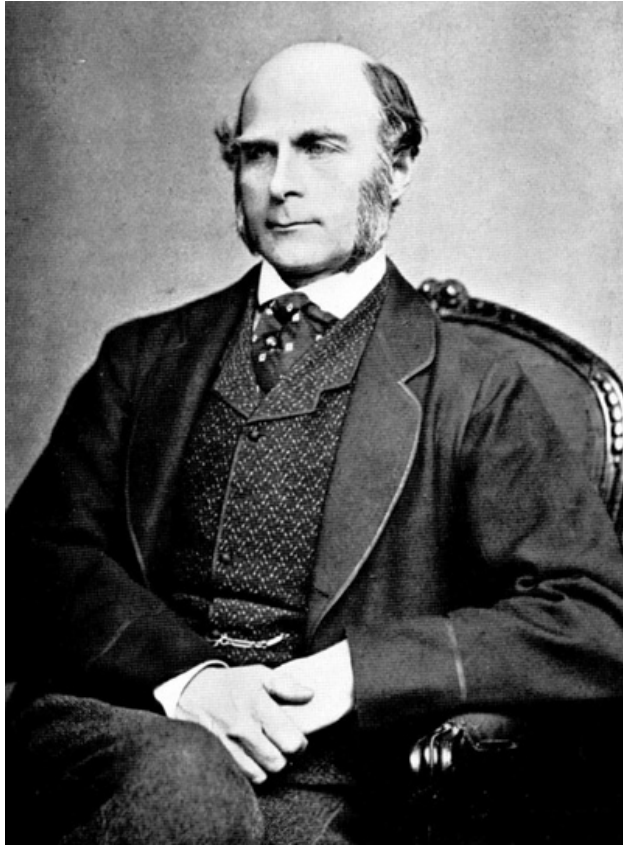
1965 Inheritance of liability to certain diseases estimated from incidence among relatives., *Ann. Hum. Genet.***29**:51ff.

1960 *Introduction to Quantitative Genetics*. Edinburgh: Oliver and Boyd.

The (Really!) BIG Problem

Families are a mixture of
genetic and social
factors

Francis Galton (1822-1911)



1869: Hereditary Genius

1883: Inquiries into Human Faculty and its Development

1884-5: Anthropometric Laboratory at "National Health Exhibition"

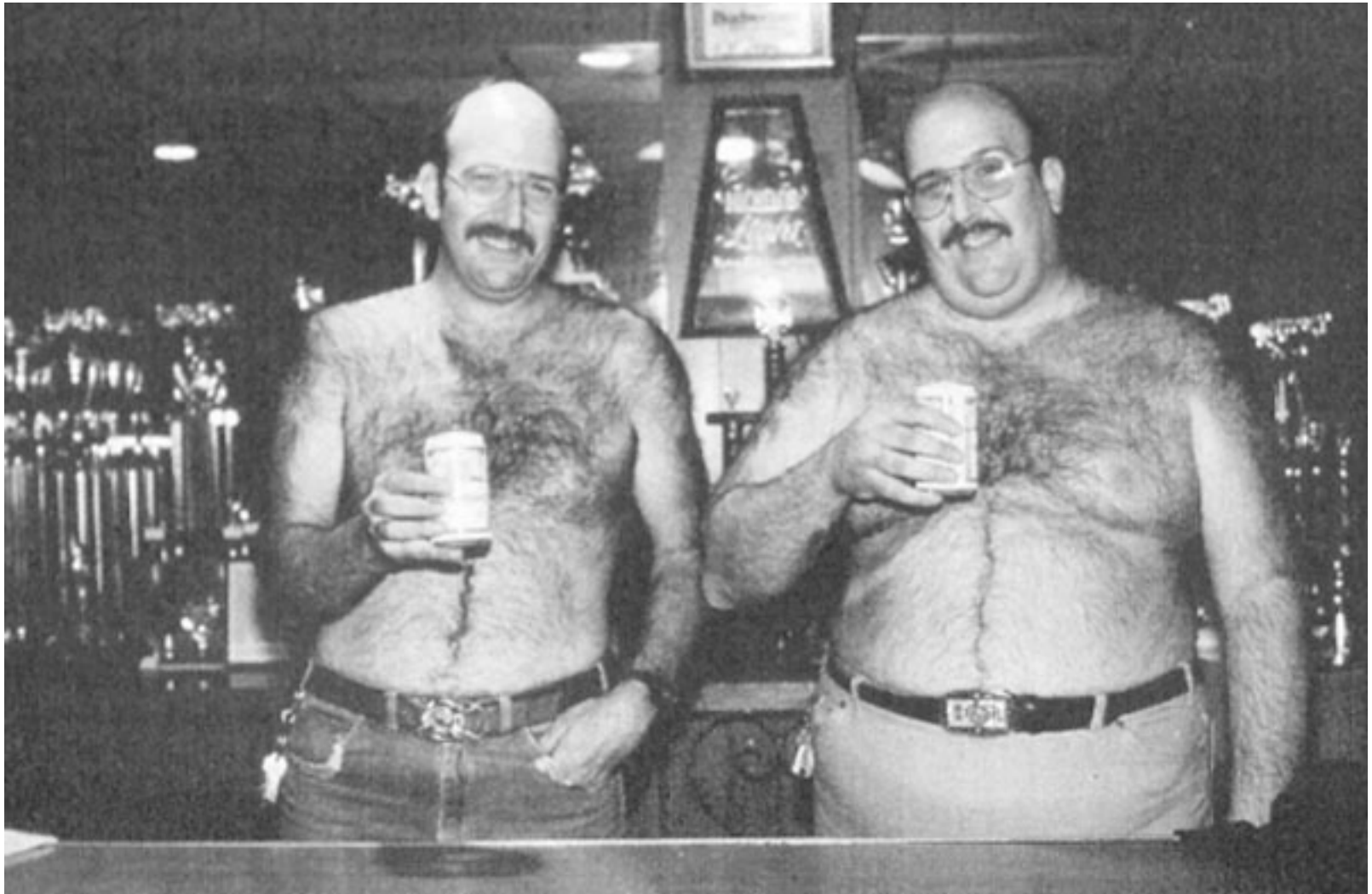
Galton's Solution:

Twins

(Though Augustine may
have got there first –
5th cent.)

One (?ideal) solution

Twins separated at
birth



MZ twins reared apart - note the same way of supporting their cans of beer

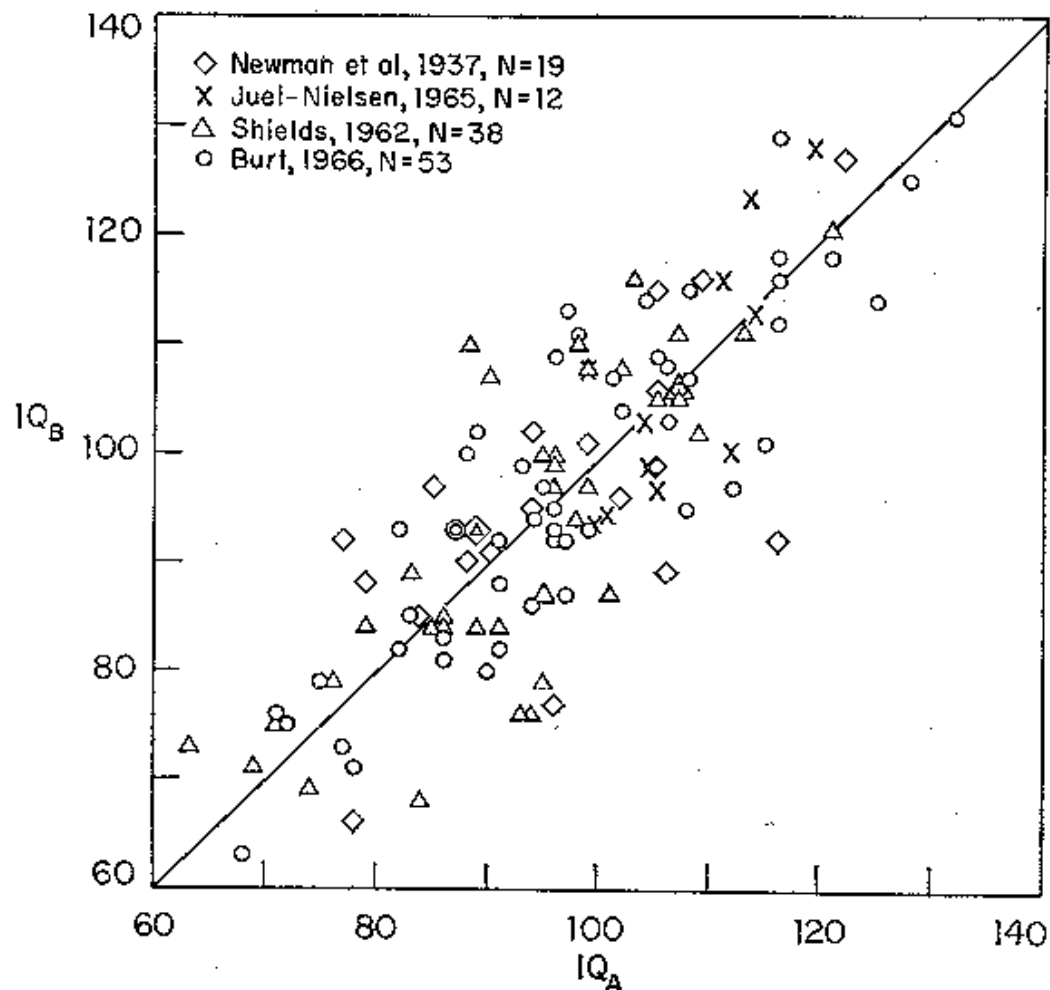


FIGURE 2. Scatter diagram showing correlation between IQs of 122 sets of co-twins (A and B assigned at random). The obtained intra-class correlation (r_i) is 0.82. The diagonal line represents perfect correlation ($r_i = 1.00$).

But separated MZs are rare

An easier alternative:

Identical and non-identical twins reared together
Galton (again!).....Or was it ?

Early Research on Human Genetics Using
the Twin Method: Who Really Invented
the Method?

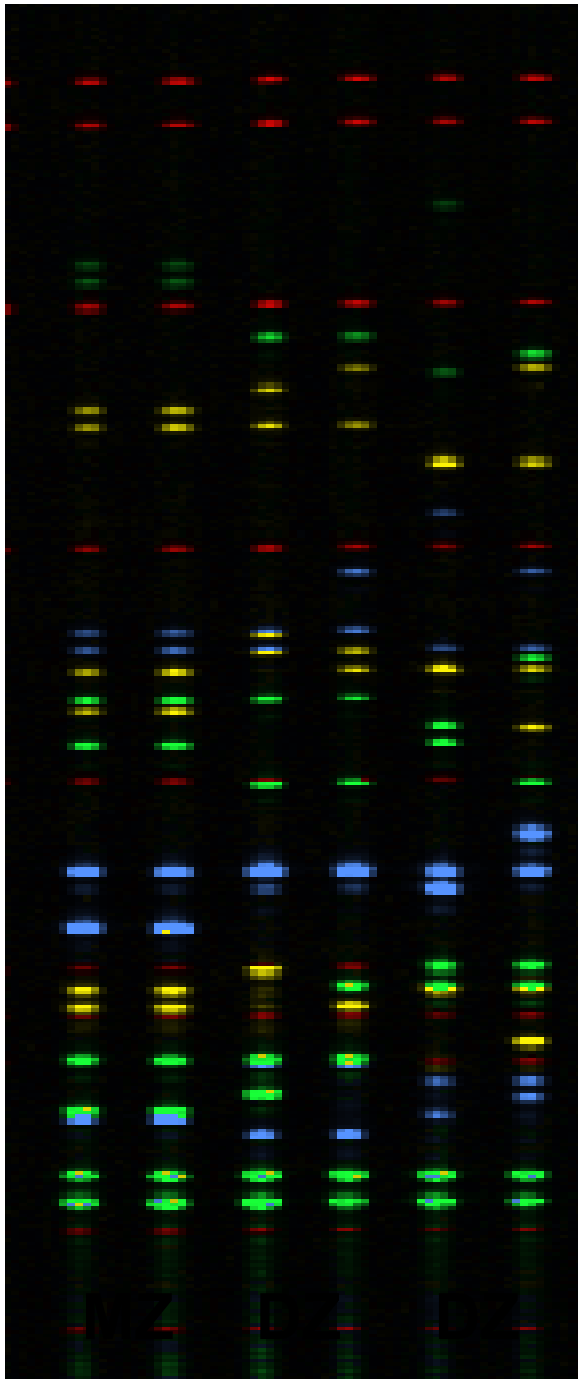
Oliver Mayo

CSIRO Livestock Industries, Australia

Twin Research and Human Genetics 12: 237–245, 2009

Poll, H. (1914). Über Zwillingsforschung als Hilfsmittel
menschlicher Erbkunde. *Zeitschrift für Ethnologie* 46, 87–105.

Identity at marker loci - except for rare mutation



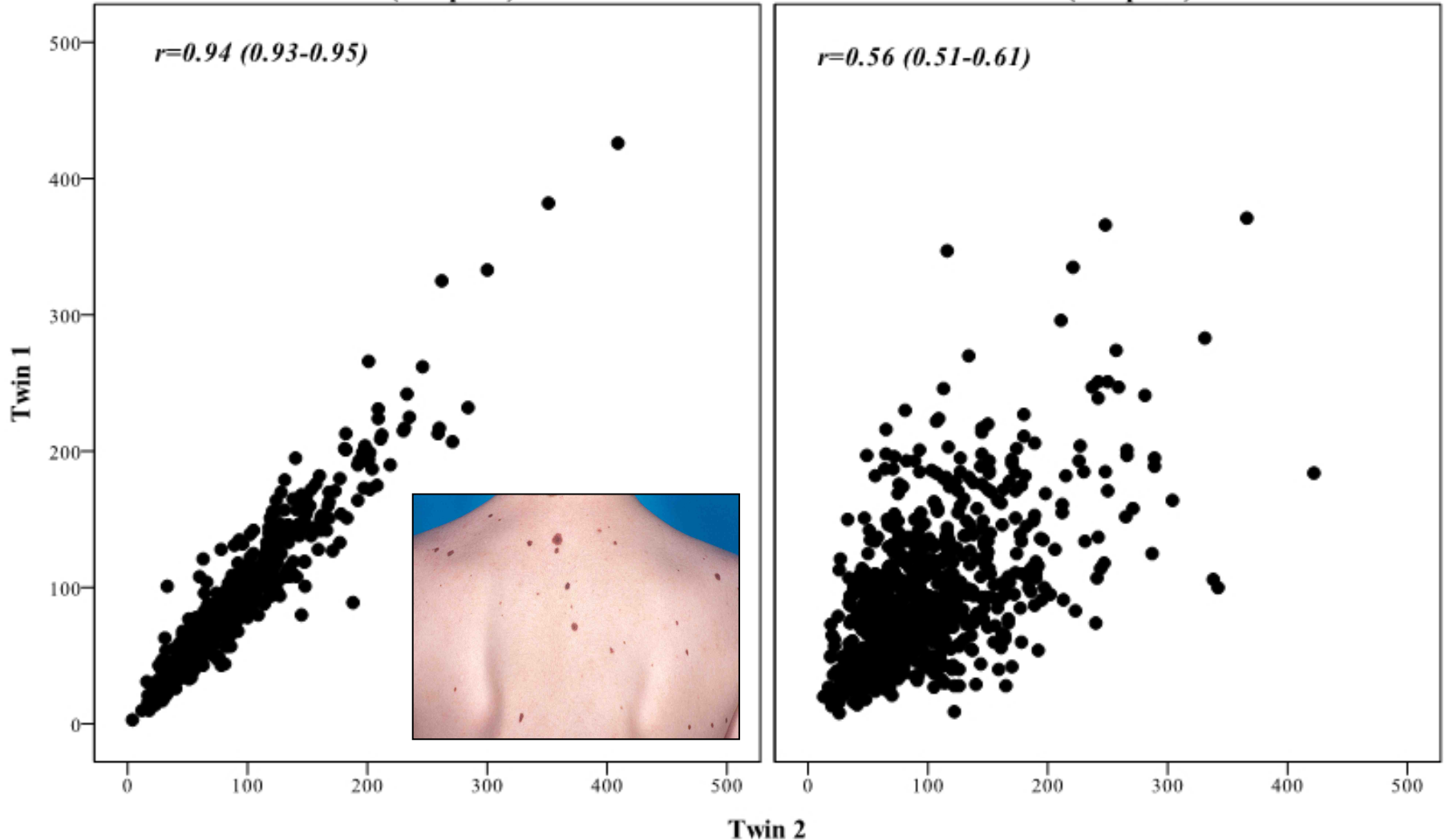
**MZ and DZ twins:
determining zygosity using
ABI Profiler™ genotyping
(9 STR markers + sex)**

Twin correlations for total mole count

Twin Correlations (95%CI) for Total Mole Count (Age 12 years)

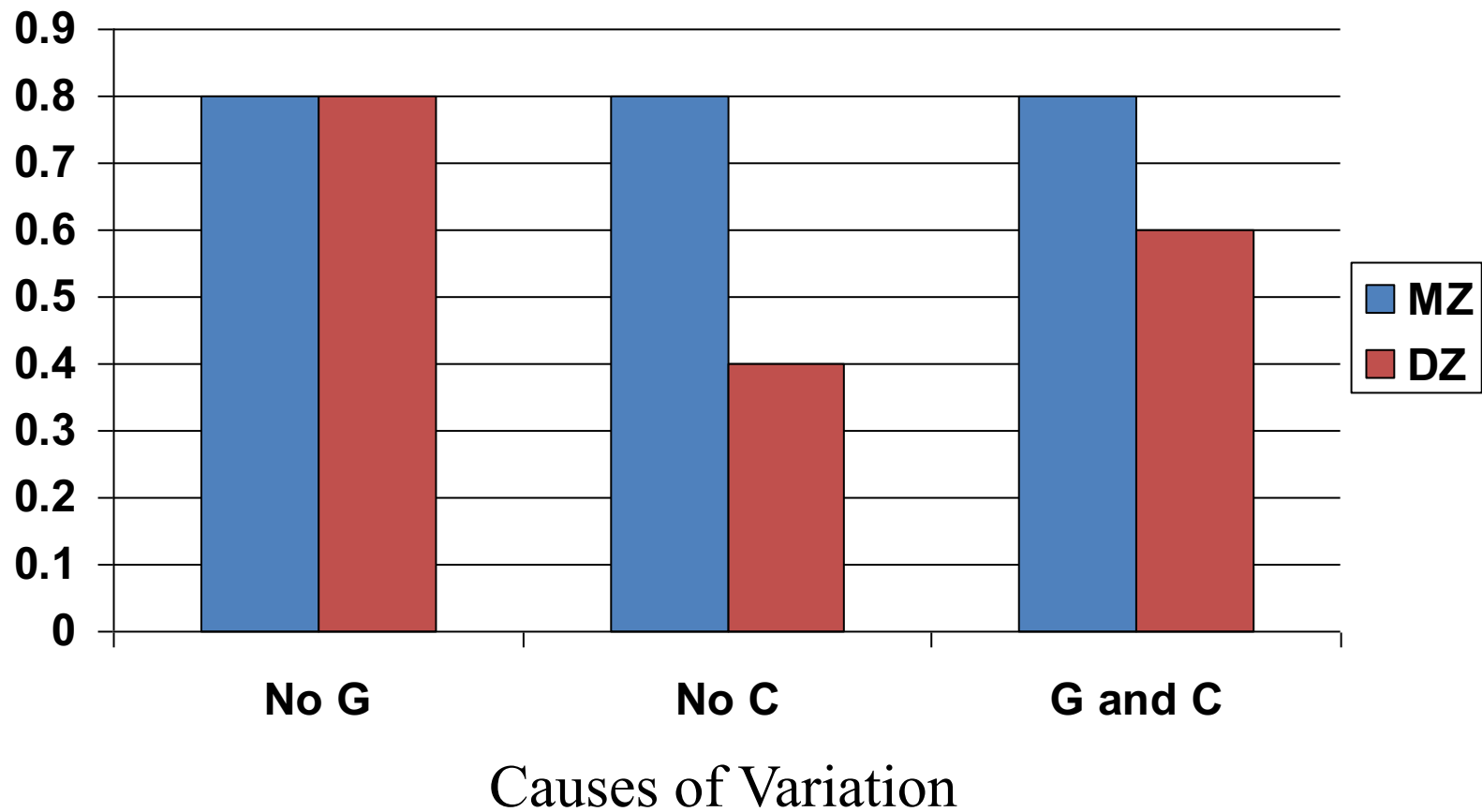
MZ (422 pairs)

DZ (762 pairs)

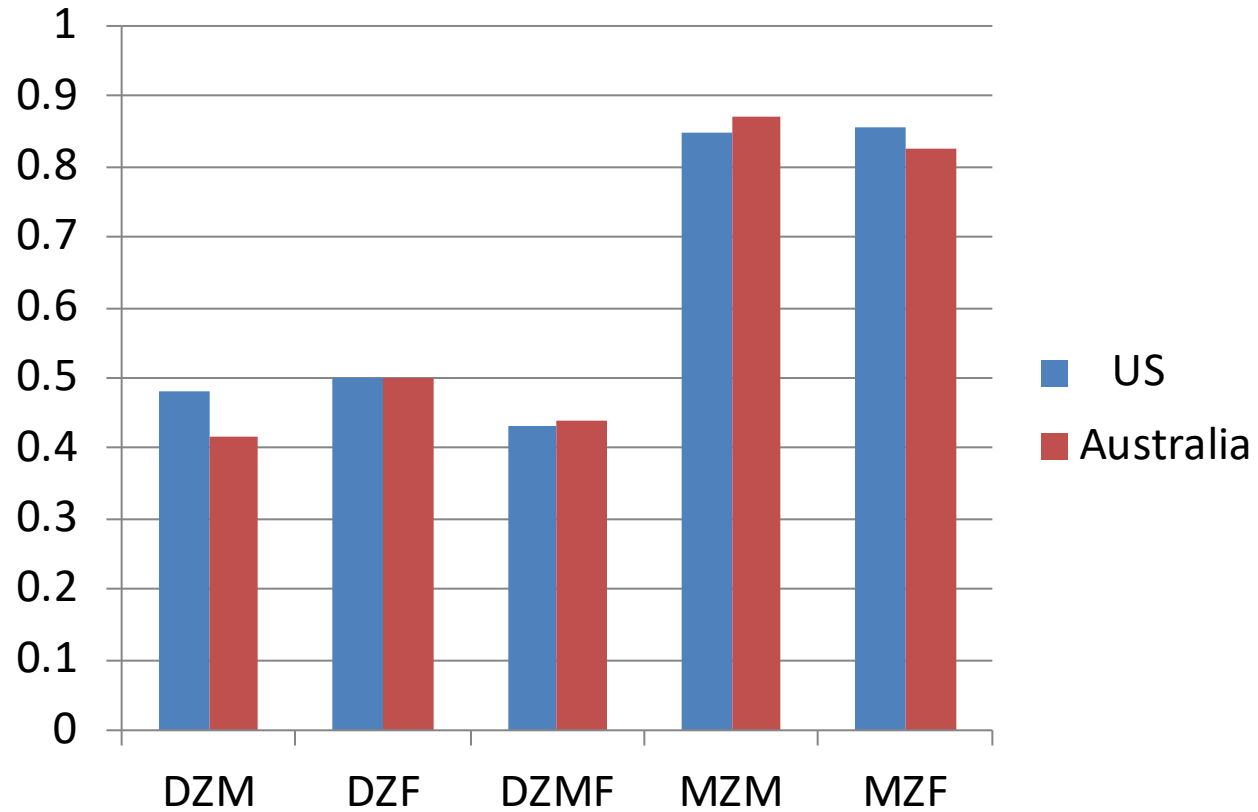


Three scenarios

Twin Correlation



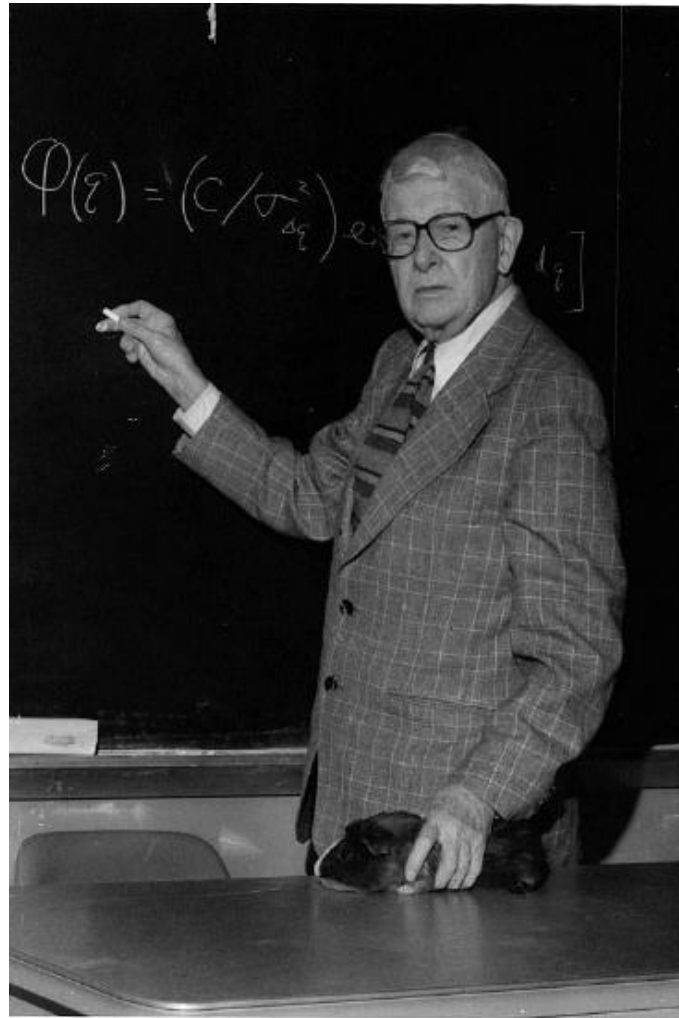
Twin Correlations for Adult Stature (Virginia 30,000 and Australia 22,000)



Structural equation modeling

- Both continuous and categorical variables
- Systematic approach to hypothesis testing
- Tests of significance
- Can be extended to:
 - More complex questions
 - Multiple variables
 - Other relatives

Sewall Wright (1889-1988)

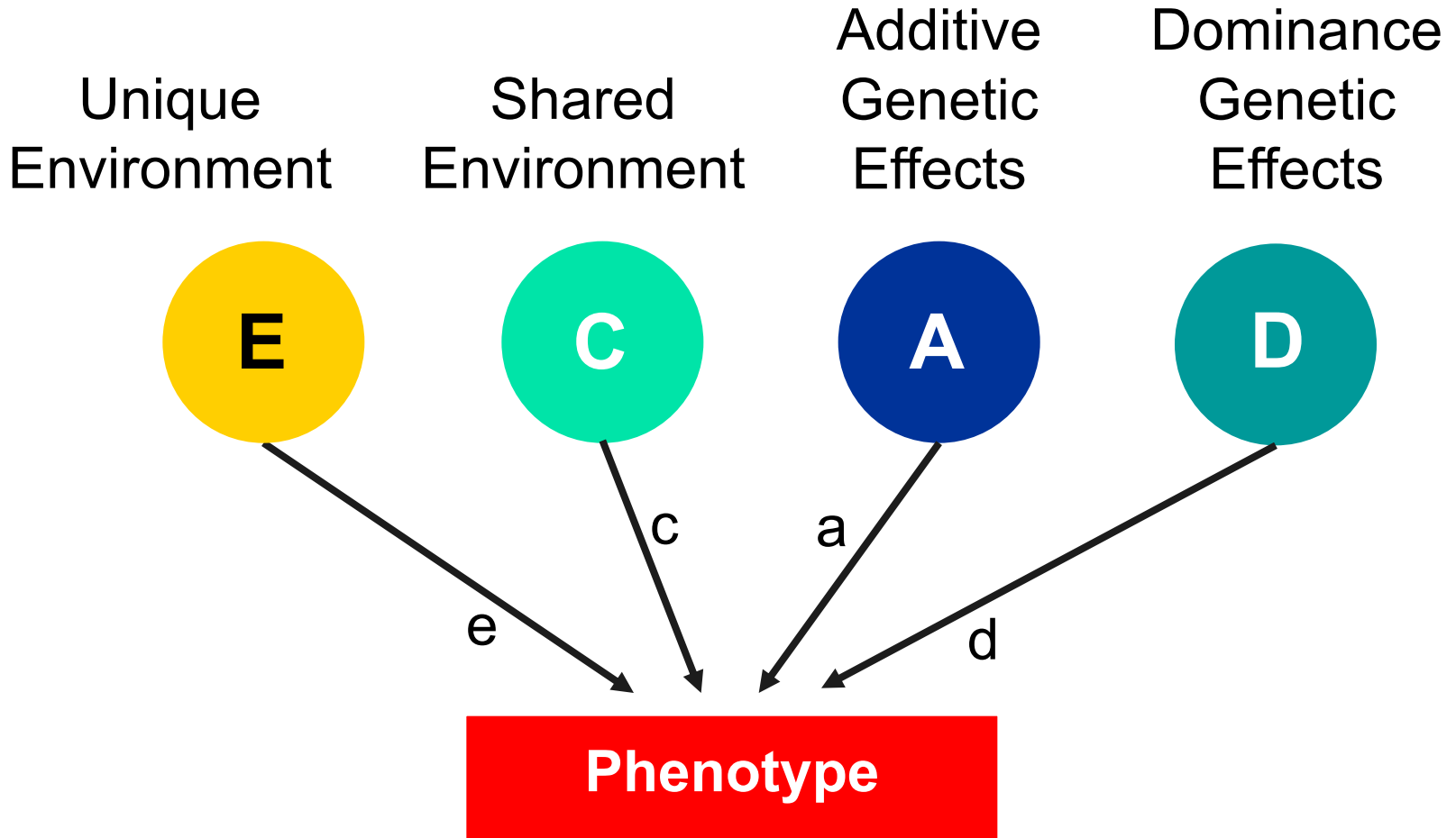


1984. Evolution and the Genetics of Populations:

Genetics and Biometric Foundations (4 vols.)

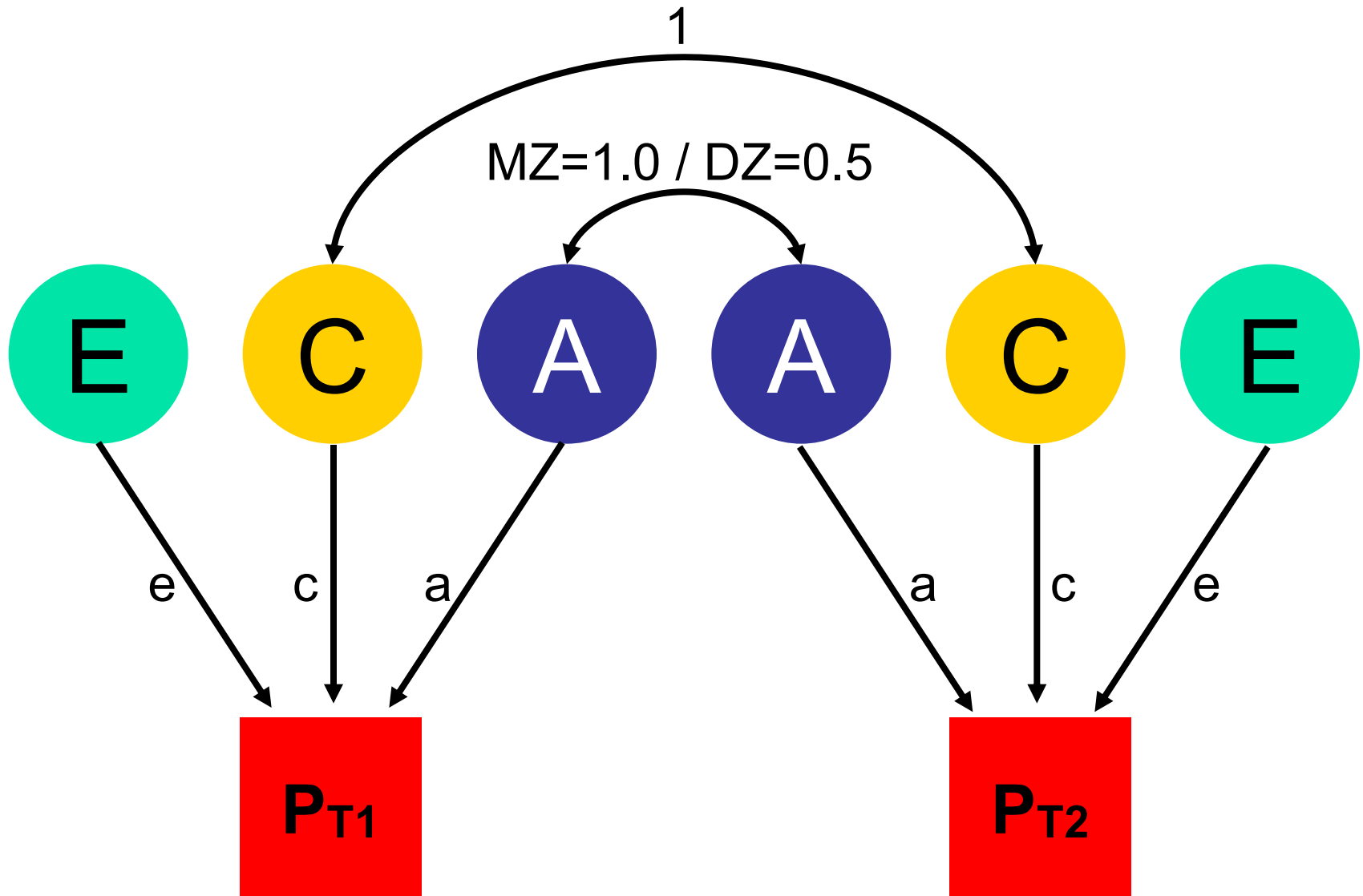
1934. The method of path coefficients. *Ann. Math. Statist.* 5: 161-215

Variance components

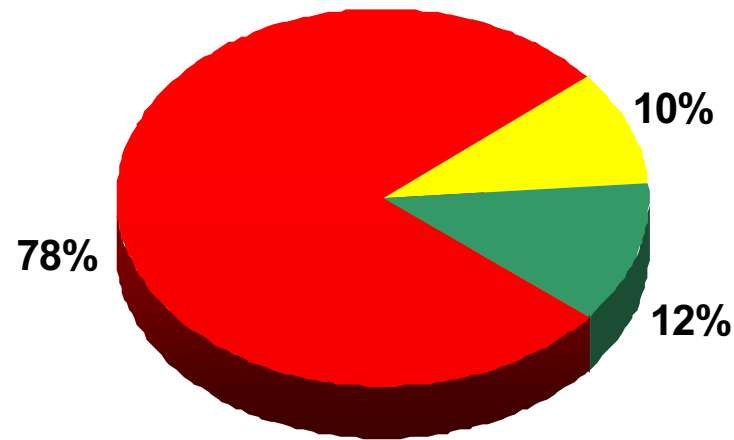


$$P = eE + aA + cC + dD$$

ACE Model for twin data



Sources of variation in exam (QCST) results



 **Additive genetic**

 **Shared environment**

 **Non-shared environment**

Problems with the ACE model (in the classical twin design)

- C and D are negatively confounded and a model with ACDE is not identified – can only fit one or the other of C/D-
 - If $r_{mz} = r_{dz} > \text{pure C}$: CE model fits best
 - If $r_{mz} = 2r_{dz} > \text{pure A}$: AE model fits best
 - If $r_{mz} < 2r_{dz} > \text{A plus C}$: ACE model fits best
 - If $r_{mz} > 2r_{dz} > \text{A plus D}$: ADE model fits best
 - If $r_{mz} = r_{dz} = 0 > \text{all E}$: no familial resemblance

Other problems (2)

- C can be inflated by age, sex differences; also batch effects; correct for these using Means model in Mx
- Low power to detect D
- D subsumes epistasis and other non-additive genetic effects
- E includes measurement error – value of test-retest data

Other problems (3)

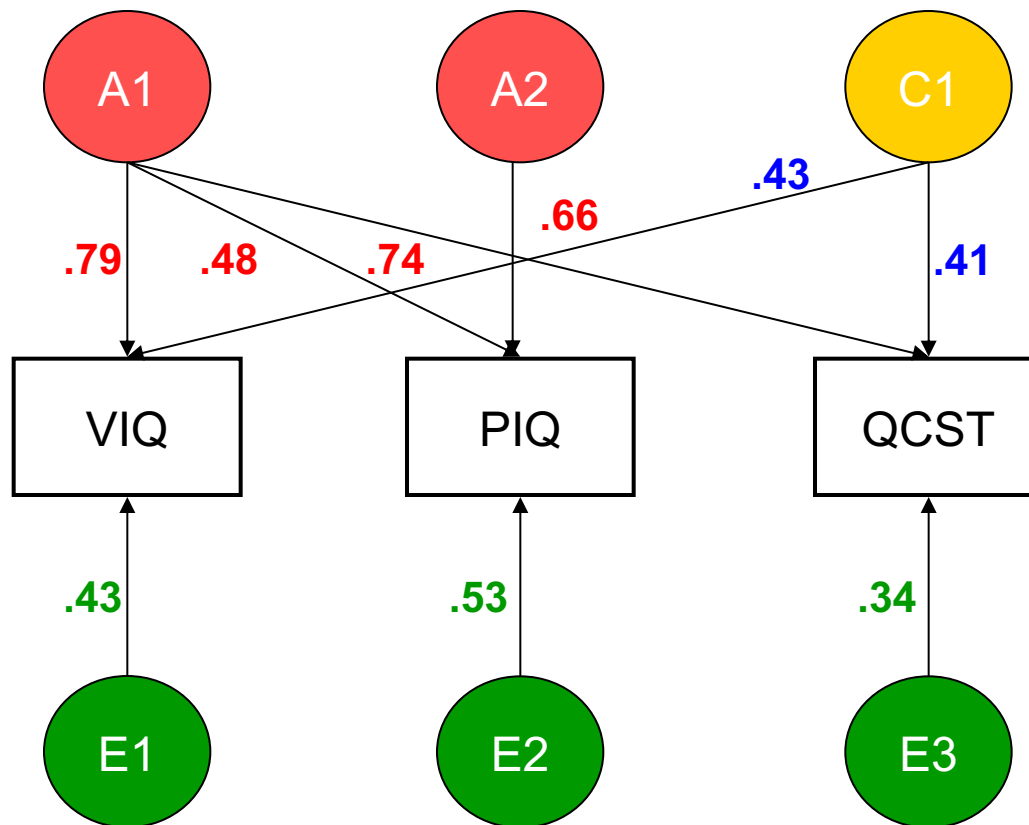
- C can be inflated by phenotypic assortative mating (like marrying like)
 - Important to collect spouse data to estimate marital correlation
 - Height ~ 0.2
 - Education, smoking, drinking, politics ~ 0.6

**CAN ACCOUNT FOR SOME/MOST/ALL OF "C"
ESTIMATED FROM ACE MODEL IN TWIN DATA**

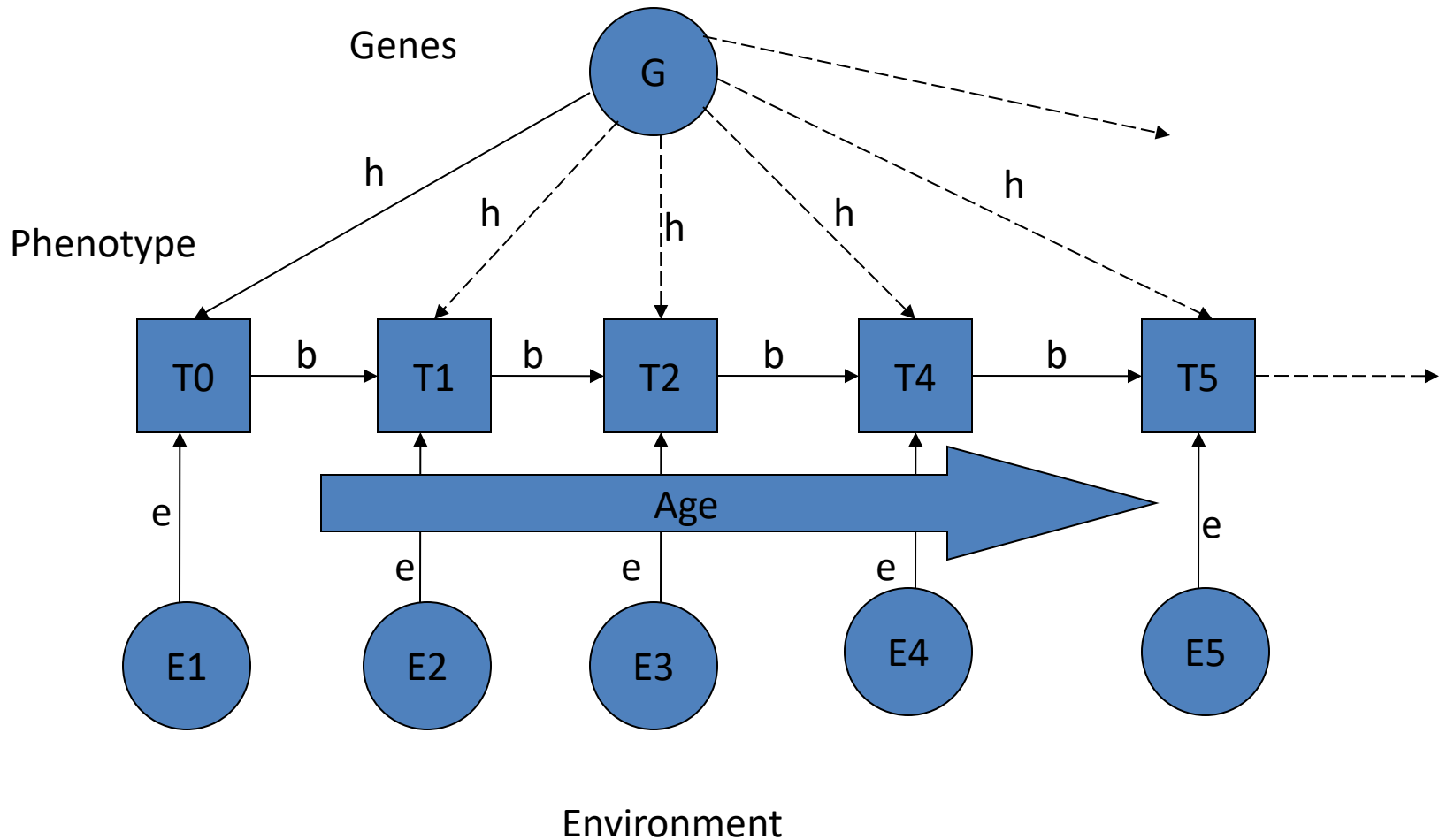
How twin studies changed research agendas

1. Autism – “caused by cold mothers”
10/11 MZ pairs concordant, 2/11 DZs concordant
2. ADHD – “caused by food dyes”
Twin studies found $h^2 \sim 0.8$
3. Multiple sclerosis - “caused by a virus”
MZ concordance 26%, DZ concordance 2%

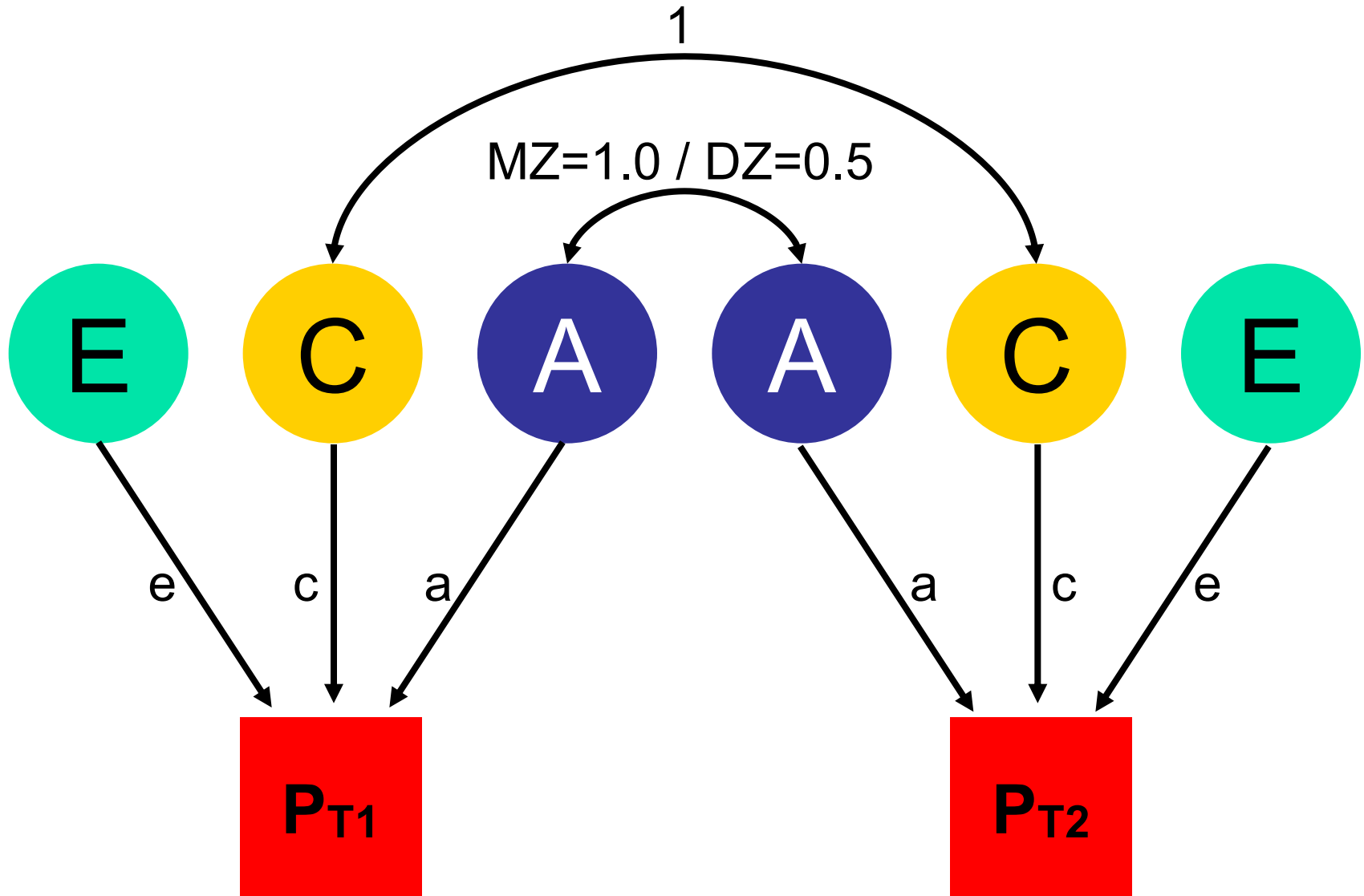
Are there common genetic and environmental factors influencing Verbal IQ, Performance IQ and QCST results?



Genetic variation in **developmental change**: time series with common genes and time-specific environmental “innovations”



ACE Model for twin data



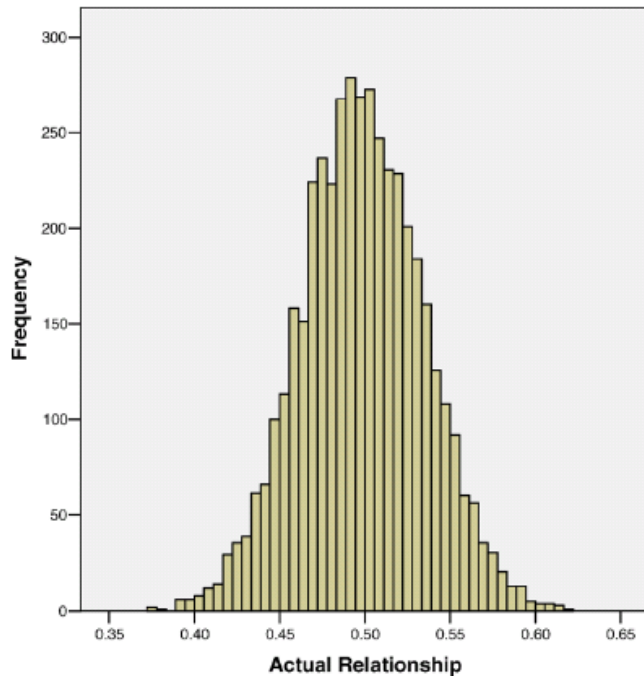


Figure 1. Empirical Distribution of Actual Additive Genetic Relationships of 4,401 Quasi-Independent Pairs of Full Sibs
Histogram of the genome-wide additive genetic relationships of full-sib pairs estimated from genetic markers.
DOI: 10.1371/journal.pgen.0020041.g001

Classical twin design revisited: Heritability estimation without MZ twins

Why do we use the average sib value of $r_a = 0.5$ when we can estimate the (almost) exact values for each sib pair from genetic marker data ?

OPEN ACCESS Freely available online

PLoS GENETICS

Assumption-Free Estimation of Heritability from Genome-Wide Identity-by-Descent Sharing between Full Siblings

Peter M. Visscher^{*}, Sarah E. Medland, Manuel A. R. Ferreira, Katherine I. Morley, Gu Zhu, Belinda K. Cornes, Grant W. Montgomery, Nicholas G. Martin

Genetic Epidemiology Group, Queensland Institute of Medical Research, Brisbane, Australia

Table 2. ML Estimates of Heritability of Height from Genome-Wide IBD Sharing between Sib Pairs

Data	Model	Estimates (95% CI)	
		f^2	h^2
Adolescents ($n = 931$)	FAE	0.00 (0.00–0.43)	0.80 (0.00–0.90)
	FE	0.40 (0.34–0.45)	
Adults ($n = 2,444$)	FAE	0.00 (0.00–0.18)	0.80 (0.43–0.86)
	FE	0.39 (0.36–0.43)	
Combined ($n = 3,375$)	FAE	0.00 (0.00–0.17)	0.80 (0.46–0.85)
	FE	0.39 (0.36–0.42)	

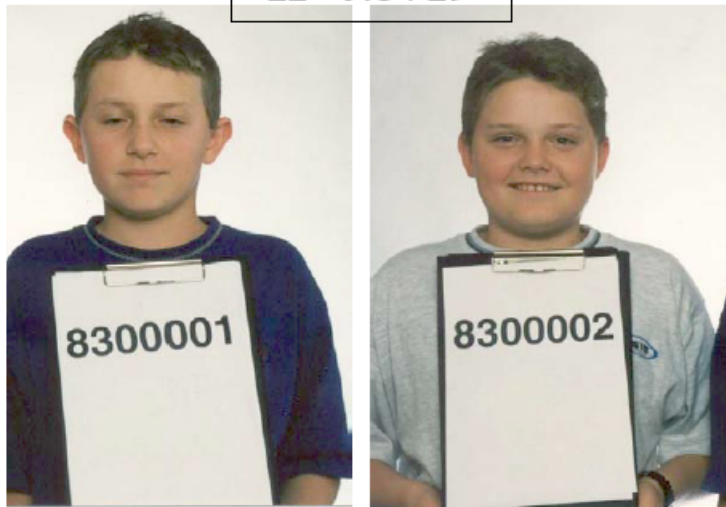
8188001,02
H=0.5677



8473001,02
H=0.5577



8300001,02
H=0.5719



8582601,02
H=0.5640



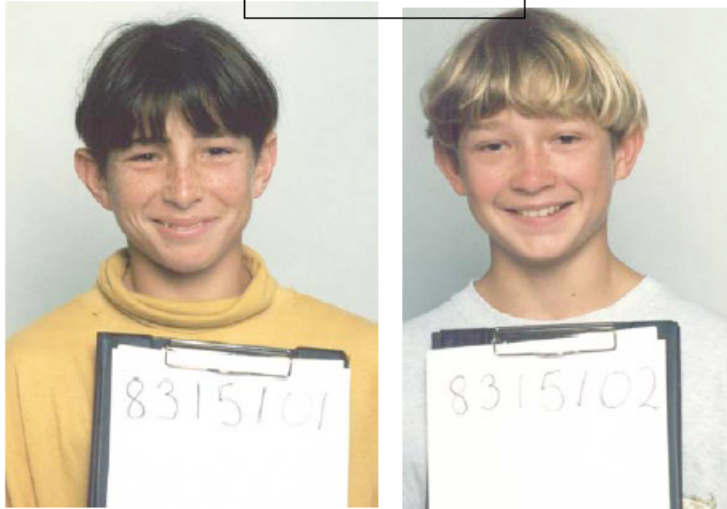
8040201,02
H=0.4351



8069101,02
H=0.4291



8315101,02
H=0.4320

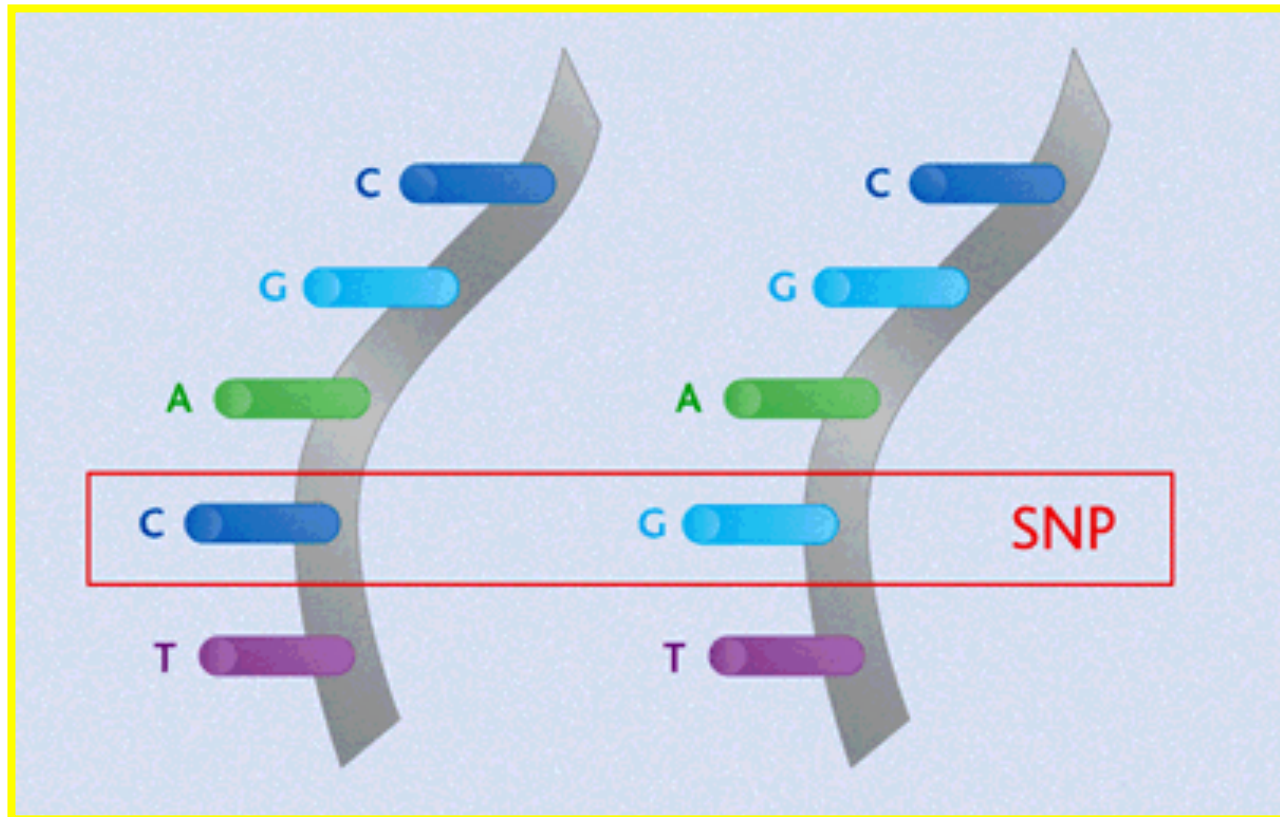


8525101,02
H=0.4385

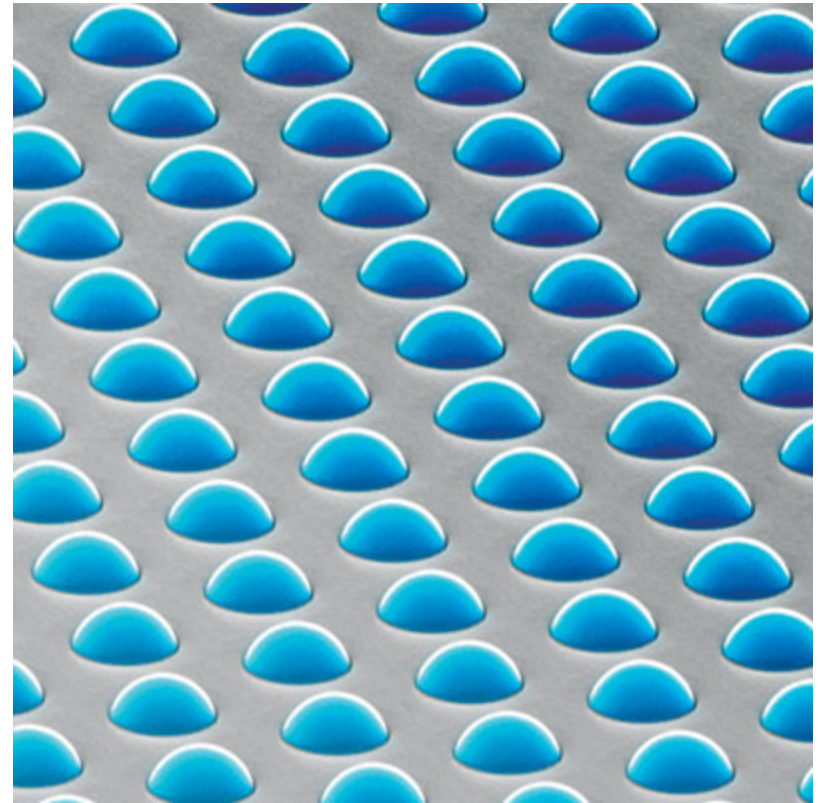


Finding the genes - association

Looks for correlation between specific alleles and phenotype (trait value, disease risk) using single nucleotide polymorphisms (SNPs)



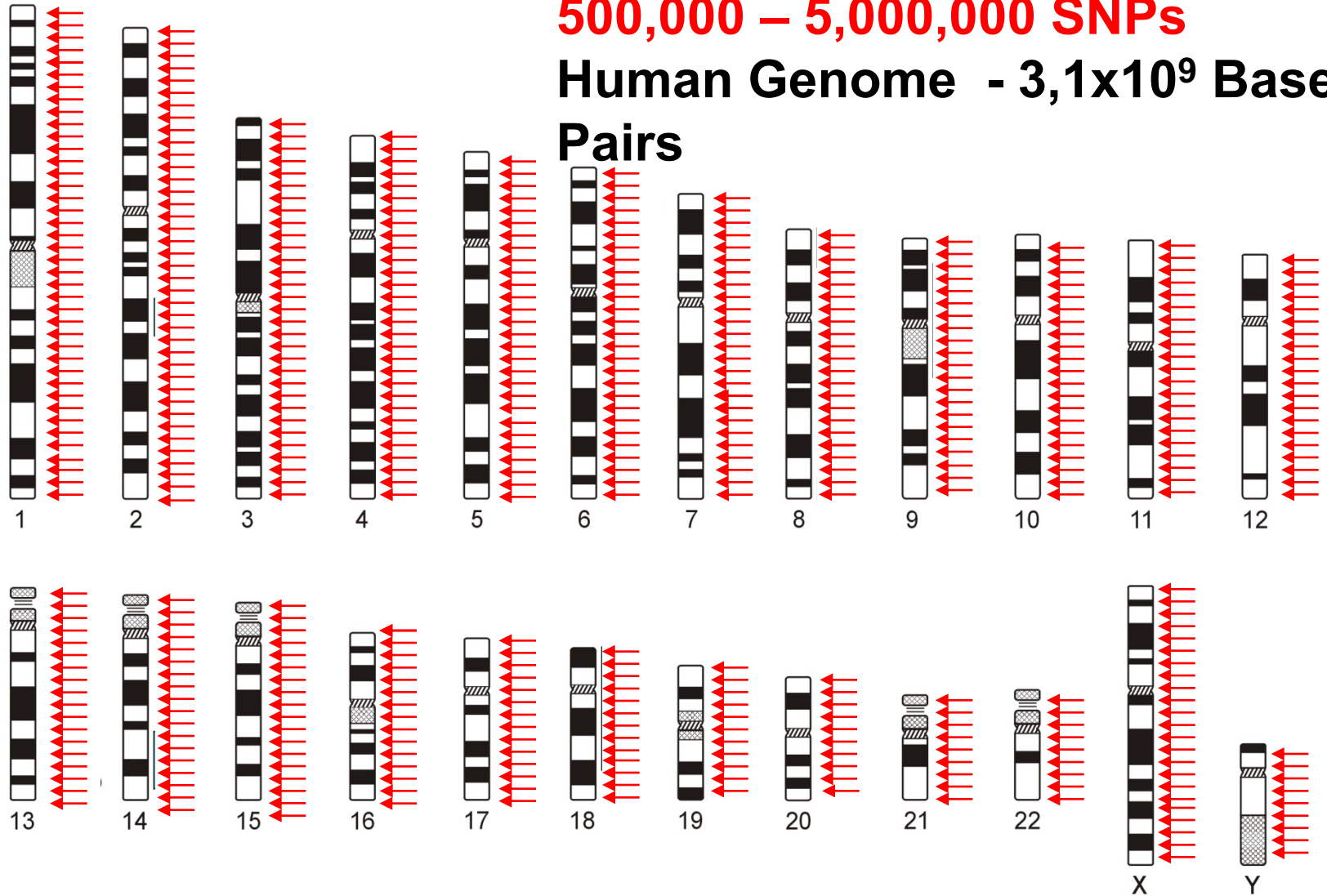
High density SNP arrays – up to 1 million SNPs



Genome-Wide Association Studies

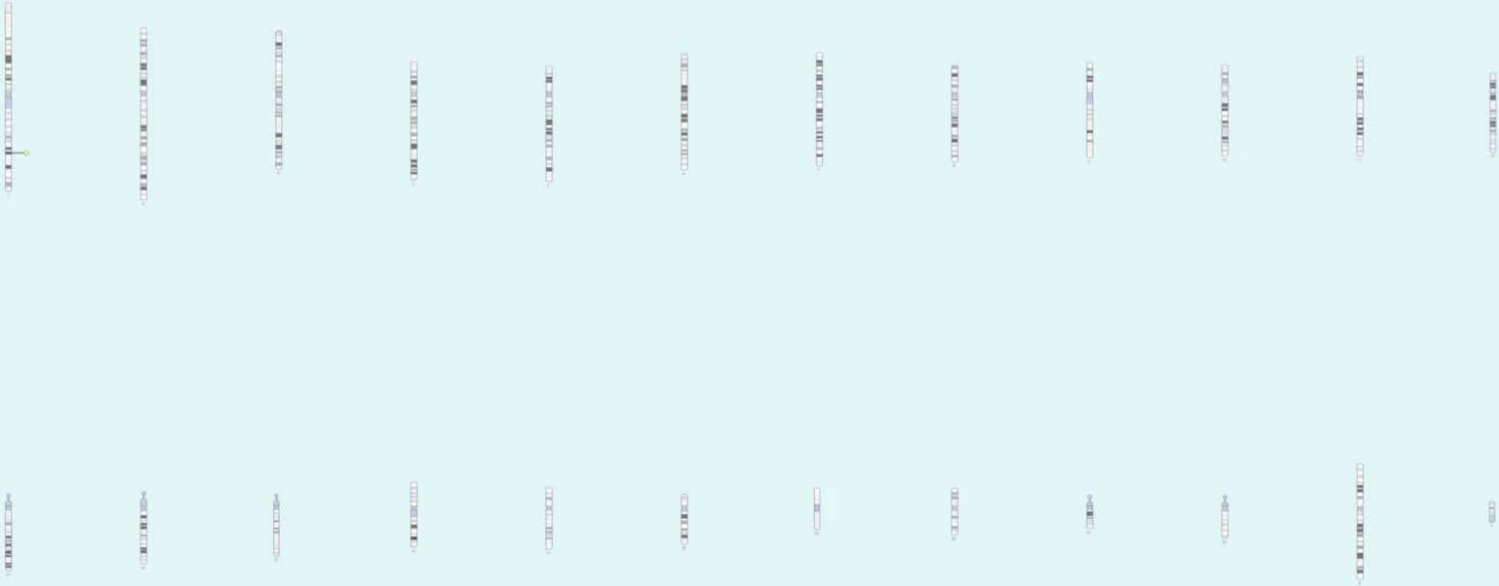
500,000 – 5,000,000 SNPs

Human Genome - $3,1 \times 10^9$ Base Pairs



The success of GWAS

2006 Jan



www.ebi.ac.uk/gwas

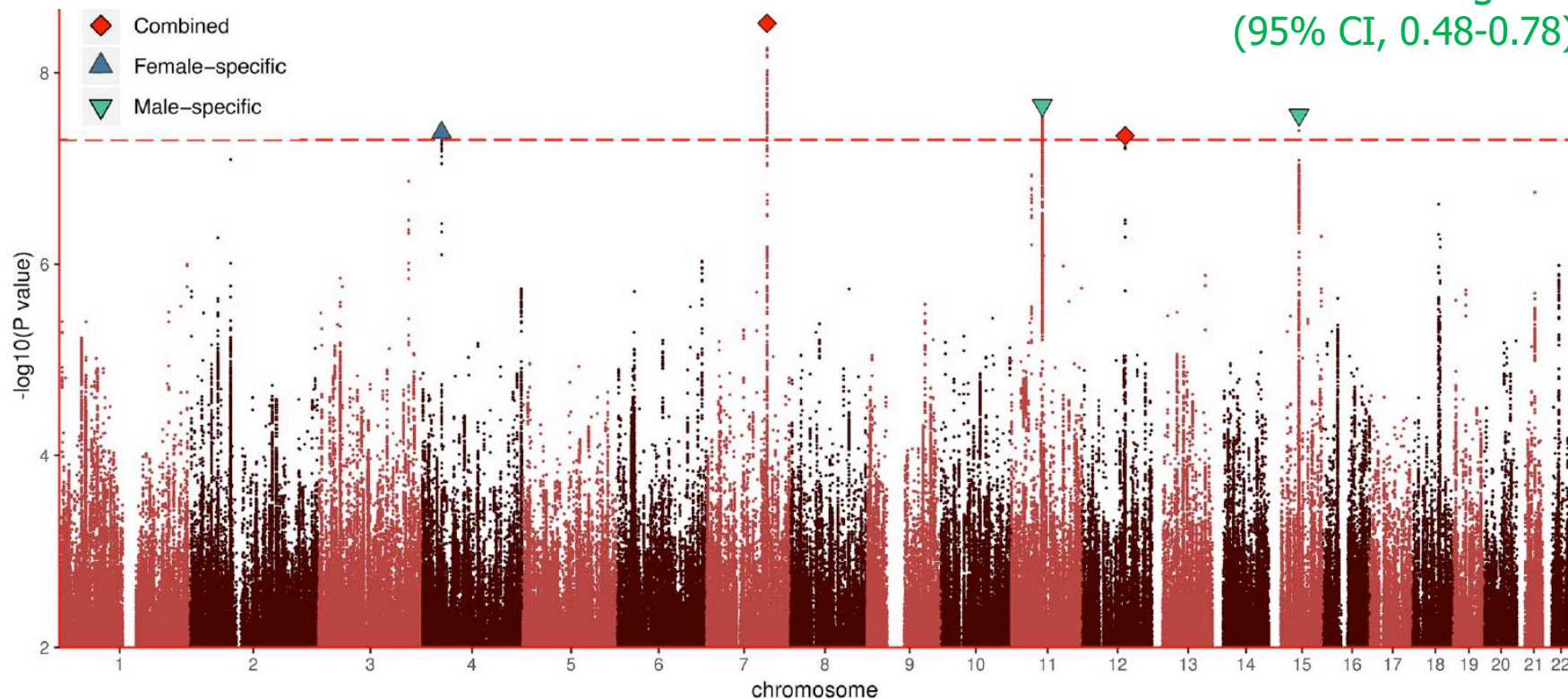
Large-scale GWAS reveals insights into the genetic architecture of same-sex sexual behavior

Total n = 477,522
(26,827 reporting same-sex sexual behavior)

Andrea Ganna^{1,2,3,4*}, Karin J. H. Verweij^{5*}, Michel G. Nivard⁶, Robert Maier^{1,2,3}, Robbee Wedow^{1,3,7,8,9,10,11}, Alexander S. Busch^{12,13,14}, Abdel Abdellaoui⁵, Shengru Guo¹⁵, J. Fah Sathirapongsasuti¹⁶, 23andMe Research Team¹⁶, Paul Lichtenstein⁴, Sebastian Lundström¹⁷, Niklas Långström⁴, Adam Auton¹⁶, Kathleen Mullan Harris^{18,19}, Gary W. Beecham¹⁵, Eden R. Martin¹⁵, Alan R. Sanders^{20,21}, John R. B. Perry^{12†}, Benjamin M. Neale^{1,2,3†}, Brendan P. Zietsch^{22††}

GCTA $h^2 = 32.4\%$
(95% CI 10.6 - 54.3)

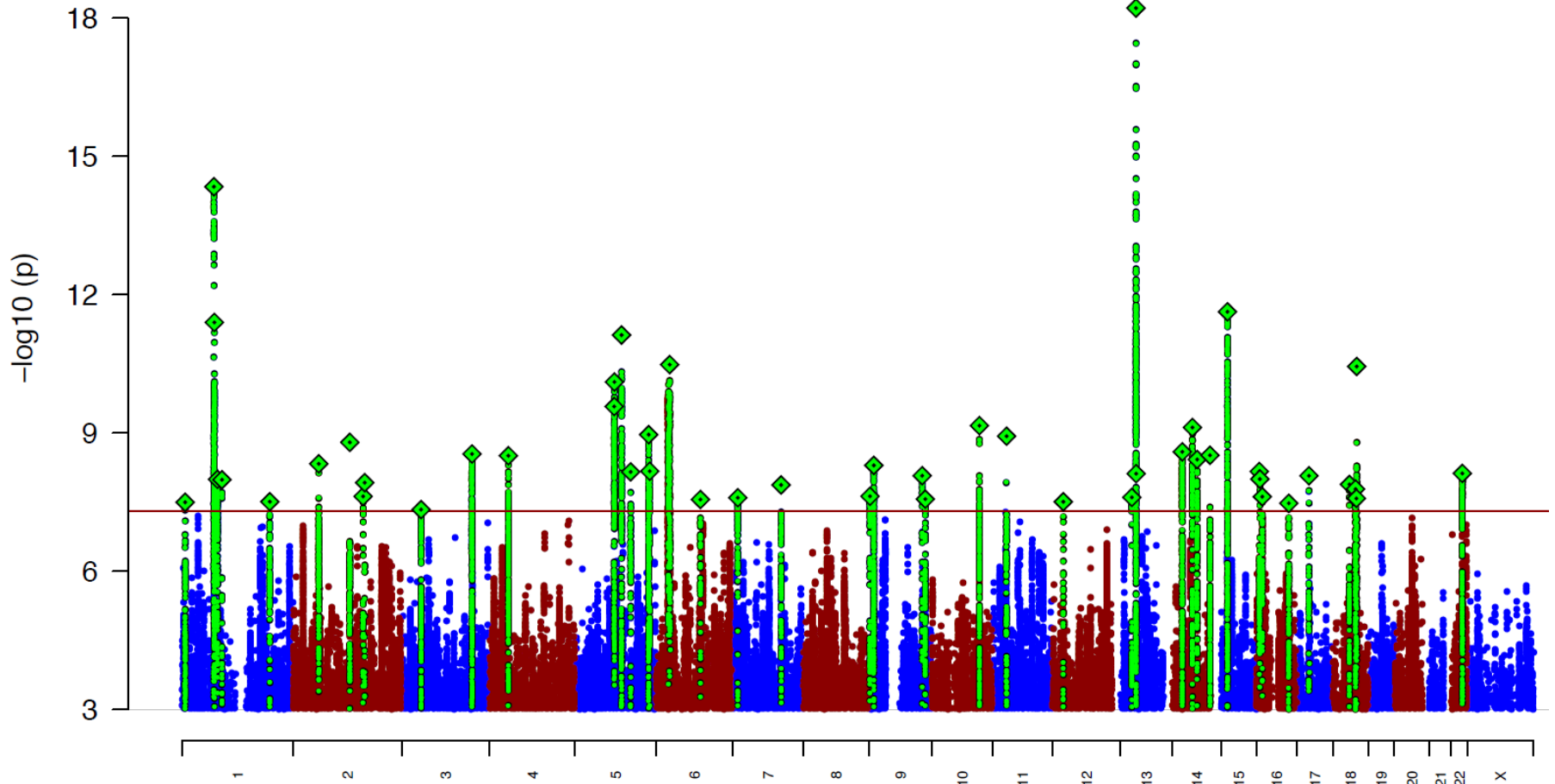
Across sex $R_g = 0.63$
(95% CI, 0.48-0.78)



Major Depressive Disorder

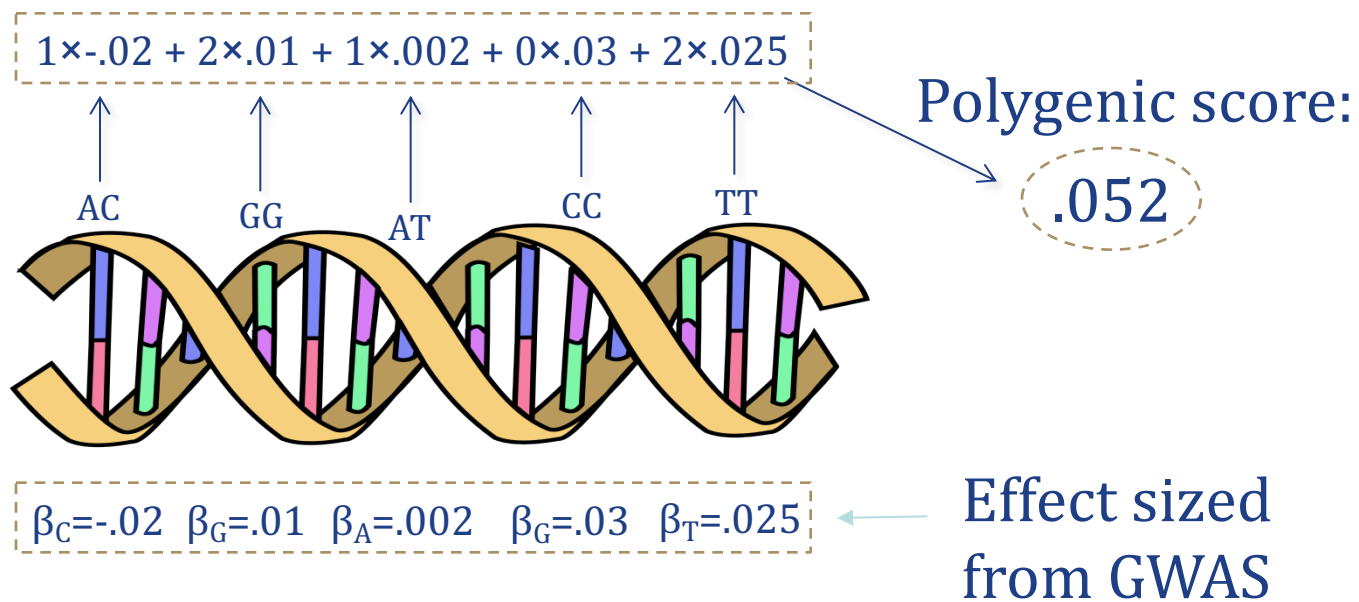
130,664 MDD cases , 330,470 controls → 44 independent loci ($P < 5e-8$)

Genes that are targets of antidepressant medications were strongly enriched for MDD association signals ($P=8.5e-10$)

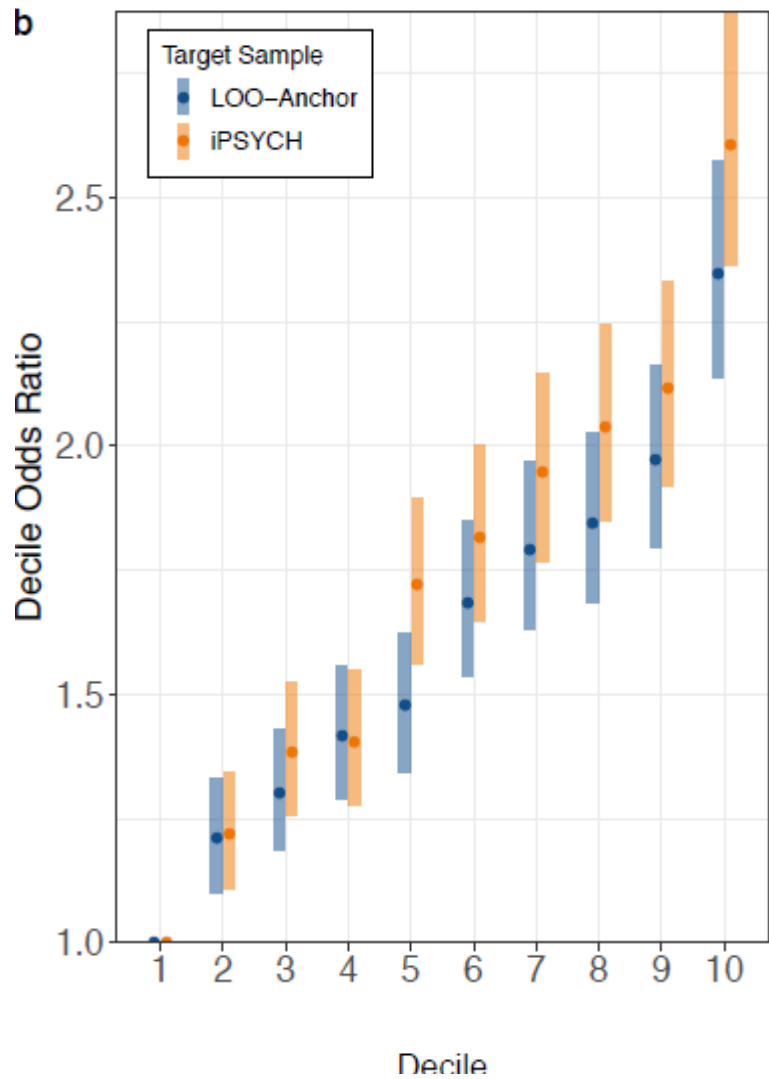


Polygenic Risk Scores

Polygenic Risk Scores capture (part of) someone's genetic "risk" by summing all risk alleles weighted by the effect sizes estimated in a Genome-Wide Association Study (GWAS)



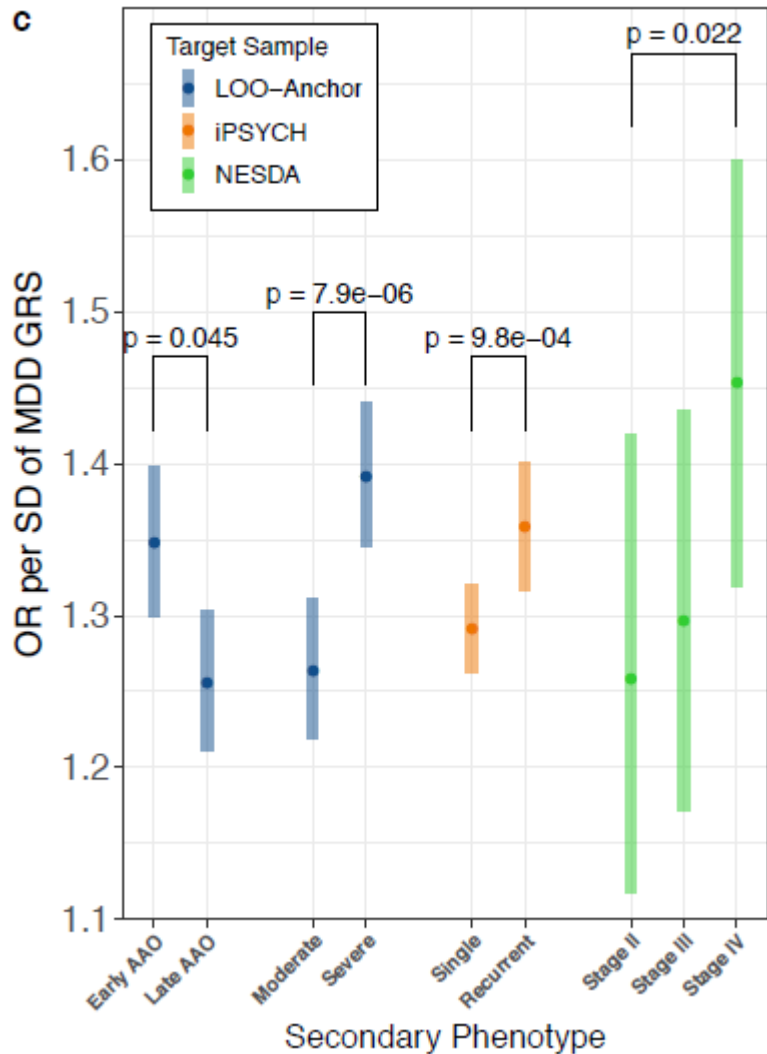
MDD Polygenic Risk Score predicts risk in independent samples



Odd ratios of MDD per PRS decile relative to the first decile for iPSYCH and anchor cohorts.

Interdecile risk ~2.5

MDD Polygenic Risk Score predicts age at onset, recurrence, and severity in independent samples



MDD PRS (from out-of-sample discovery sets) were significantly higher in MDD cases with:

- earlier age at onset; more severe MDD symptoms (based on number of criteria endorsed)*
- recurrent MDD compared to single episode*
- chronic/unremitting MDD (“Stage IV” compared to “Stage II”, first-episode MDD)*

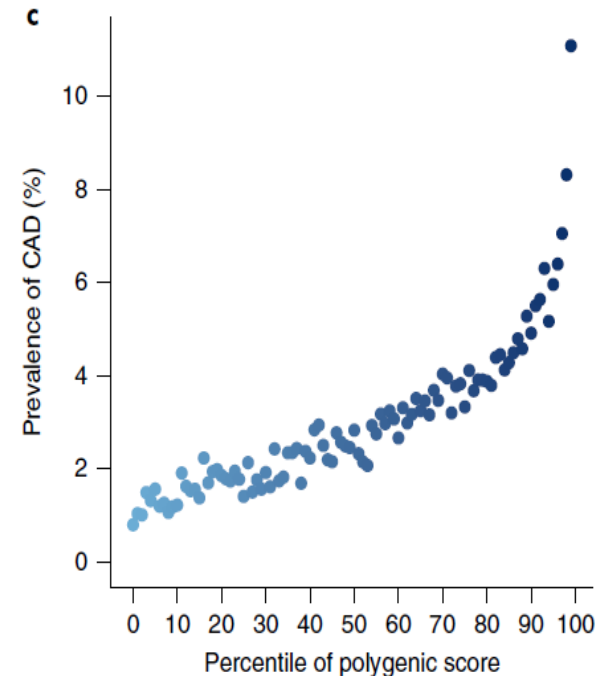
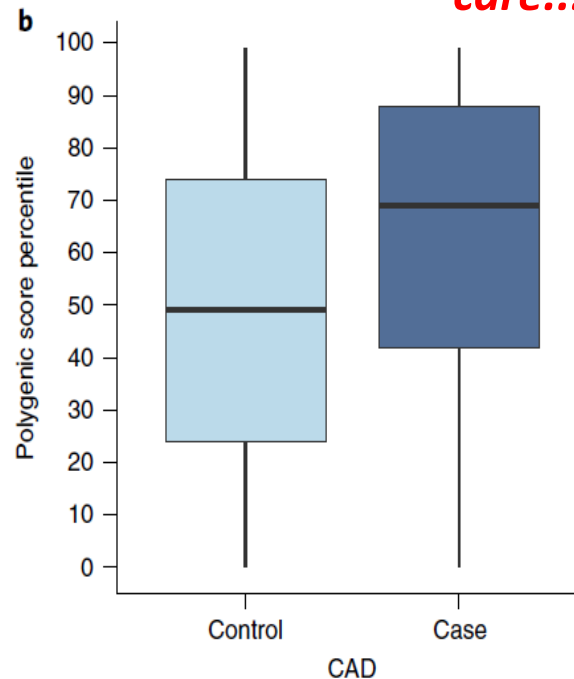
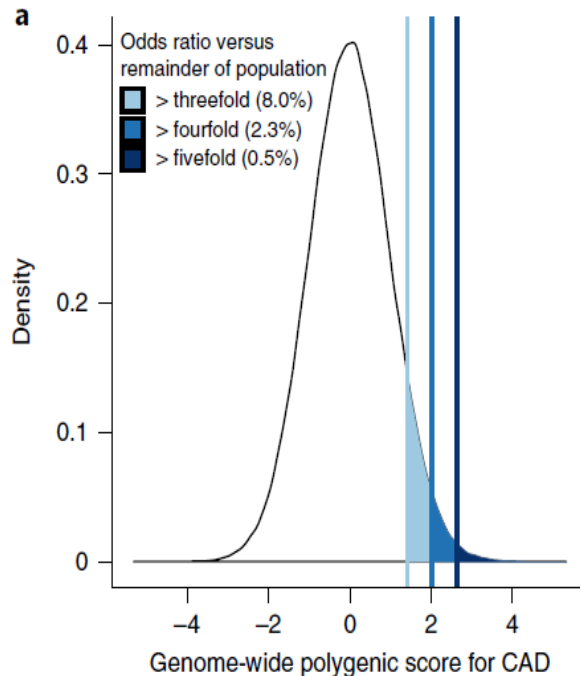
Error bars represent 95% CI

- CAD 8.0%,
- atrial fibrillation 6.1%
- type 2 diabetes 3.5%
- IBD 3.2%
- breast cancer 1.5%

Genome-wide polygenic scores for common diseases identify individuals with risk equivalent to monogenic mutations

Amit V. Khera^{1,2,3,4,5}, Mark Chaffin^{4,5}, Krishna G. Aragam^{1,2,3,4}, Mary E. Haas⁴, Carolina Roselli⁴, Seung Hoan Choi⁴, Pradeep Natarajan^{2,3,4}, Eric S. Lander⁴, Steven A. Lubitz^{2,3,4}, Patrick T. Ellinor^{2,3,4} and Sekar Kathiresan^{1,2,3,4*}

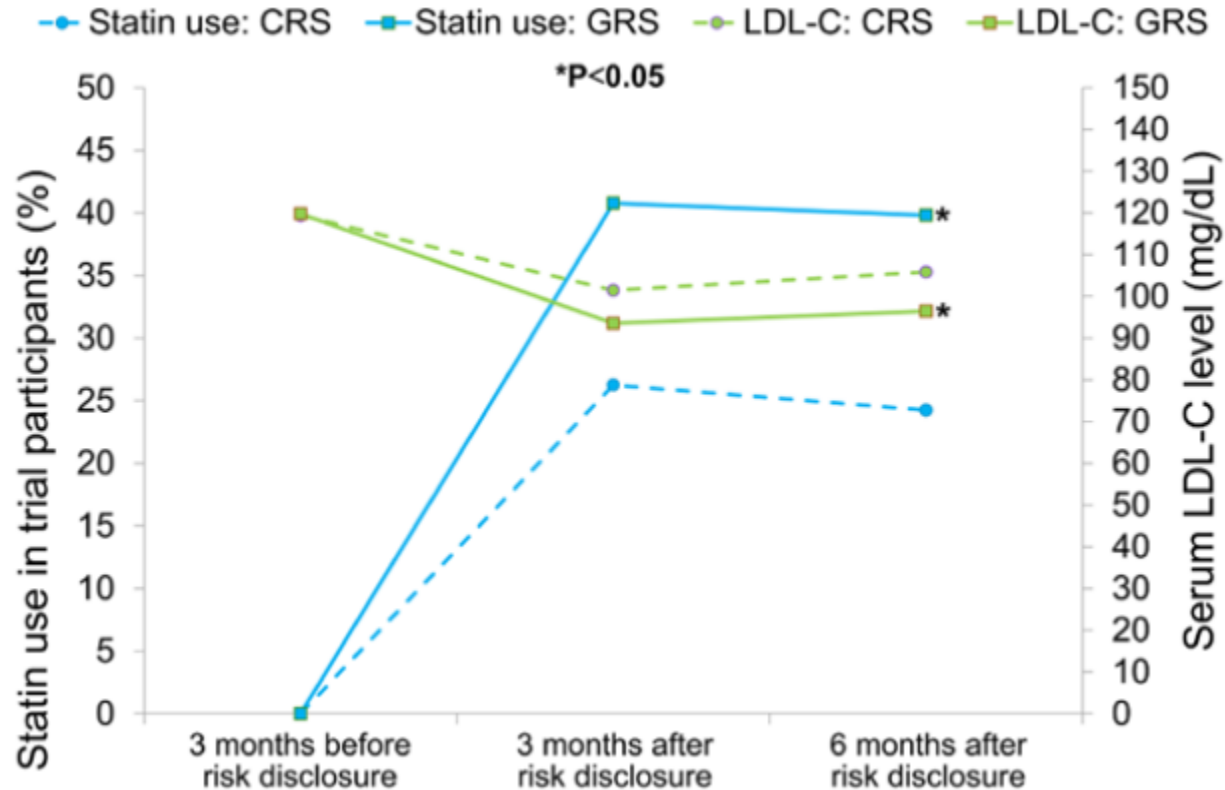
“We propose that it is time to contemplate the inclusion of polygenic risk prediction in clinical care...”



Effect of Disclosing Genetic Risk for Coronary Heart Disease on Information Seeking and Sharing

The MI-GENES Study (Myocardial Infarction Genes)

Sherry-Ann N. Brown, MD, PhD; Hayan Jouni, MD; Tariq S. Marroush, MD; Iftikhar J. Kullo, MD



Statin use significantly higher in patients given genetic risk score than conventional risk score

GxE Interaction and Correlation

- GxE: **SENSITIVITY** to E controlled by G
- rGE: **EXPOSURE** to E correlated with (“depends on”) G

Lots of good plant and animal models for both

DIATHESIS-STRESS MODEL IN DEPRESSION

Depression = Diathesis + Stress + Diathesis * Stress

(Predisposition, Vulnerability) (Disruption of psychological equilibrium)

Hypothesised contribution to risk



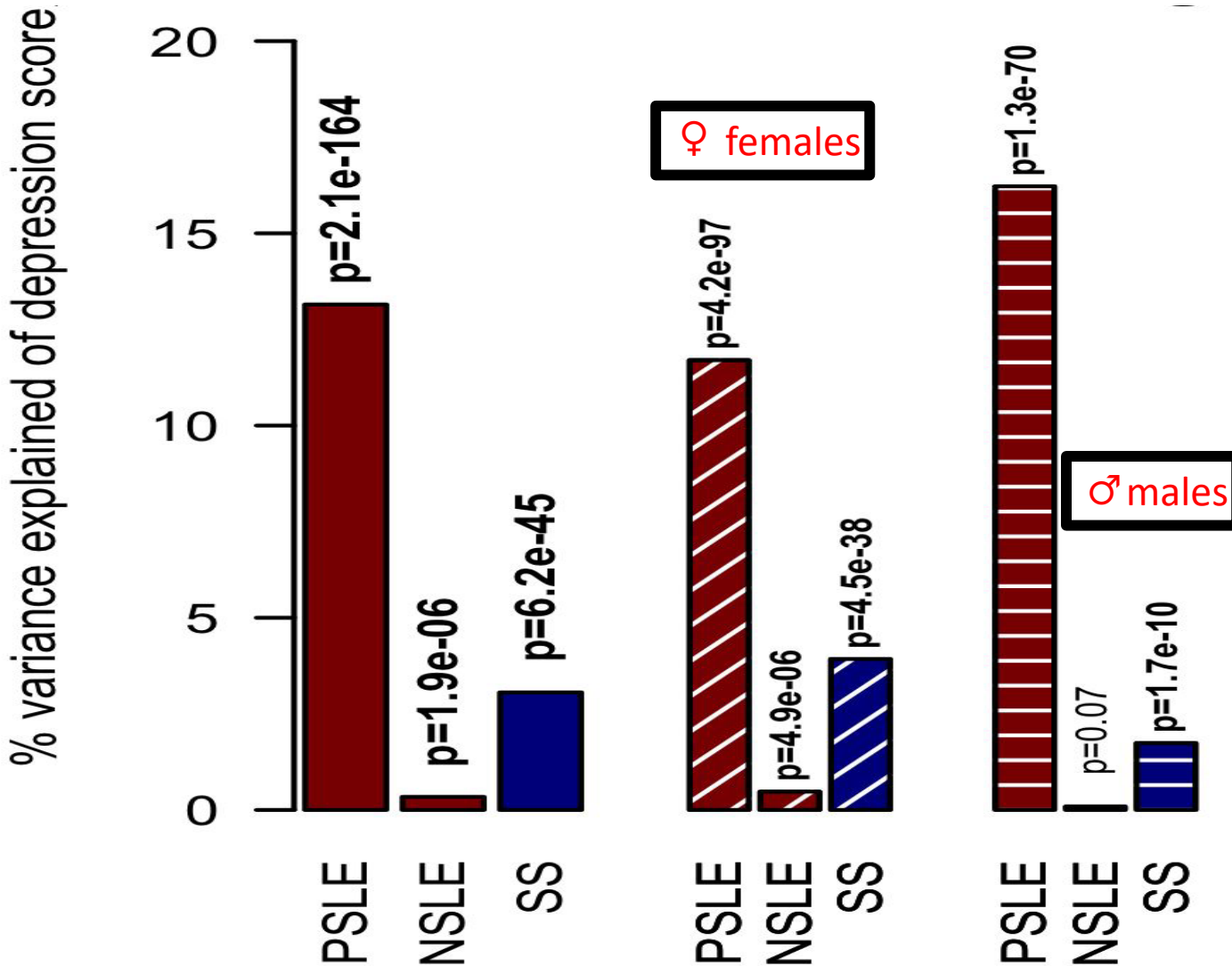
Our model:

Depression = PRS + PSLE + NSLE + SS + PRS * PSLE + PRS*NSLE + PRS*SS

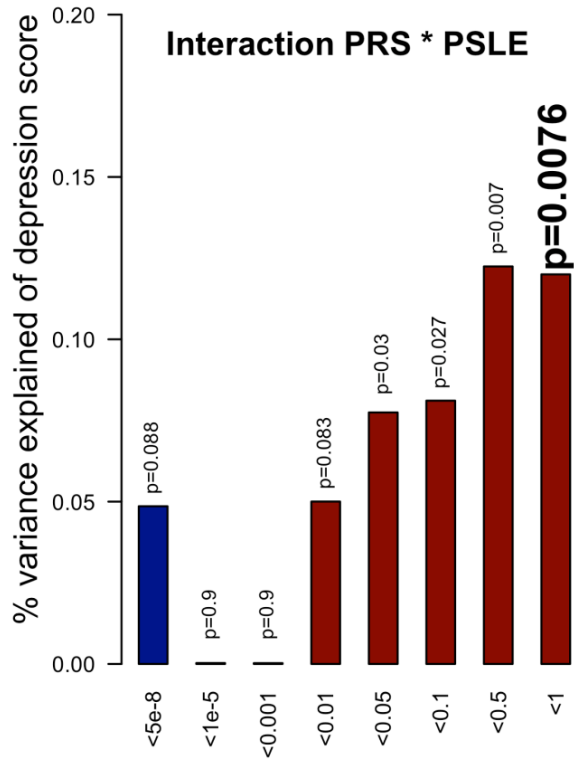
- **PRS**: Polygenic Risk Score
- **PSLE**: Personal Stressful Life Event, **NSLE**: Network Stressful life event, **SS** Social Support
- Familial relatedness modelled using a kinship matrix or a genetic relatedness matrix calculated from SNPs.
- All analyses controlled for age, age², sex, age*sex and age²*sex interactions, study, array, and the first four genetic principal components
- Scores regressed for covariates to avoid false positive interactions

MAIN EFFECTS OF STRESS SCORES

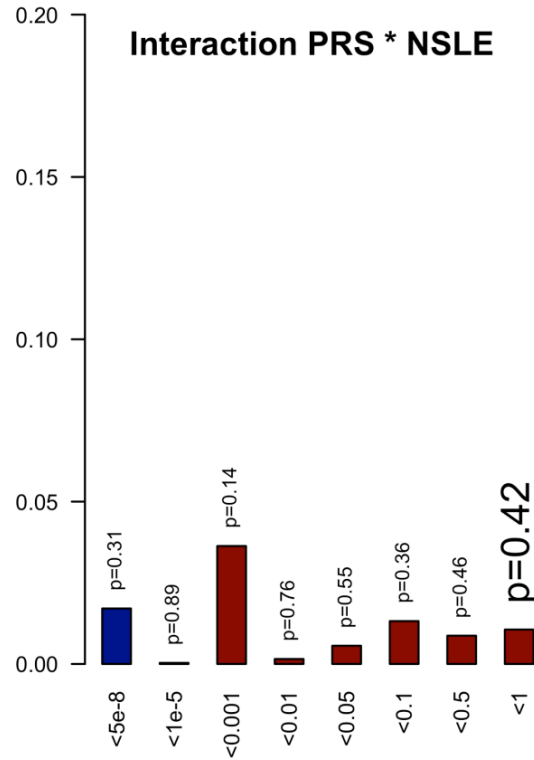
- Depression score associated with Personal (PSLE) and network (NSLE) stressful life events and perceived social support (SS). Note sex differences.



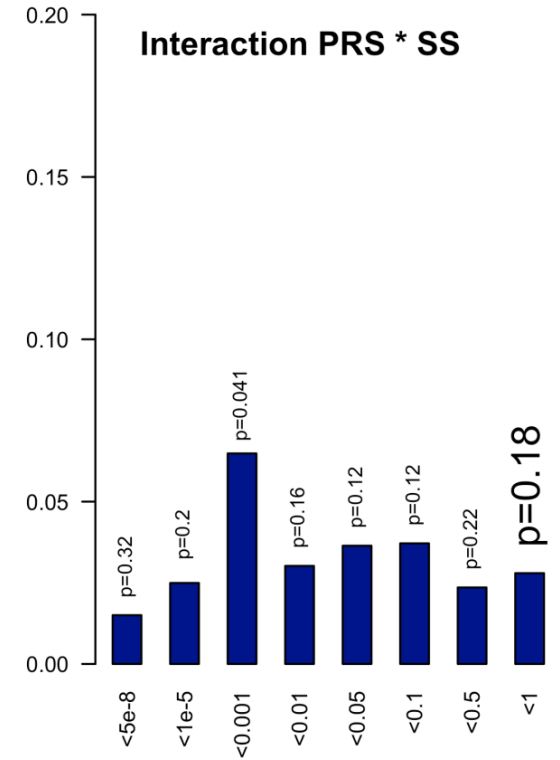
TEST OF INTERACTION



Significant interaction



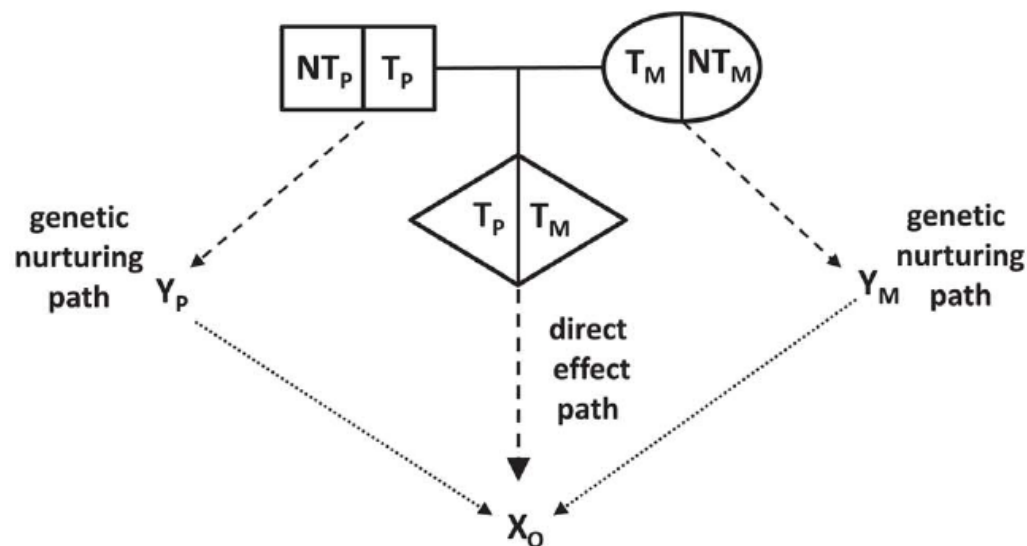
No interaction



No interaction

The nature of nurture: Effects of parental genotypes

Augustine KongKari Stefansson



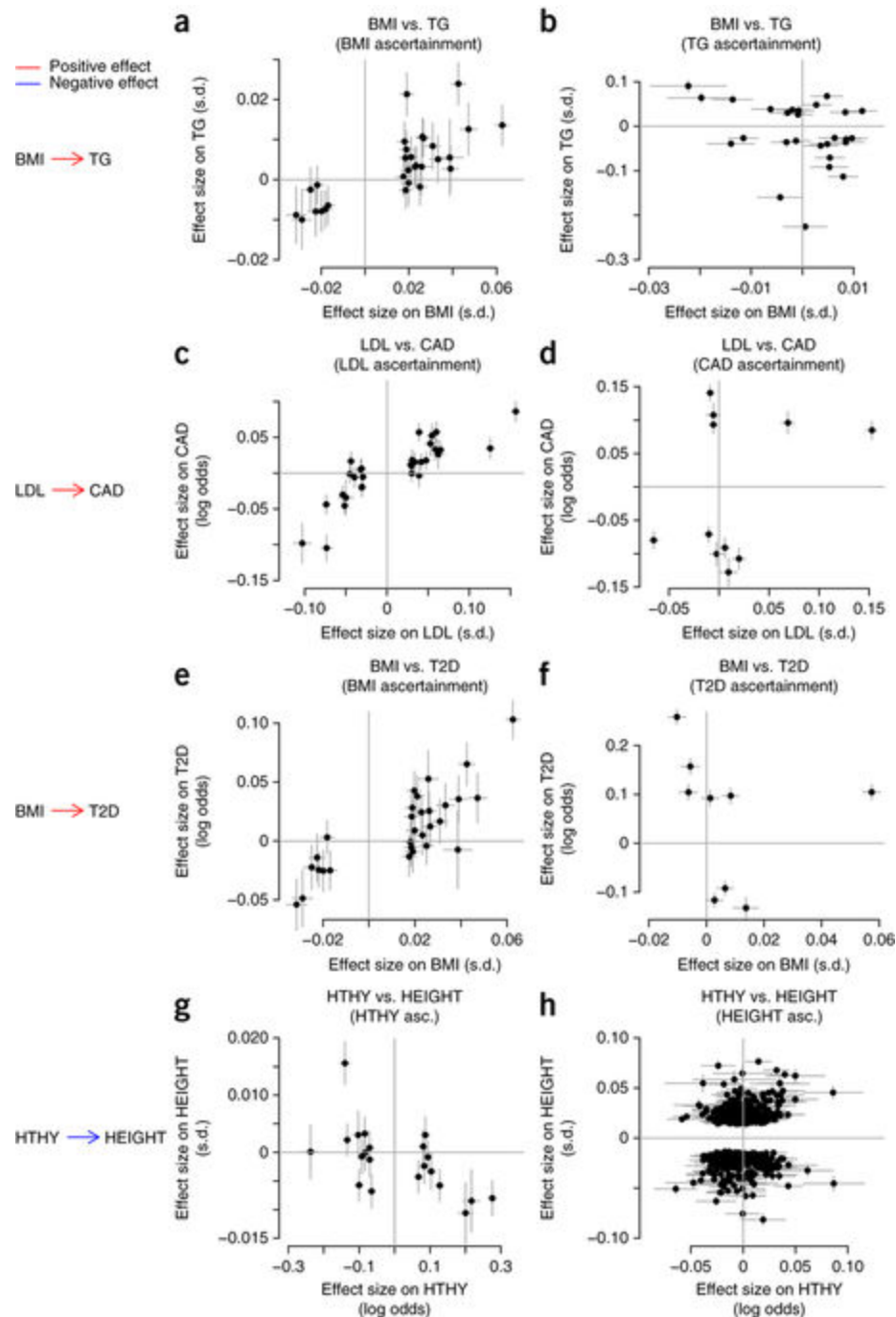
Nontransmitted alleles can affect a child through their impacts on the parents and other relatives, a phenomenon we call “genetic nurture.” Using results from a meta-analysis of educational attainment, we find that the polygenic score computed for the nontransmitted alleles of 21,637 probands with at least one parent genotyped has an estimated effect on the educational attainment of the proband that is 29.9% ($P = 1.6 \times 10^{-14}$) of that of the transmitted polygenic score.

Detection and interpretation of shared genetic influences on 42 human traits

Joseph K Pickrell, Tomaz Berisa, Jimmy Z Liu, Laure Séguirel, Joyce Y Tung & David A Hinds.

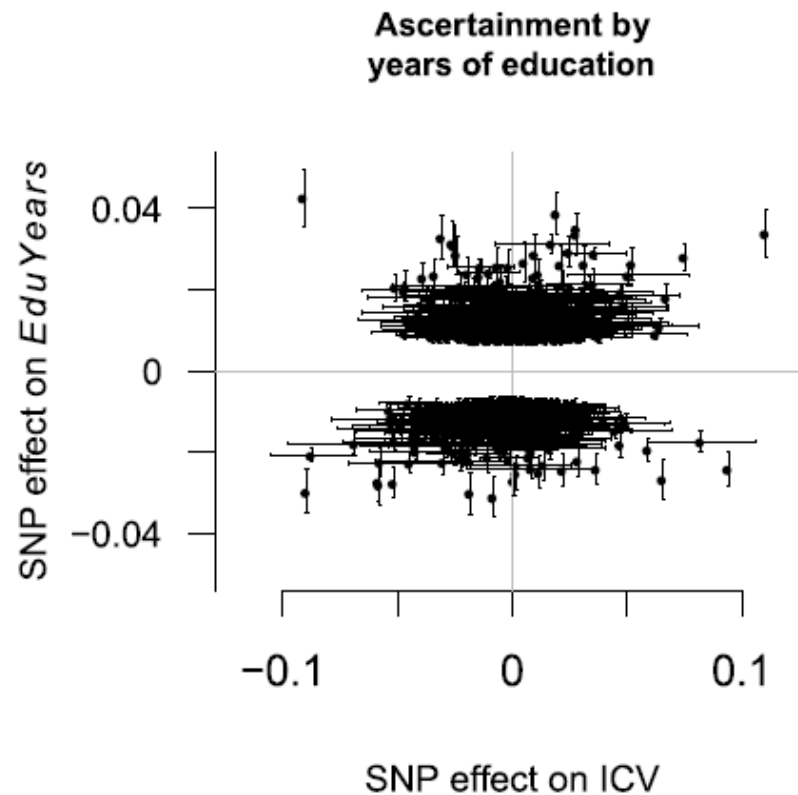
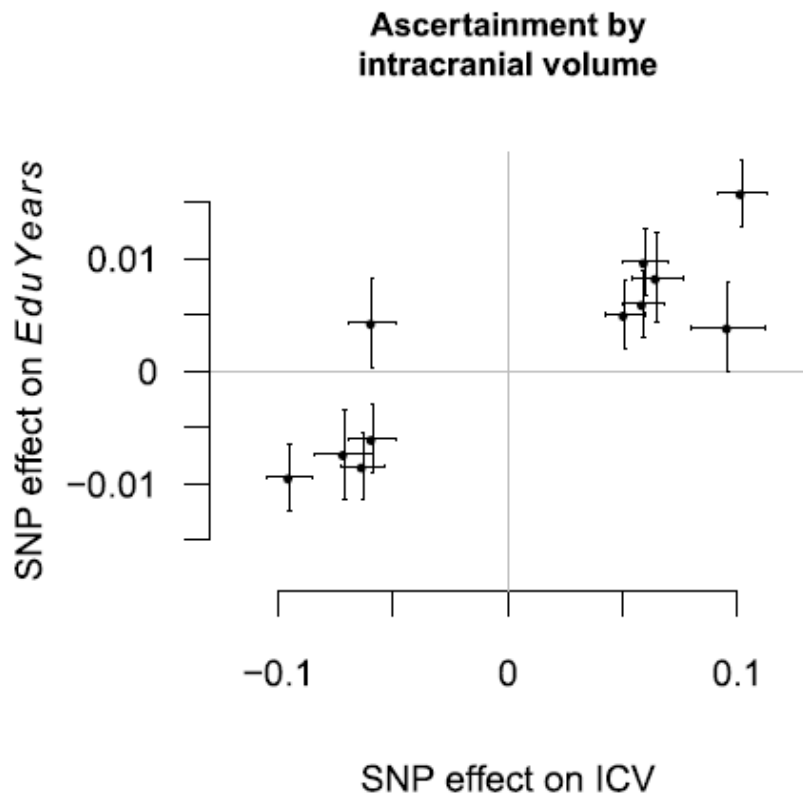
Nature Genetics 48; 709–717, 2016

Powerful GWAS for traits A and B can help determine direction of causation

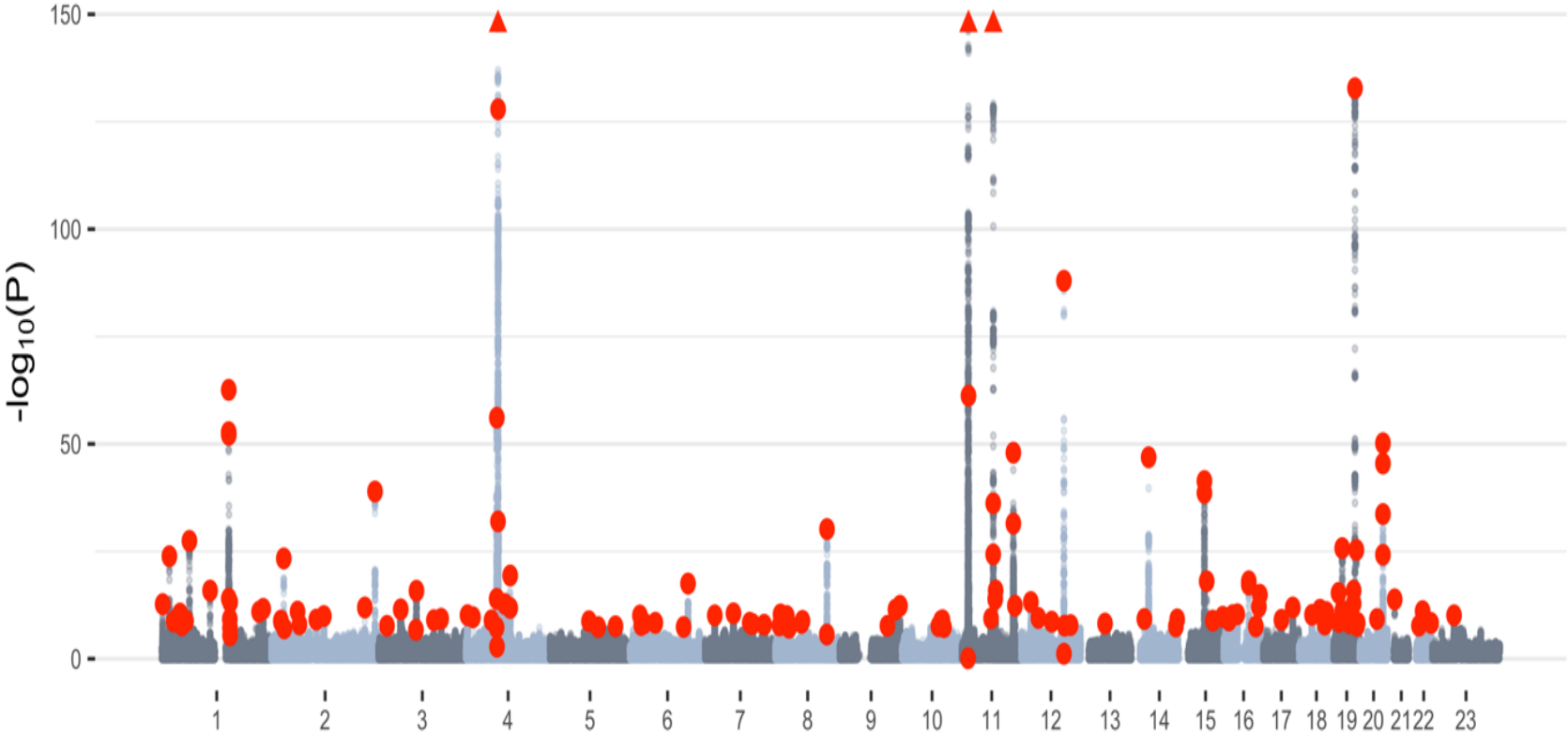


The causal influence of brain size on human intelligence: Evidence from within-family phenotypic associations and GWAS modeling

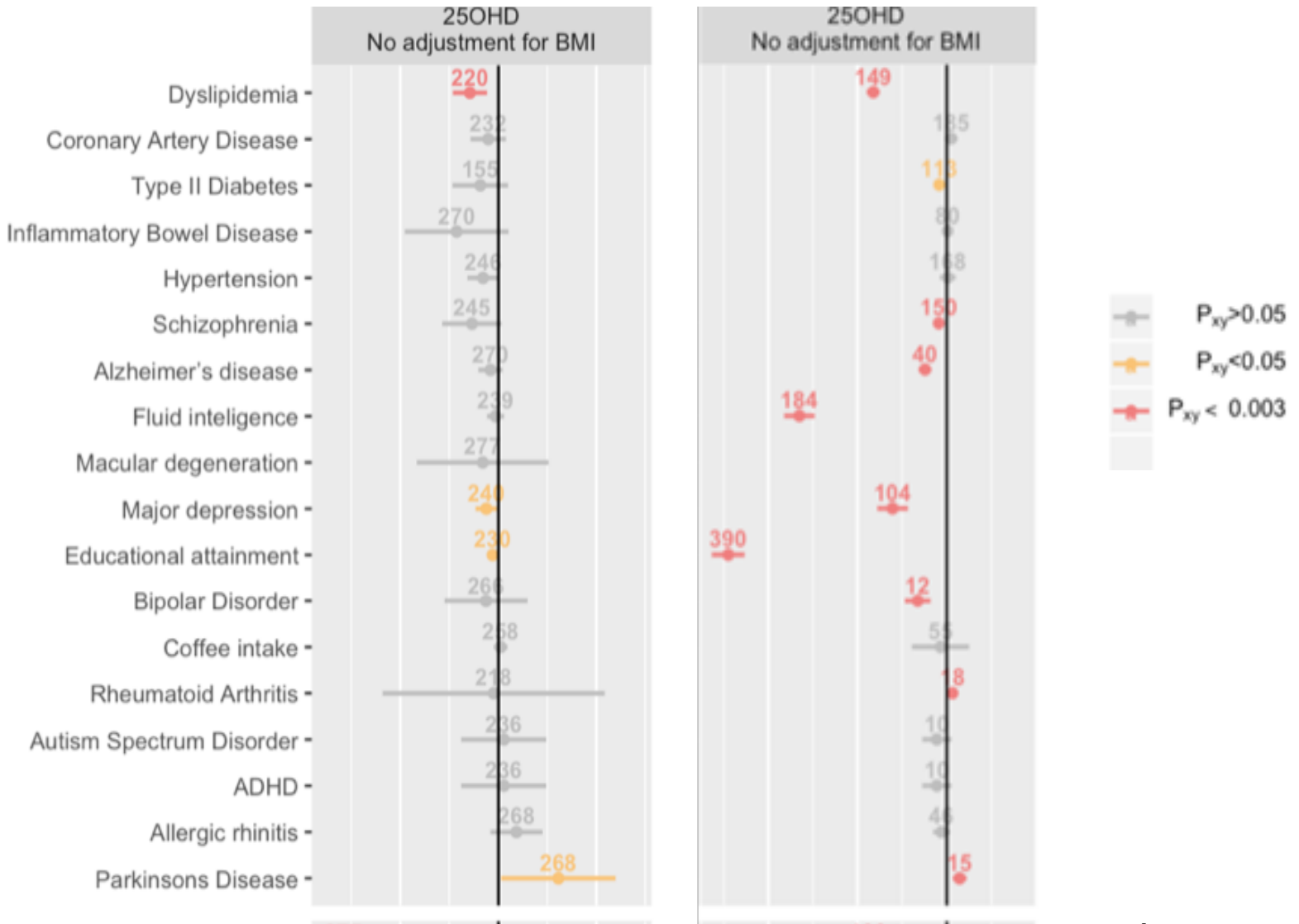
James J. Lee^{a,*}, Matt McGue^a, William G. Iacono^a, Andrew M. Michael^{b,c}, Christopher F. Chabris^b



Manhattan plot of the 25OHD (vitamin D) GWAS in the UK Biobank: n=417,580, 143 loci



Bidirectional Generalized Summary data level Mendelian Randomization (GSMR) between 25 hydroxyvitamin D concentrations and selected phenotypes



Inference on heritability and selection from WGS data

Heritability estimates from different studies:

Source	Height	BMI	Year	EA	SCZ
Pedigree studies	~ 0.8	~ 0.4 - 0.6	< 2010	~0.4	~0.7
Genotyped SNPs	~ 0.45 (0.03)	~ 0.16 (0.03)	2010+	~0.15	~0.25
Imputed SNPs	~ 0.56 (0.02)	~ 0.27 (0.02)	2015+	~0.2	~0.3
Whole Genome	???			???	???

Can we recover the 'still missing' heritability using WGS data?

Investigation using TOPMed WGS



TOPMed overview

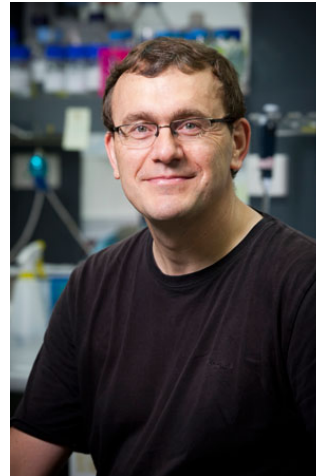
NIH project, whole genome sequencing of entire cohorts in the US (~30 cohorts)

- ~40% Europeans
- ~30% African americans
- ~16% latin@s
- ~9% Asians
- ~4% unknown

Median sequencing depth 30x

Multiple data freezes:

- 54k samples for freeze 5
- 140k samples for freeze 8

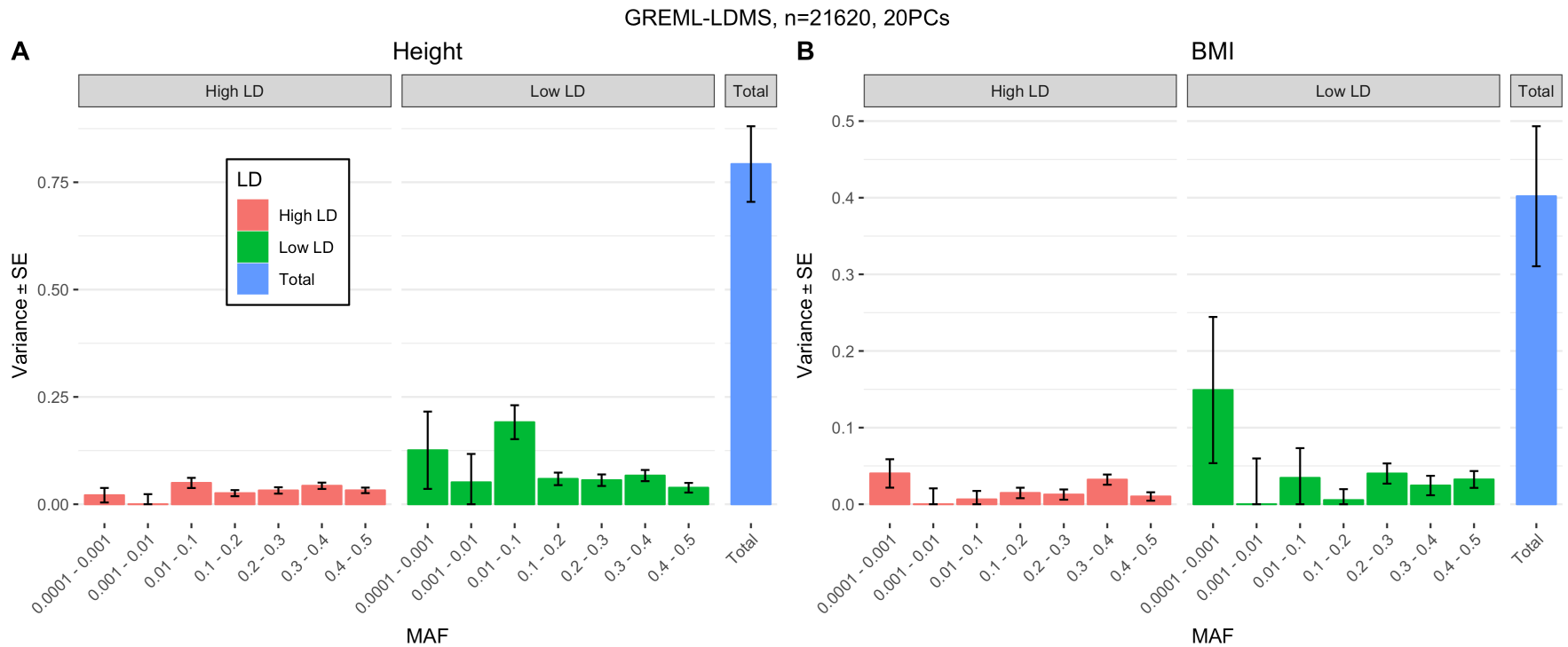


Peter Visscher



Pierrick Wainschtein

GREML-LDMS using WGS data: recovery of pedigree heritability?



Estimates using 20PCs as fixed effects:

- Height:

$$h_{WGS}^2 = 0.79 (0.09)$$

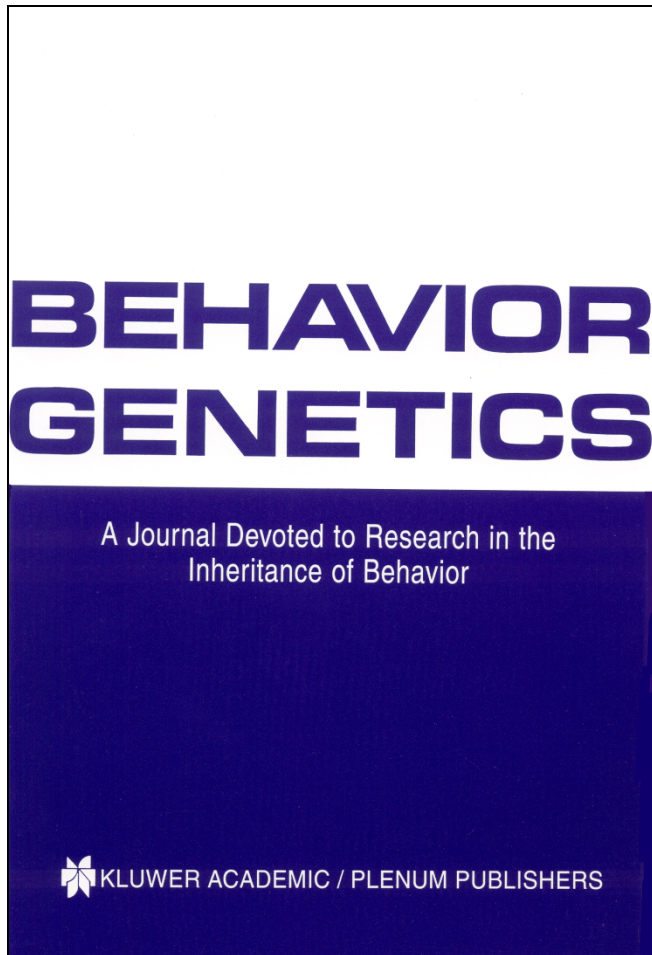
- BMI:

$$h_{WGS}^2 = 0.40 (0.09)$$

Estimates close to pedigree estimates

Large role for low LD and low MAF variants

We also run two journals (1)



- Editor: John Hewitt
- Editorial assistant
Christina Hewitt
- Publisher: Kluwer
/Plenum
- Fully online
- <http://www.bga.org>

twin research and human genetics

The official journal of the International Society for Twin Studies

Covering all areas of
human genetics
with an emphasis on
twin studies, genetic
epidemiology, psychiatric
and behavioral genetics.
And research on
multiple births in the
fields of epidemiology,
genetics, and pathology,
fetal pathology, obstetrics
and pediatrics.

Volume **8** Number **1**

February 2005

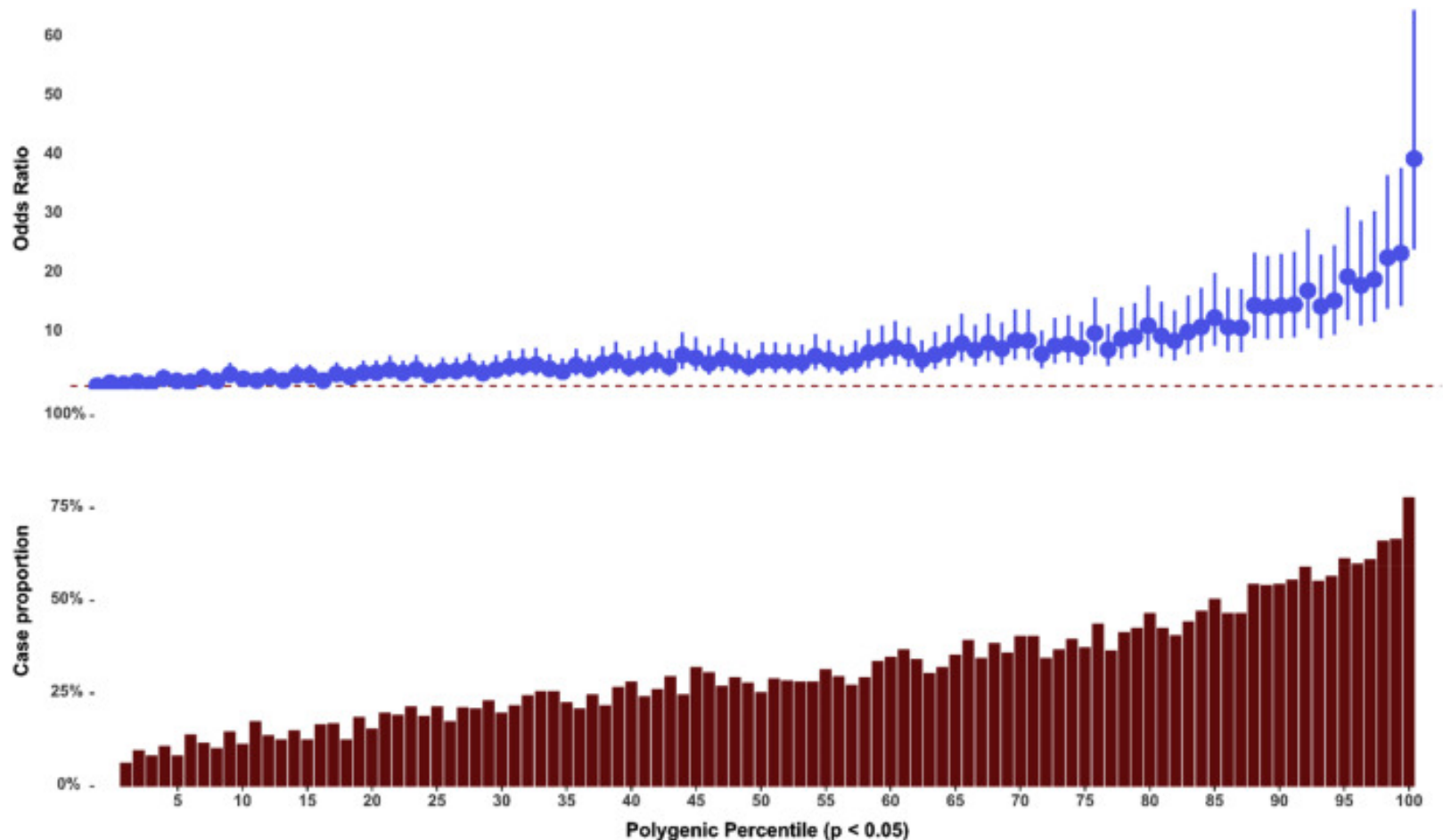
Editor
Nicholas G. Martin

- Editor: Nick Martin
- Publisher:
Cambridge
University Press
- Fully online
- Fast turnaround
- **First submission
free to workshop
participants!!!!!!😊**

Schizophrenia GWAS: polygenic centile analysis

Independent discovery sample (29,415 cases and 40,101 controls) used to predict to CLOZUK target sample (11,260 cases and 24,542 controls)

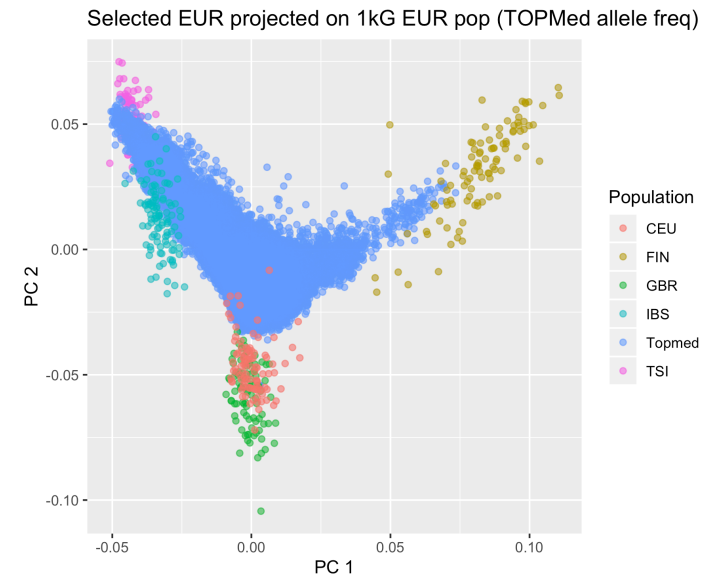
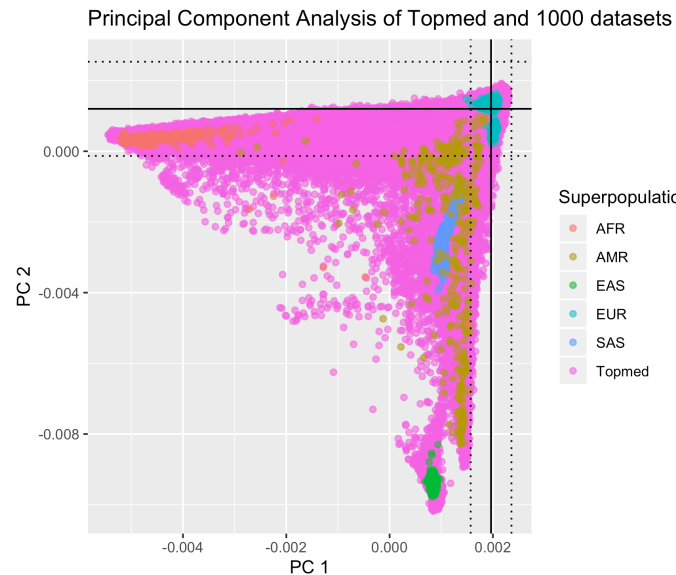
Risk of schizophrenia in top 1% is ~40-fold higher than in bottom 1%



European ancestry selection

Selection of European ancestry samples

1. Using HM3 SNPs (1.2M)
2. From 1kG population references
3. Selected samples within 6sd of EUR pop



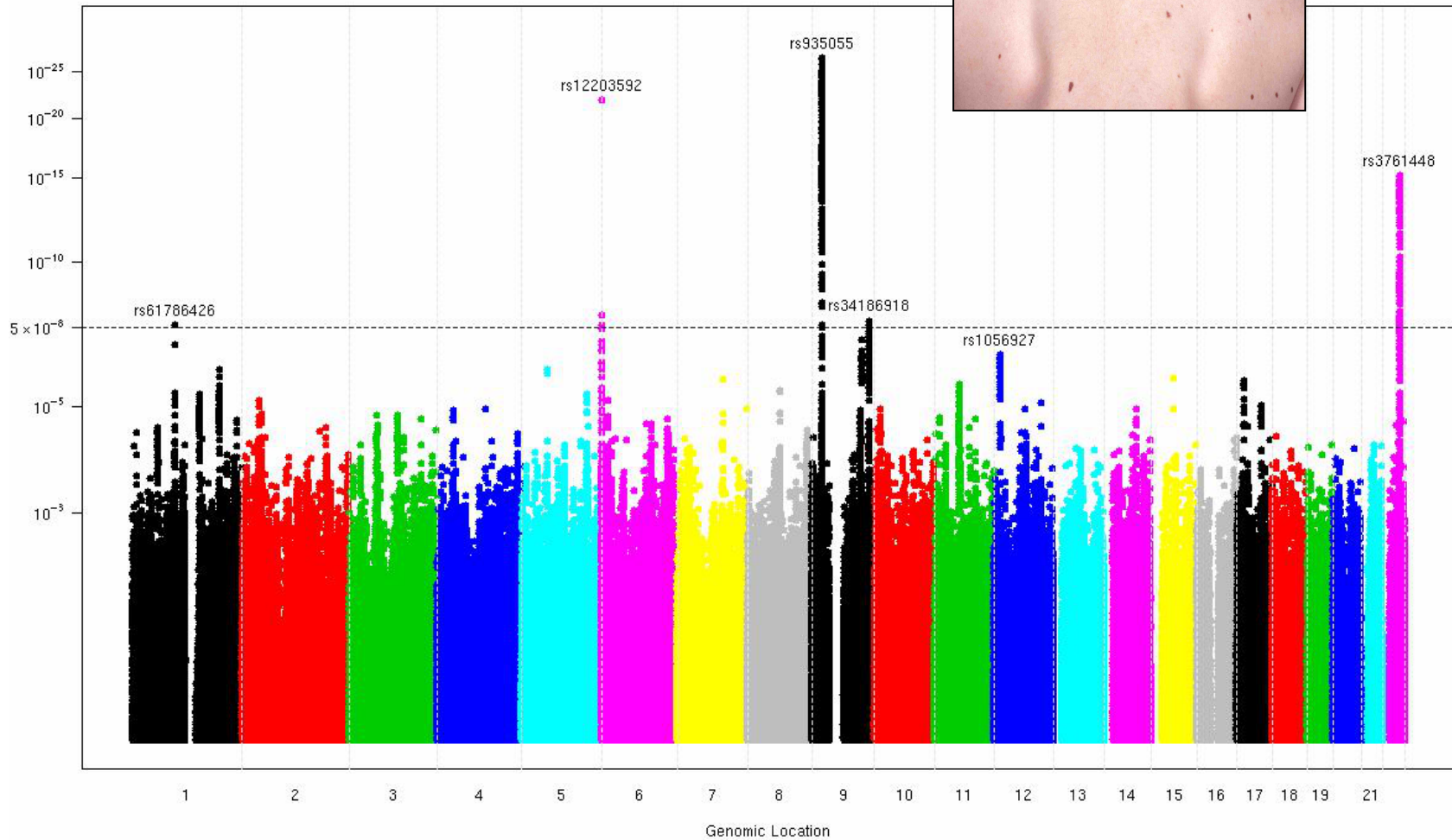
Dataset

Analysis parameters:

- 21,620 unrelated individuals from European ancestry
- Genomic Relationship Matrices (GRMs) fitted:
 - 1 GRM from HM3 SNPs
 - 7 GRMs from all variants split by MAF
 - 14 GRM split by MAF and LD (7 AF bins * 2 LD bins)
- MAF threshold of 0.0001 (each variant seen > 2x)
- 47.1M variants
- Phenotypes standardized $N(0, 1)$ within each cohort by sex

RINT = rank based inverse normal transformation

GWAS for moliness(11 studies, 23,000 subjects, 5 hits)



David Duffy, Gu Zhu

The support of human genetic evidence for approved drug indications

Matthew R Nelson¹, Hannah Tipney², Jeffery L Painter¹, Judong Shen¹, Paola Nicoletti³, Yufeng Shen^{3,4}, Aris Floratos^{3,4}, Pak Chung Sham^{5,6}, Mulin Jun Li^{6,7}, Junwen Wang^{6,7}, Lon R Cardon⁸, John C Whittaker² & Philippe Sanseau²

Table 1 The relative value of genetic support for the probability that a target-indication pair progresses along the drug development pipeline, based on historical drug trial information

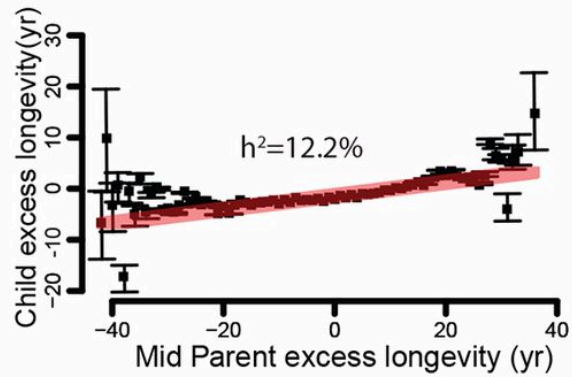
Progression	$\rho(\text{progress} \text{genetic support})/(\text{progress} \text{no genetic support})$		
	GWASdb and OMIM	GWASdb	OMIM
Phase I to phase II	1.2 (1.1–1.3)	1.2 (1.1–1.3)	1.2 (1.1–1.3)
Phase II to phase III	1.5 (1.3–1.7)	1.4 (1.2–1.7)	1.6 (1.3–1.9)
Phase III to approval	1.1 (1.0–1.2)	1.0 (0.8–1.2)	1.1 (0.9–1.3)
Phase I to phase III	1.8 (1.5–2.1)	1.8 (1.4–2.1)	1.9 (1.5–2.3)
Phase I to approval	2.0 (1.6–2.4)	1.8 (1.3–2.3)	2.2 (1.6–2.8)

“We estimate that selecting genetically supported targets could double the success rate in clinical development. Therefore, using the growing wealth of human genetic data to select the best targets and indications should have a measurable impact on the successful development of new drugs.”

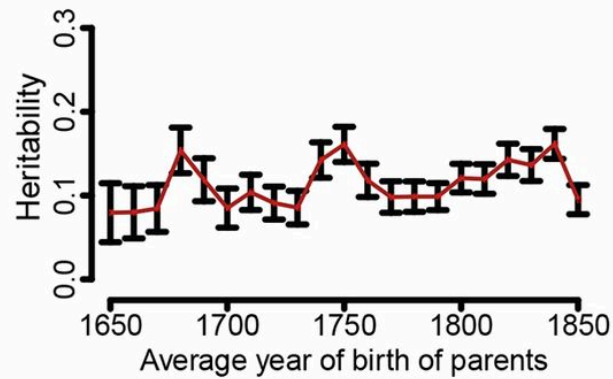
Quantitative analysis of population-scale family trees with millions of relatives

Joanna Kaplanis *et al.* *Science* 01 Mar 2018:

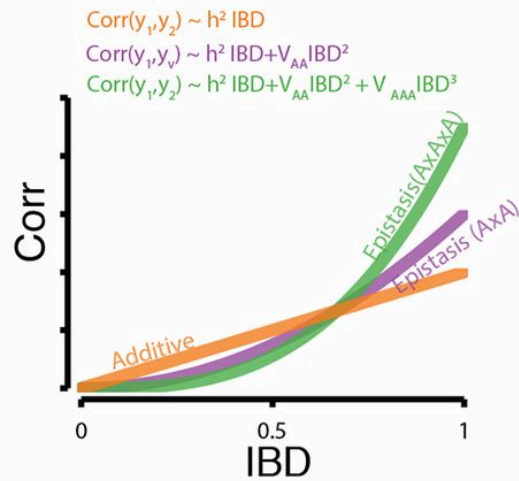
A



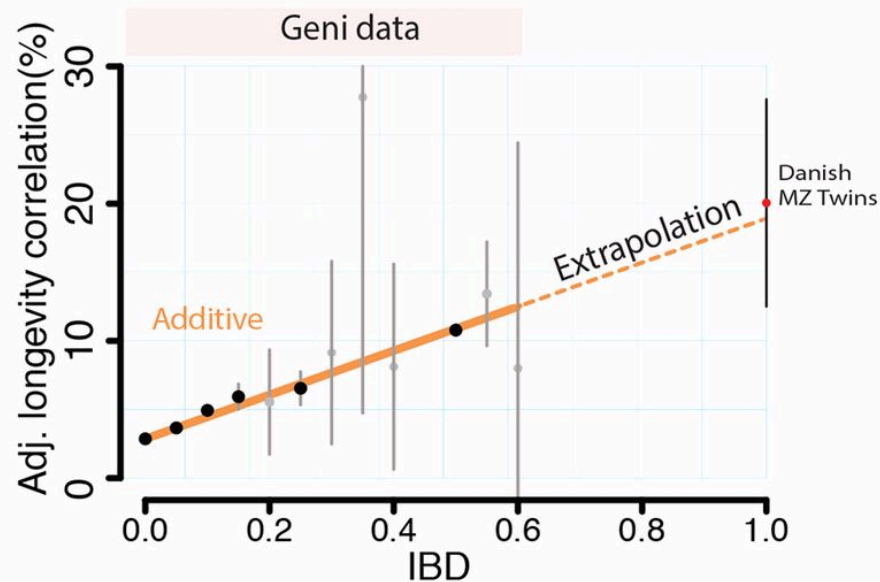
B



C

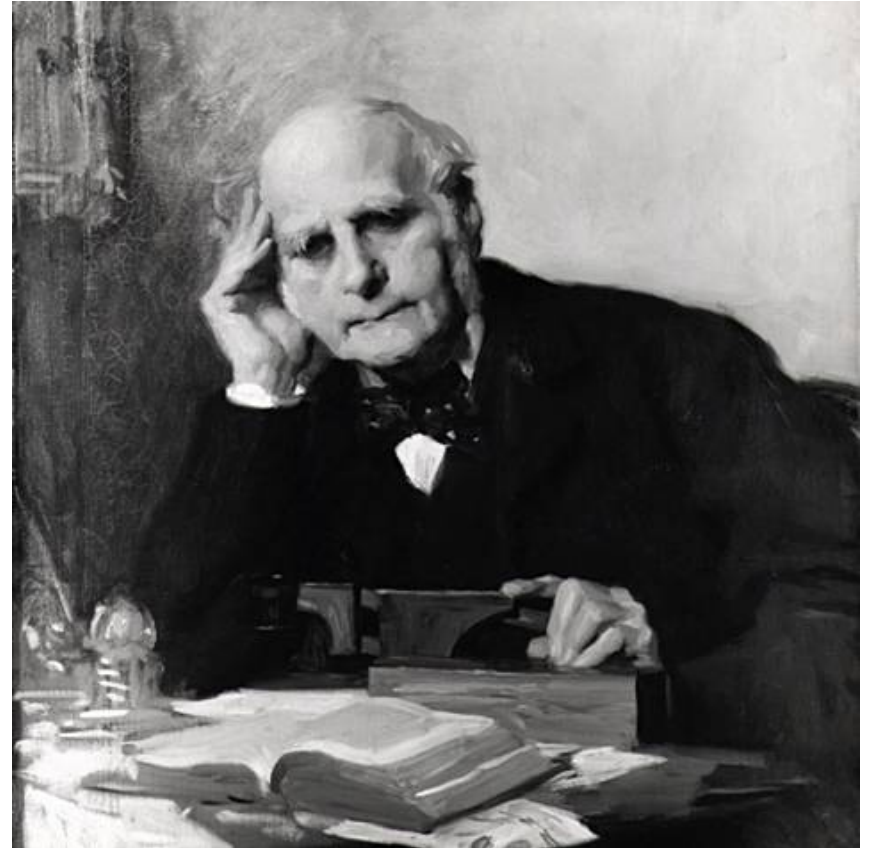
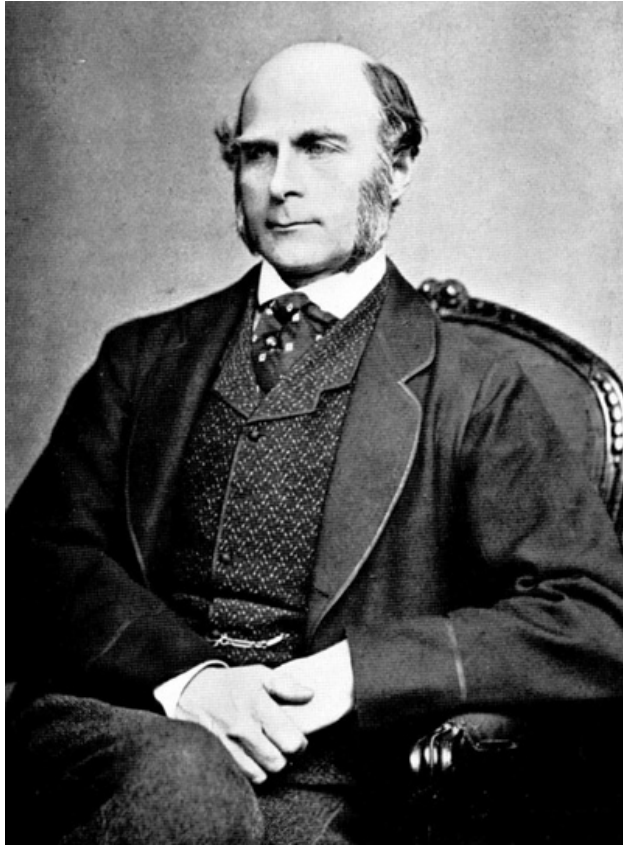


D



“HEREDITY”

Francis Galton (1822-1911)

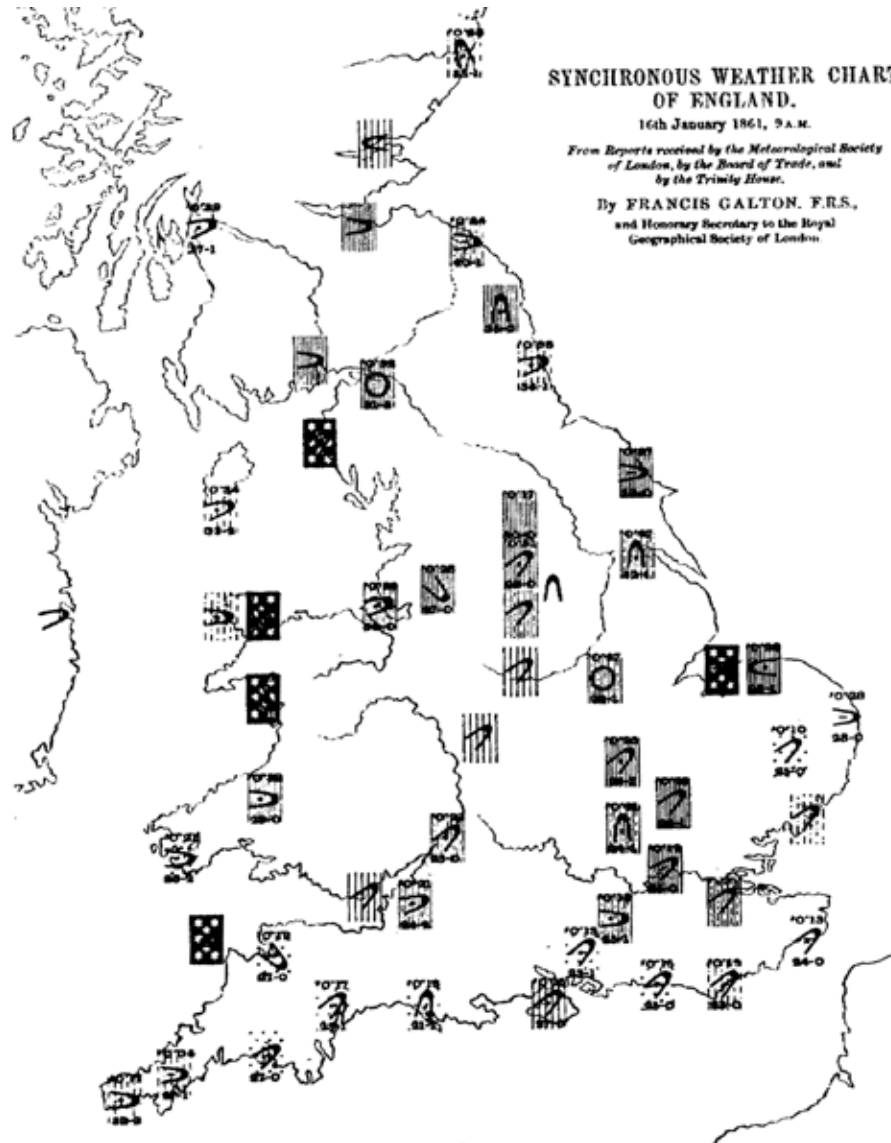


1869: Hereditary Genius

1883: Inquiries into Human Faculty and its Development

1884-5: Anthropometric Laboratory at "National Health Exhibition"

Galton's Other Work e.g. Meteorology



Hereditary Genius (1869, p 317)

	Judges, p. 61.	Statesmen, p. 109.	Commanders, p. 148.	Literary, p. 171.	Scientific, p. 195.	Poets, p. 227.	Artists, pp. 238 and 249.	Divines, p. 275.	Illustrious and Eminent Men of all Classes.		
	B.	B.	B.	B.	B.	B.	B.	B.	B.	C.	D.
Father	26	33	47	48	26	20	32	28	31	100	31
Brother	35	39	50	42	47	40	50	36	41	150	27
Son	36	49	31	51	60	45	89	40	48	100	48
Grandfather	15	28	16	24	14	5	7	20	17	200	8
Uncle	18	18	8	24	16	5	14	40	18	400	5
Nephew	19	18	35	24	23	50	18	4	22	400	5
Grandson	19	10	12	9	14	5	18	16	14	200	7
Great-grandfather .	2	8	8	3	0	0	0	4	3	400	1
Great-uncle	4	5	8	6	5	5	7	4	5	800	1
First cousin	11	21	20	18	16	0	1	8	13	800	2
Great-nephew	17	5	8	6	16	10	0	0	10	800	1
Great-grandson . . .	6	0	0	3	7	0	0	0	3	400	1
All more remote . .	14	37	44	15	23	5	18	16	31	?	...

Galton's Anthropometric Laboratory (1884-1885)

ANTHROPOMETRIC LABORATORY

**For the measurement in various
ways of Human Form and Faculty.**

Entered from the Science Collection of the S. Kensington Museum.

This laboratory is established by Mr. Francis Galton for the following purposes:—

1. For the use of those who desire to be accurately measured in many ways, either to obtain timely warning of remediable faults in development, or to learn their powers.
2. For keeping a methodical register of the principal measurements of each person, of which he may at any future time obtain a copy under reasonable restrictions. His initials and date of birth will be entered in the register, but not his name. The names are indexed in a separate book.
3. For supplying information on the methods, practice, and uses of human measurement.
4. For anthropometric experiment and research, and for obtaining data for statistical discussion.

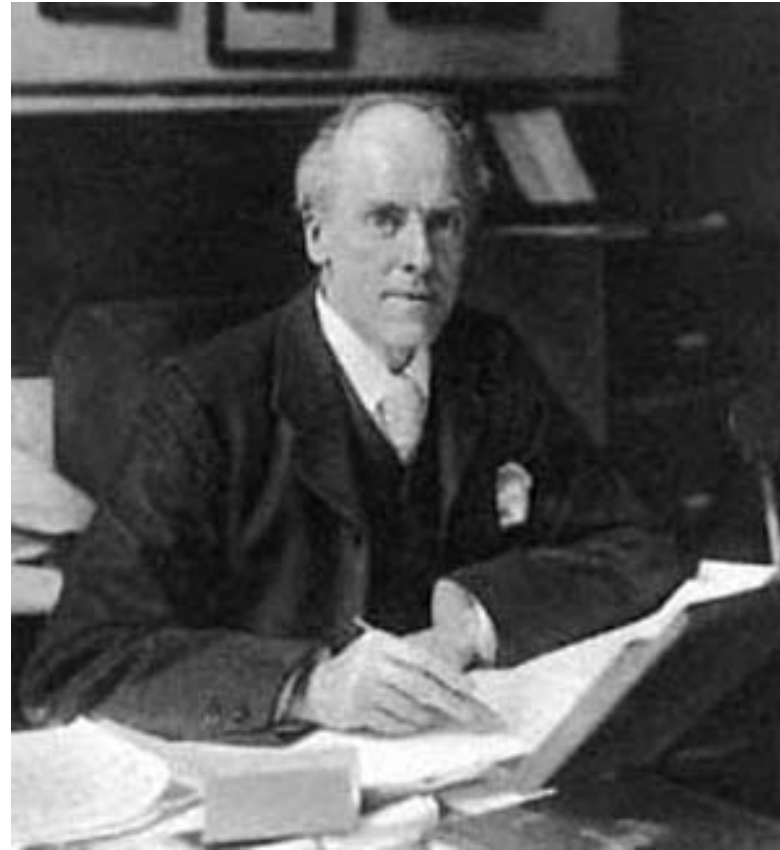
Charges for making the principal measurements:
THREEPENCE each, to those who are already on the Register.
FOURPENCE each, to those who are not:— one page of the Register will thenceforward be assigned to them, and a few extra measurements will be made, chiefly for future identification.

The Superintendent is charged with the control of the laboratory and with determining in each case, which, if any, of the extra measurements may be made, and under what conditions.



Francis Galton's First Anthropometric Laboratory at the International Health Exhibition, South Kensington, 1884-5.

Karl Pearson (1857-1936)



1903: On the Laws of Inheritance in Man: I Physical Characteristics (with Alice Lee)

1904: II Mental and Moral Characteristics

1914: The Life, Letters and Labours of Francis Galton

FAMILY MEASUREMENTS.

Professor KARL PEARSON, of University College, London, would esteem it a great favour if any persons in a position to do so, would assist him by making one set (or if possible several sets) of anthropometric measurements on their own family, or on families with whom they are acquainted. The measurements are to be made use of for testing theories of heredity, no names, except that of the recorder, are required, but the Professor trusts to the *bona fides* of each recorder to send only correct results.

Each family should consist of a father, mother, and at least one son or daughter, not necessarily the eldest. The sons or daughters are to be at least 18 years of age, and measurements are to be made on not more than two sons and two daughters of the same family. If more than two sons or two daughters are easily accessible, then not the tallest but the eldest of those accessible should be selected.

To be of real service the whole series ought to contain 1000—2000 families, and therefore the Professor will be only too grateful if anyone will undertake several families for him.

Copies of this paper, together with cards for recording data, may be obtained from

or from the above-named Professor.

Pearson and Lee's diagram for measurement of "span" (finger-tip to finger-tip distance)

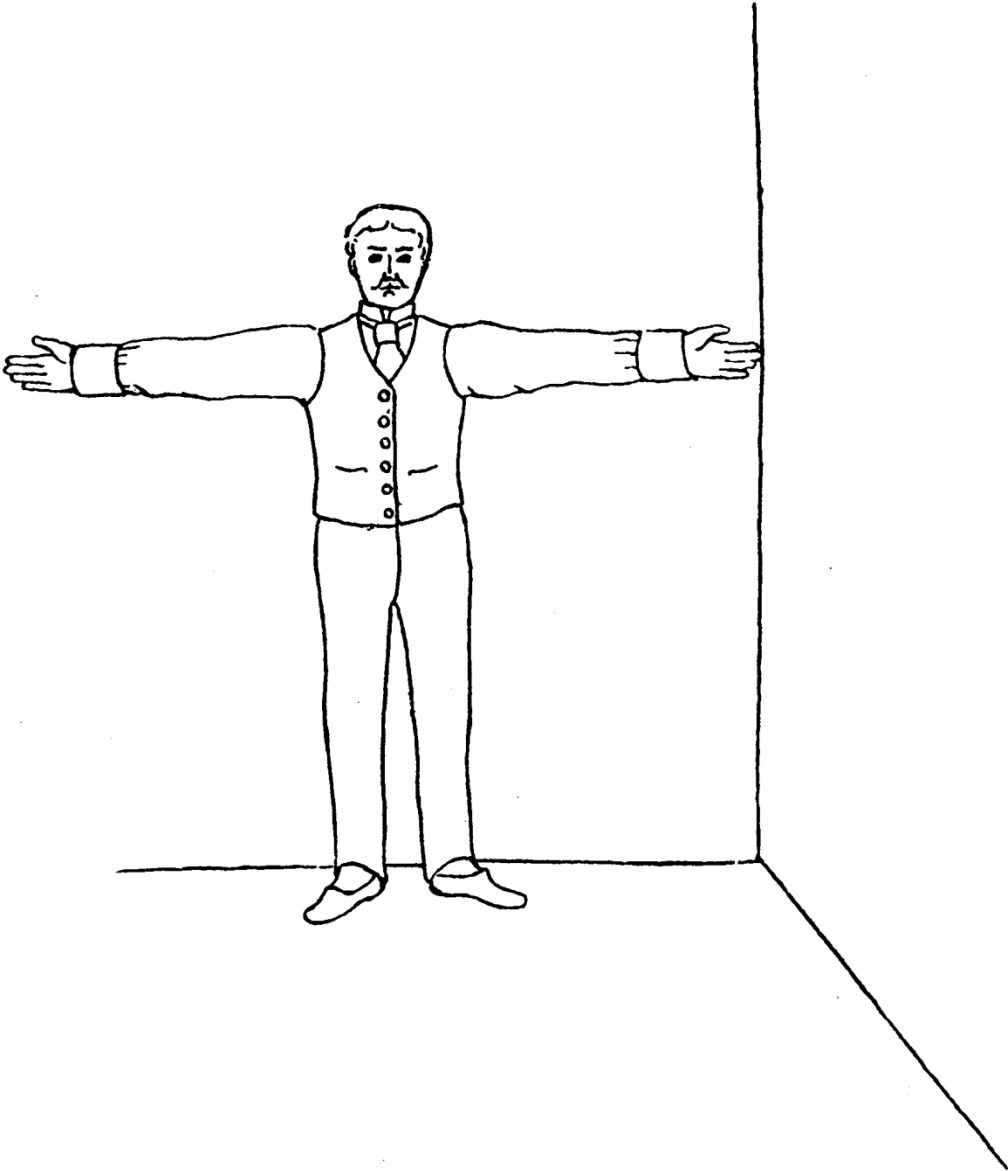


TABLE IV.

Coefficients of Heredity. Parents and Offspring.

Character	Father and		Mother and	
	Son	Daughter	Son	Daughter
Stature	$\cdot514 \pm \cdot015$	$\cdot510 \pm \cdot013$	$\cdot494 \pm \cdot016$	$\cdot507 \pm \cdot014$
Span	$\cdot454 \pm \cdot016$	$\cdot454 \pm \cdot014$	$\cdot457 \pm \cdot016$	$\cdot452 \pm \cdot015$
Forearm	$\cdot421 \pm \cdot017$	$\cdot422 \pm \cdot015$	$\cdot406 \pm \cdot017$	$\cdot421 \pm \cdot015$

From Pearson and Lee (1903) p.378

Correlation Coefficients for Direct Fraternal Heredity.

Character	Brother and Brother	Sister and Sister	Brother and Sister	Mean
Stature	$\cdot511 \pm \cdot028$	$\cdot537 \pm \cdot022$	$\cdot553 \pm \cdot013$	$\cdot534$
Span	$\cdot549 \pm \cdot026$	$\cdot555 \pm \cdot021$	$\cdot525 \pm \cdot013$	$\cdot543$
Forearm	$\cdot491 \pm \cdot029$	$\cdot507 \pm \cdot023$	$\cdot440 \pm \cdot015$	$\cdot479$
Mean	$\cdot517$	$\cdot533$	$\cdot506$	$\cdot519$
Eye Colour*	$\cdot517 \pm \cdot020$	$\cdot446 \pm \cdot023$	$\cdot462 \pm \cdot022$	$\cdot475$
Total mean	$\cdot517$	$\cdot511$	$\cdot495$	$\cdot508$

From Pearson and Lee (1903) p.387

Assortative Mating. Based on 1000 to 1050 Cases of Husband and Wife.

	Husband's Character	Wife's Character	Correlation and Probable Error	Symbol
Direct	Stature Span Forearm	Stature Span Forearm	$\cdot2804 \pm \cdot0189$ $\cdot1989 \pm \cdot0204$ $\cdot1977 \pm \cdot0205$	r_{12} r_{34} r_{56}

From Pearson and Lee (1903) p. 373

Modern Data

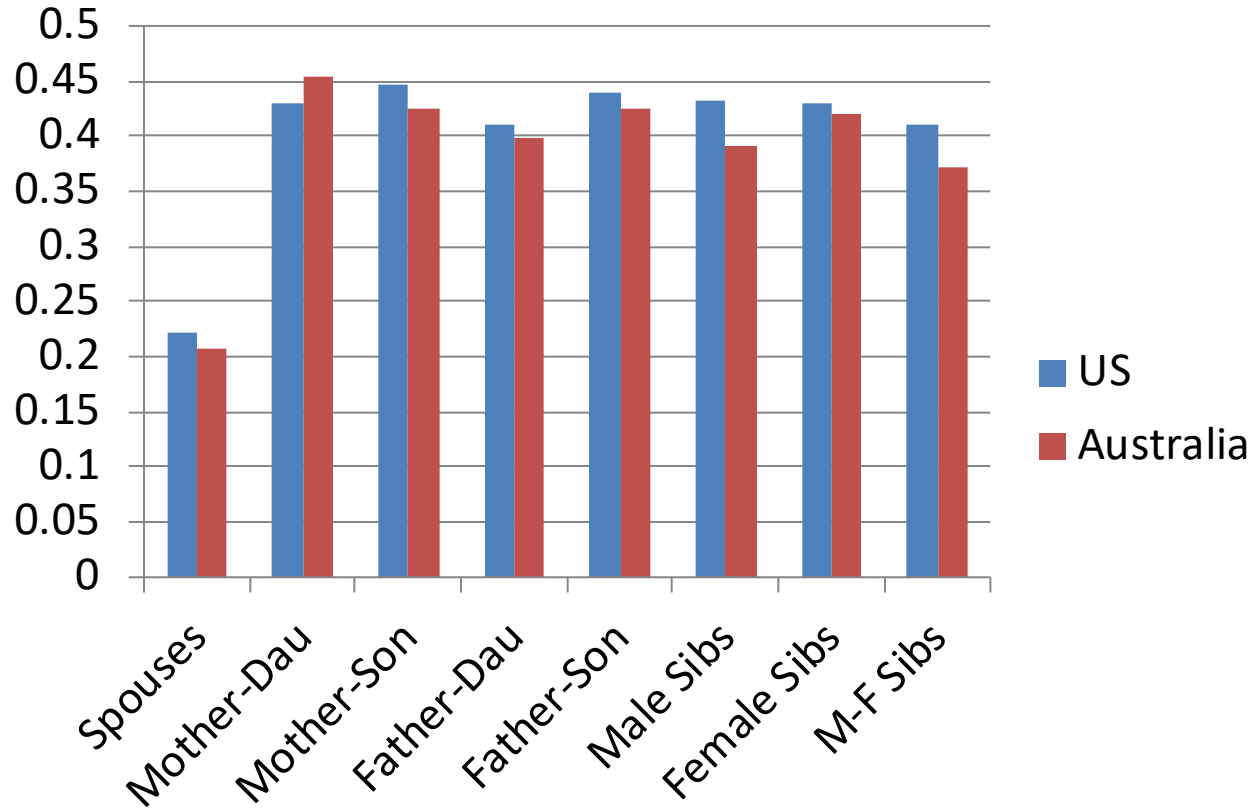
The Virginia 30,000
(N=29691)

The Australia 22,000
(N=20480)

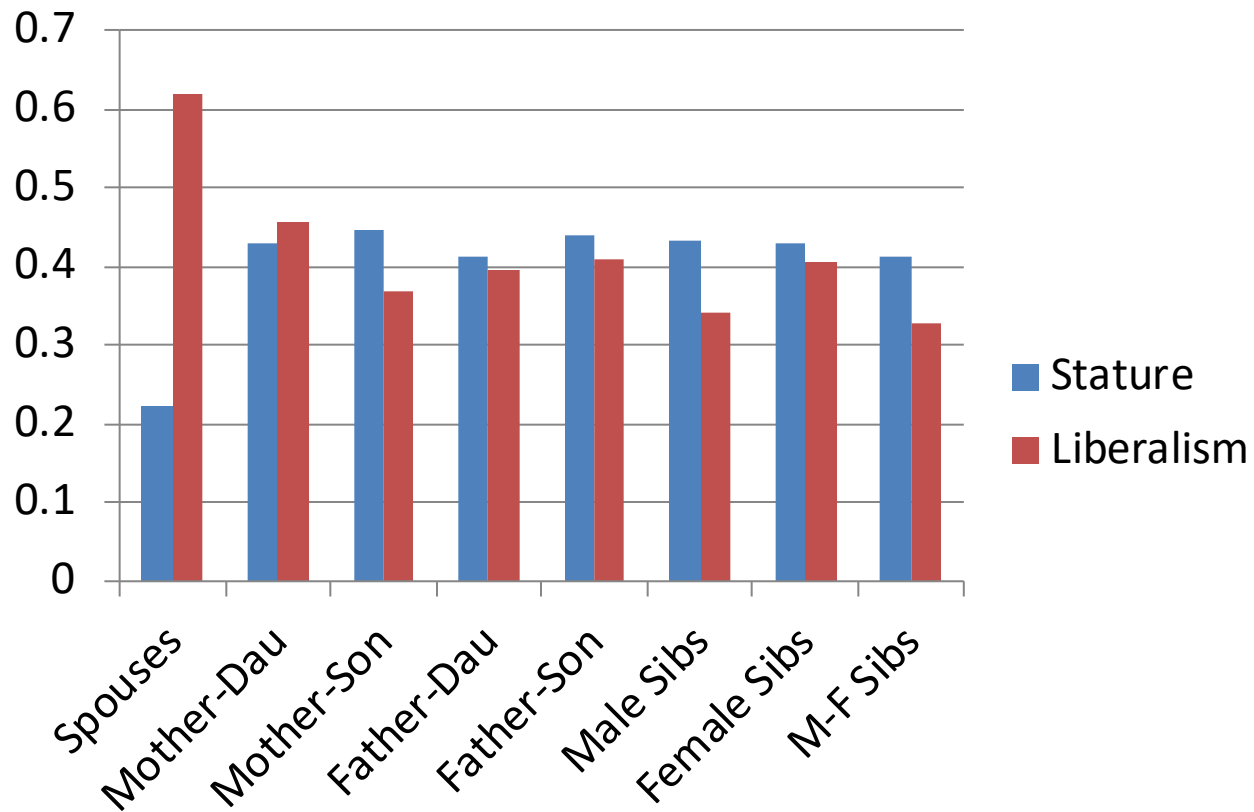
Overall sample sizes

Relationship	# of pairs
Parent-offspring	25018
Siblings	18697
Spouses	8287
DZ Twins	5120
MZ Twins	4623

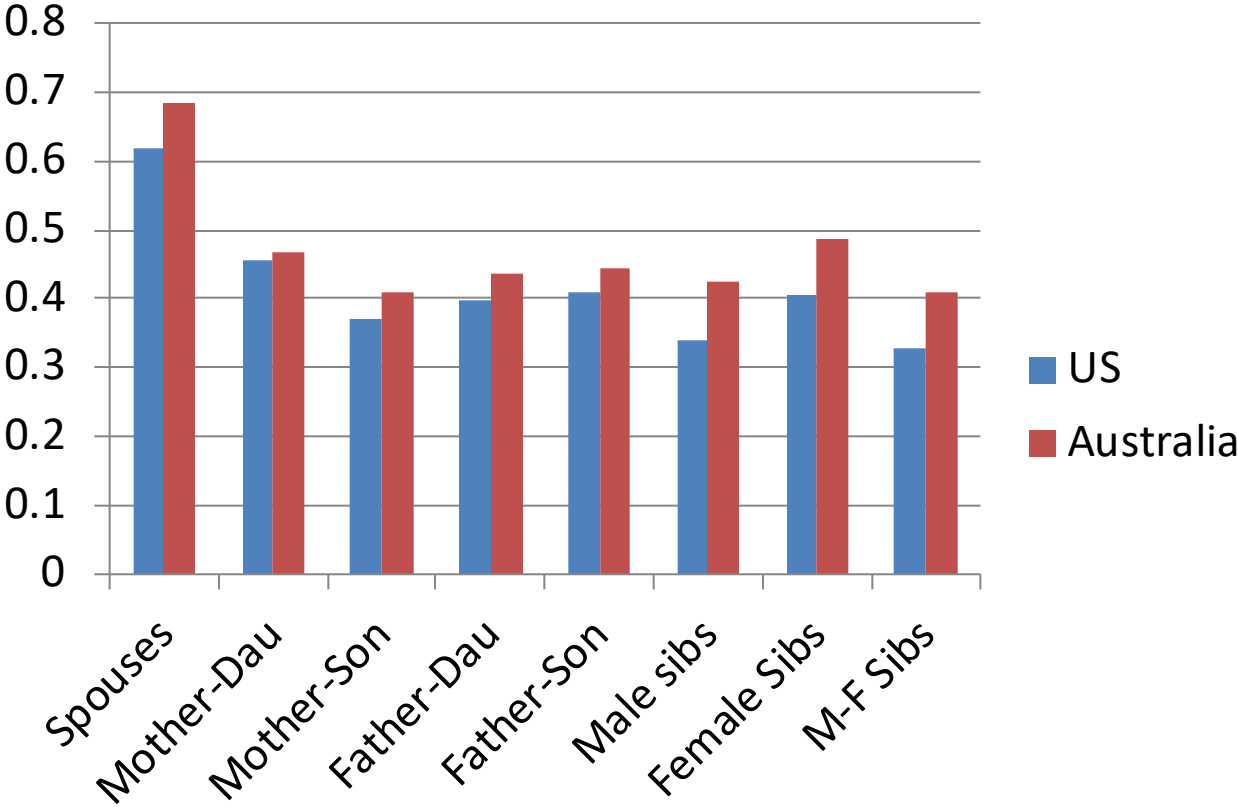
Nuclear Family Correlations for Stature (Virginia 30,000 and OZ 22,000)



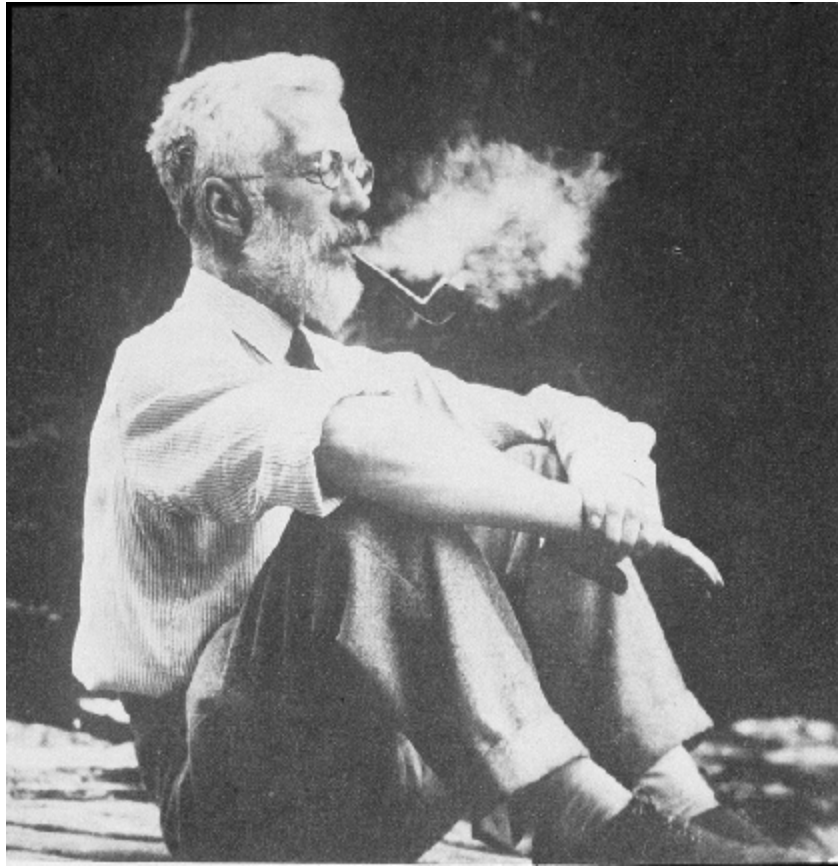
Nuclear Family Correlations for Stature and Liberalism/Conservatism (Virginia 30,000)



Nuclear Family Correlations for Liberalism/Conservatism (Virginia 30,000 and Australia 22,000)



Ronald Fisher (1890-1962)



1918: On the Correlation Between Relatives on the Supposition of Mendelian Inheritance

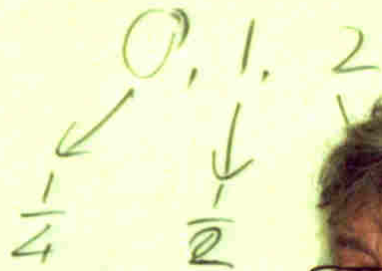
1921: Introduced concept of “likelihood”

1930: The Genetical Theory of Natural Selection

1935: The Design of Experiments

Fisher developed mathematical theory
that reconciled Mendel's work with
Galton and Pearson's correlations

Heb's l.b.d.

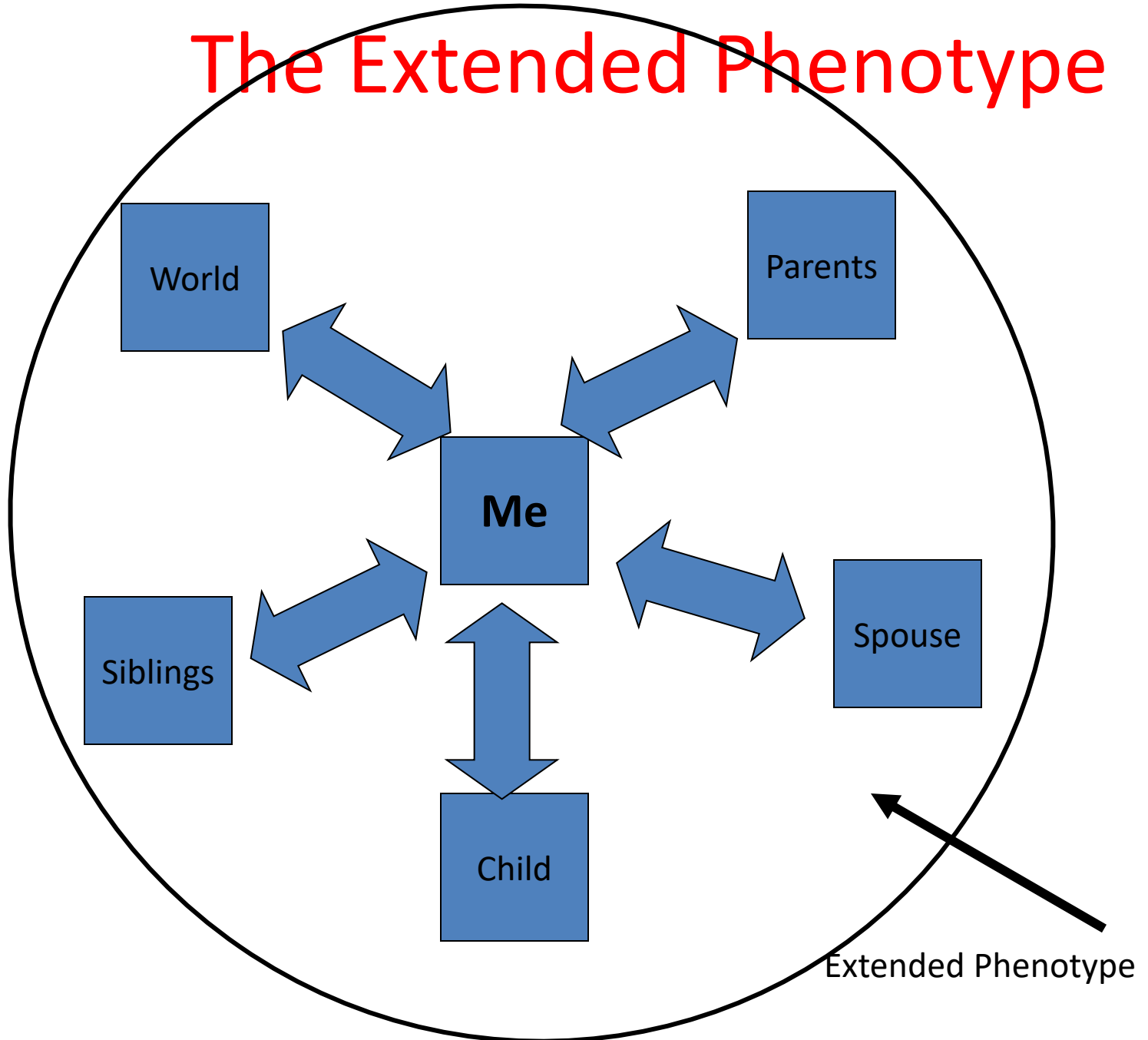


$$\mu - r_0 = \psi$$

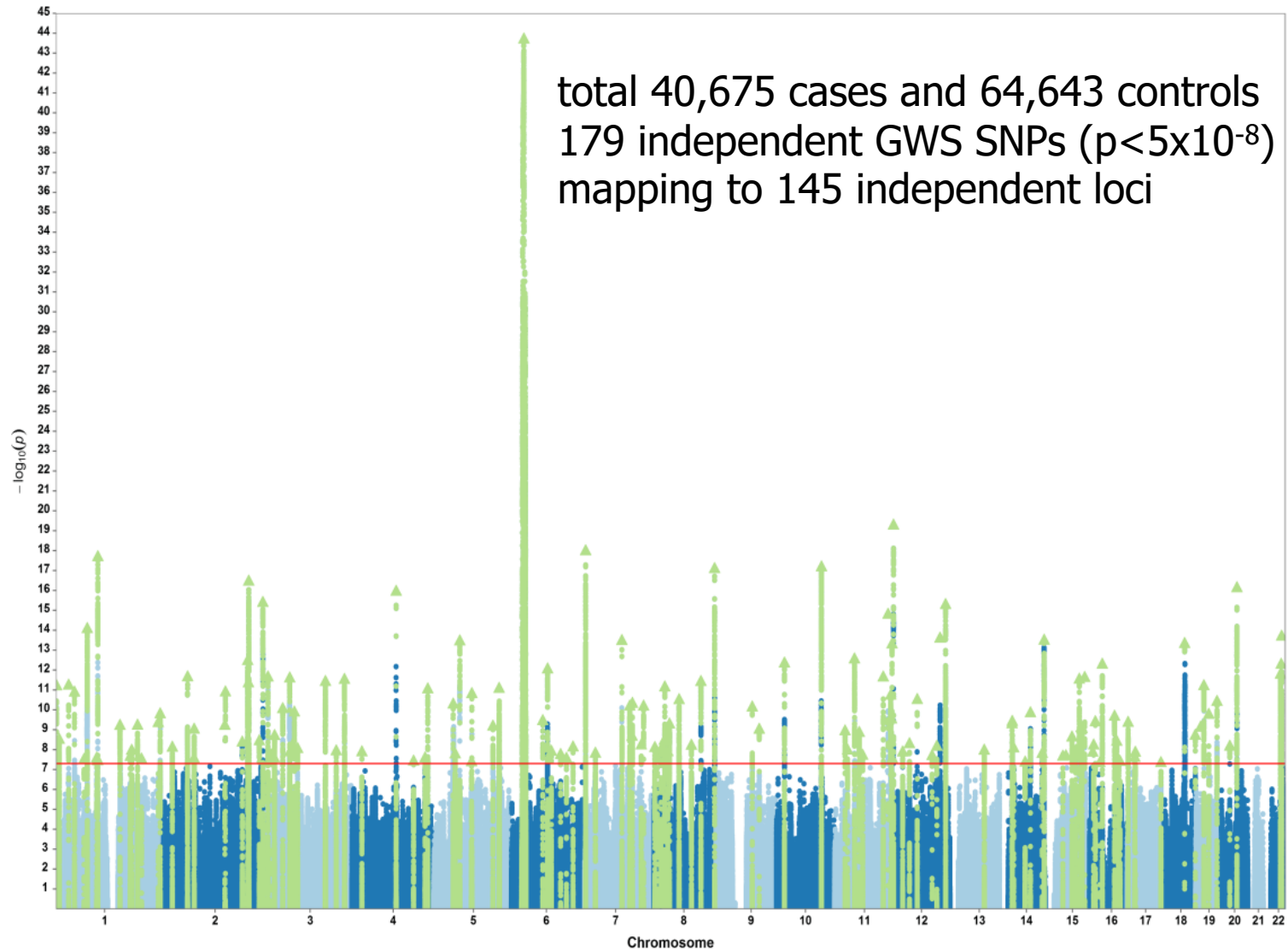
$$d = r_2 - r_0 =$$

Lindon Eaves

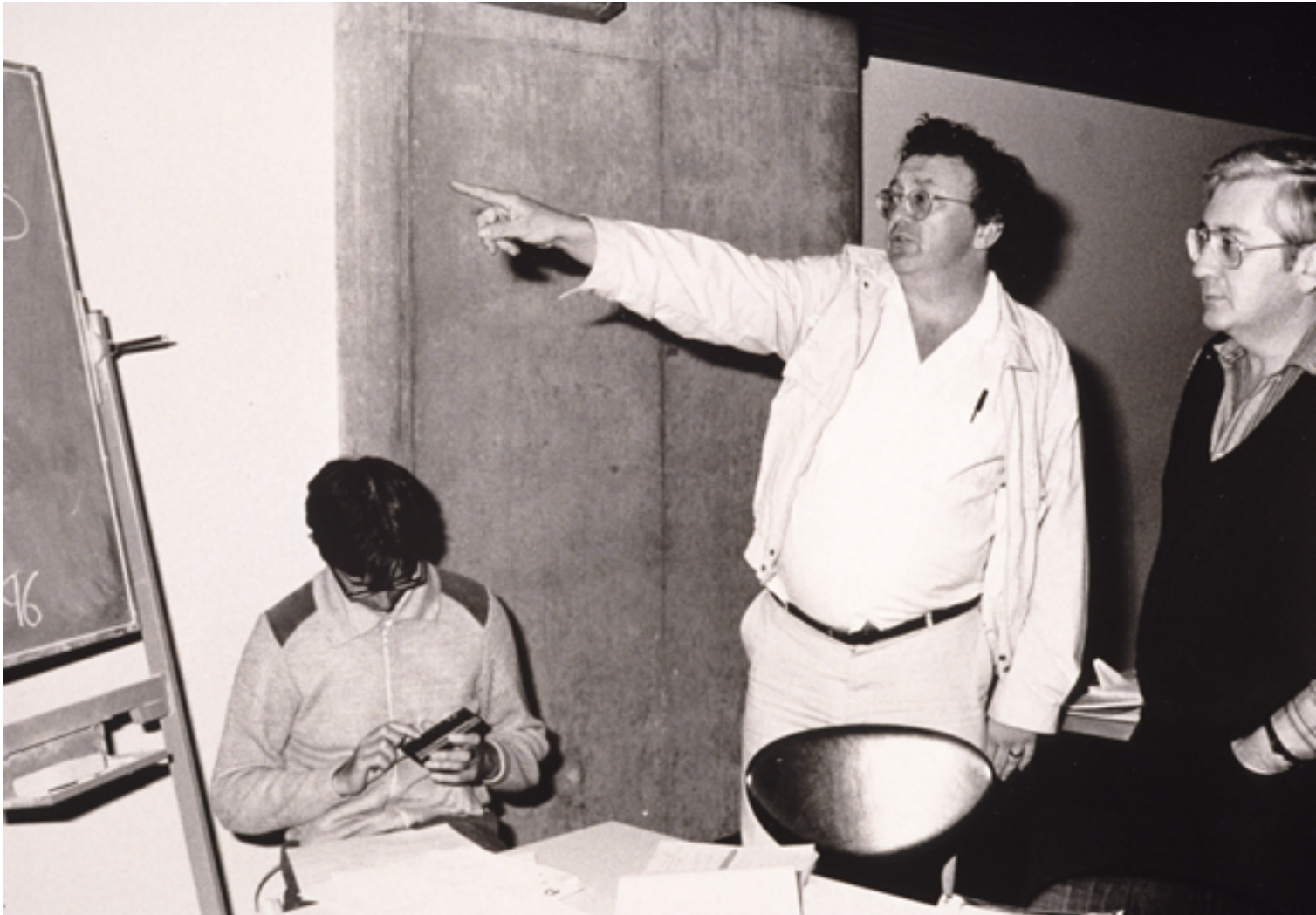
The Extended Phenotype



Schizophrenia GWAS associations from the meta-analysis of CLOZUK and PGC



Lindon Eaves



Andrew Heath

David Fulker

Polygenic Risk Scores for Prediction of Breast Cancer and Breast Cancer Subtypes

Nasim Mavaddat,^{1,*} Kyriaki Michailidou,^{1,2} Joe Dennis,¹ Michael Lush,¹ Laura Fachal,³ Andrew Lee,¹

Compared with women in the middle quintile, those in the highest 1% of risk had 4.37- and 2.78-fold risks, and those in the lowest 1% of risk had 0.16- and 0.27-fold risks, of developing ER-positive and ER-negative disease, respectively. This PRS is a powerful and reliable predictor of breast cancer risk that may improve breast cancer prevention programs.

