# Model assumptions & extending the twin model

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## Acknowledgments

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Files you will need are in Faculty drive: /matt/Assumptions\_2012

- Assumptions\_mck.pdf (the powerpoint presentation)
- CTD.ACDE-param.indet\_2012.R (OpenMx script)

#### Also: 4 PDF papers describing details of what we go over here:

- Keller\_Coventry\_CTD\_Indeterminacy\_2005: why parameter indeterminacy exists
- Coventry\_Keller\_CTD.ETFD.Comparison\_2005 comparison of extended twin model estimates and CTD estimates across ~ 30 phenotypes
- Keller\_Cascade\_2009 step-by-step explanation of how extended twin models work
- Keller\_ETFD\_2010 simulation work showing the effects of violations of assumptions on 4 different thopes of models

## Structural Equation Modeling (SEM) in BG

#### SEM is great because...

- Directs focus to effect sizes, not "significance"
- Forces consideration of causes and consequences
- Explicit disclosure of assumptions
- Potential weakness...
  - Parameter reification: "Using the CTD we found that 50% of variation is due to A and 20% to C."
  - Should you believe that 50% of variation is truly additive genetic?

#### True parameters vs. Estimated parameters

A C D E: true (unknowable) values of A, C, D, E in the population (short for VA, VC, VD, and VE)

 $\hat{A} \ \hat{C} \ \hat{D} \ \hat{E}$  : \*estimated\* values of A, C, D, E.

Ĉ D Ê will differ from A, C, D, E due to:
1) sampling variability
2) bias

#### Quiz Question 1

I)  $\hat{C}$  and  $\hat{D}$  cannot be estimated simultaneously in the classical twin design (i.e., the design that uses MZ and DZ twins only) model because: [choose all that apply]

a) these two estimates are too highly correlated (multicolinearity problems)

- b) they can be estimated simultaneously; you just have to fix  $\hat{A}$  to some specific value
- c) there are more informative statistics than parameters to be estimated
- d) there are fewer informative statistics than parameters to be estimated



Why can't we estimate  $\hat{C} \And \hat{D}$  at same time using twins only?

Solve the following two equations for  $\hat{A}$ ,  $\hat{D}$ ,  $\hat{C}$ : CVmz = A + D + CCVdz = I/2A + I/4D + C

- 3 unknowns, 2 informative equations. It can't be done. The model is "unidentified".
- In practice, you can detect non-identification by noting that (a) model estimates depend on starting values AND (b) all final models have identical likelihoods

#### Indeterminacy: Practical 1

- Open up CTD.ACDE-param.indet\_2012.R in R
- Run this script until you see "# END PRACTICAL I." Don't close the script or R, as we'll use this same script again for Practicals 2, 3, & 4.
- Write down your -2 log likelihood and your estimates of A, C, and D
- Compare these to your neighbor's results
- WHY is this occurring?

## The CTD: Two statistics give info about within-family resemblance



#### ACE Model



- Assume D = 0. Solve for  $\hat{A} \& \hat{C}$ CVmz = A + C CVdz = I/2 A + C
- 2 unknowns, 2 independently informative equations:  $\hat{A} = 2(CVmz-CVdz)$  $\hat{C} = 2CVdz-CVmz$

Note: if we tried to estimate  $\hat{D}$ , it would necessarily hit the 0 boundary anyway and the model wouldn't fit as well (because  $\hat{D}$  'wants' to go negative), so it makes sense to solve for  $\hat{C}$ 

#### The CTD: ADE Model



#### PRACTICAL 2: ADE Algebra & Indeterminacy

- Assume C = 0. Solve for  $\hat{A} \& \hat{D}$ CVmz = A + D CVdz =  $\frac{1}{2}A + \frac{1}{4}D$ Solve for these yourselves algebraically now
- Then reopen CTD.ACDE-param.indet.R in R & run FROM "# START PRACTICAL 2" TO "# END PRACTICAL 2"
- Did you get roughly the same answer as your algebra suggested?

What happened to estimates of C & D in the DCE model?

**Review Question 1 again** 

I)  $\hat{C}$  and  $\hat{D}$  cannot be estimated simultaneously in the classical twin design (i.e., the design that uses MZ and DZ twins only) model because: [choose all that apply]

a) these two estimates are too highly correlated (multicolinearity problems)

b) they can be estimated simultaneously; you just have to fix  $\hat{A}$  to some specific value

c) there are more informative statistics than parameters to be estimated

d) there are fewer informative statistics than parameters to be estimated

**Quiz Question 2** 

2) What are the *typical* assumptions of a classical twin model? [choose all that apply]

- a) the MZ and DZ covariances are equal
- b) either D or C is equal to zero

c) only genetic factors cause MZ twins to be more similar to each other than DZ twins

- d) no assortative mating
- e) no epistasis
- f) no gene-environment interactions or correlations

What are the effects of violations of assumptions in the CTD?

Either D or C is equal to zero: A is overestimated and D and C are underestimated

Only genetic factors cause MZ twins to be more similar to each other than DZ twins: A and D are overestimated and C is underestimated

No assortative mating: A and D are underestimated and C is overestimated

No epistasis: D or A is overestimated and C is underestimated

No gene-environment interactions or correlations: AxC: A overestimated; AxE: E overestimated; Cov(A,C): depends

#### Bias in parameter estimates

- In ACE Models:  $\hat{A} = A + 3/2D$  $\hat{C} = C - \frac{1}{2}D$
- In ADE Models:
   Â = A + 3C
   D = D 2C

3) If the assumptions of a model are violated (i.e., not true in the real world)... [choose all that apply]

a) the interpretation of the estimated parameters should be altered; e.g.,  $\hat{A}$  should be considered an amalgam of A, D, and C and not just additive genetic effects

b) there is no point in doing the analysis at all

c) the point estimates of the estimated parameters may be biased

#### **Quiz Question 4**

4) An ACE model finds that  $\hat{A} = .30$  and  $\hat{C} = .15$ . This implies that D does not influence the trait in question, or has a minor (non-significant) effect.

a) TRUE b) FALSE 5) We run an ADE model and find that  $\hat{A} = .69$  and that  $\hat{D} = .05$ . If in truth, C = .10, what will the effect on the estimated parameters be? [choose all that apply]

- a)  $\hat{D}$  will be biased (too low)
- b)  $\hat{D}$  will be biased (too high)
- c)  $\hat{A}$  will be biased (too low)
- d)  $\hat{A}$  will be biased (too high)

e) there is no affect on the estimated parameters; however by not estimating C (aka, fixing it to zero), we underestimated C

#### PRACTICAL 3: Sensitivity analysis

- Sensitivity analysis: studying what the effects are on estimated parameters when assumptions are wrong
- In CTD.ACDE-param.indet.R, run: FROM "# START PRACTICAL 3" TO "# END PRACTICAL 3"
- Change the value of C from 0 to other values (remember, C=c^2). What happens to estimates of A and D depending on different assumed values of C?

6) In the CTD, we have two different relative covariance estimates (MZ covariance & DZ covariance). Let's say we add parents to the twin design. That gives us 2 additional relative covariance estimate to work with (parentoffspring and spousal) and allows us to [choose all that apply]

- a) estimate A, C, & D simultaneously
- b) account for the effects of assortative mating
- c) account for passive G-E covariance
- d) reduce the bias in estimates of A, C, and D vis a vis the CTD

## Classical Twin Design (CTD)

Assumption	biased up	biased down
Either D or C is zero	Α	C & D
No assortative mating	С	D
No A-C covariance	$\mathbf{C}$	D & A



# Adding parents gets us around all these assumptions



### Parents also allow differentiation of S & F



F = familial env factors passed from parents to offspring



## Nuclear Twin Family Design (NTFD)



#### **PRACTICAL 4: NTFD analysis**

- In CTD.ACDE-param.indet.R, run: FROM "# START PRACTICAL 4" TO "# END PRACTICAL 4"
- What are the estimated values of A, D, & S? [Note: S = sib environment, equivalent to C in the CTD]

#### CTD vs. NTFD vs. simulation results

TRUE values CTD estimates NTFD estimates
 A = .30 A-hat = .68 A-hat = .32
 D = .30 D-hat = .04 D-hat = .29
 S = .10 S-hat = 0 S-hat = .13

## Nuclear Twin Family Design (NTFD)



#### Assumptions:

- Only can estimate 3 of 4: A, D, S, and F (bias is variable)
- Assortative mating due to primary phenotypic assortment (bias is variable)

Include twins and their sibs, parents, spouses, and offspring...

- Gives 17 unique covariances (MZ, DZ, Sib, P-O, Spousal, MZ avunc, DZ avunc, MZ cous, DZ cous, GP-GO, and 7 inlaws)
- 88 covariances with sex effects

## Additional obs. covs with Stealth allow estimation of A, S, D, F, T



= env. factors shared only between twins



(Remember: we're not just estimating more effects. More importantly, we're

reducing the bias in estimated effects!)



Assumption	biased up	biased down
Primary assortative mating	A, D, or F	A, D, or F
No epistasis	A, D	S
No AxAge	D, S	Α



- Primary AM: mates choose each other based on phenotypic similarity
- Social homogamy: mates choose each other due to environmental similarity (e.g., religion)
  - Convergence: mates become more similar to each other (e.g., becoming more conservative when dating a conservative)

### Cascade





GeneEvolve65.zip

Note: GeneEvolve is still a 'beta-version'. Breakdowns are likely! You can help by

Reality: A=.5, D=.2







Reality: A=.4, D=.15, S=.15





#### A,D, & F estimates are highly correlated in Stealth & Cascade





Reality: A=.4, A\*A=.15, S=.15



#### Reality: A=.4, A\*Age=.15, S=.15



## Conclusions

- All models require assumptions. Generally, more assumptions = more biased estimates
- Simulations provide independent assessments of the NTFD, *Stealth*, and *Cascade* models
  - These complicated models work as designed!
  - In all models, but especially the CTD, be cautious of reifying parameter estimates!
  - A is amalgam of mostly A but also D & C. It is a decent estimate of broad sense h2.
- D & C are likely to be underestimates

#### Stealth application

Twin Research (1999) 2, 99-107 © 1999 Stockton Press All rights reserved 1369-0523/99 \$12.00

http://www.stockton-press.co.uk/tr

#### Frequency of church attendance in Australia and the United States: models of family resemblance

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Data on frequency of church attendance have been obtained from separate cohorts of twins and their families from the USA and Australia (29063 and 20714 individuals from 5670 and 5615 families, respectively). The United States sample displayed considerably higher frequency of attendance at church services. Sources of family resemblance for this trait also differed between the Australian and US data, but both indicated significant additive genetic and shared environment effects on church attendance, with minor contributions from twin environment, assortative mating and parent-offspring environmental transmission. Principal differences between the populations were in greater maternal environmental effects in the US sample, as opposed to paternal effects in the Australian sample, and smaller shared environment effects observed for both women and men in the US cohort.

Keywords: religion, church attendance, extended kinship model, twins, cultural inheritance, assortative mating, twin environment

#### Further reading on this lecture

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