Ordinal data Analysis: Liability Threshold Models

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Ordinal data

- Measuring instrument discriminates between two or a few <u>ordered</u> categories e.g.:
 - Absence (0) or presence (1) of a disorder
 - Score on a single Q item e.g. : 0 1, 0 4
- In such cases the data take the form of counts, i.e. the number of individuals within each category of response

Analysis of categorical (ordinal) variables

- The session aims to show how we can estimate correlations from simple count data (with the ultimate goal to estimate h², c², e²)
- For this we need to introduce the concept of 'Liability' or 'liability threshold models'
- This is followed by a more mathematical description of the model

Liability

Liability is a **theoretical** construct. It's the assumption we make about the distribution of a variable which we were only able to measure in terms of a few ordered categories

Assumptions:

(1) Categories reflect an imprecise measurement of an underlying *normal distribution* of liability

(2)The liability distribution has 1 or more thresholds (cut-offs) to discriminate between the categories

For disorders:



The *risk* or liability to a disorder is normally distributed, only when a certain threshold is exceeded will someone have the disorder. Prevalence: proportion of affected individuals.

For a single questionnaire item score e.g:



Does not make sense to talk about prevalence: we simply count the endorsements of each response category

The Standard Normal Distribution

Liability is a *latent* variable, the scale is arbitrary, distribution is assumed to be a *Standard Normal Distribution* (SND) or z-distribution:

- Mathematically described by the SN Probability Density function ($\Phi = phi$), a bell-shaped curve with:
 - mean = 0 and SD = 1
 - z-values are the number of SD away from the mean
- Convenience: area under curve =1, translates directly to probabilities



Standard Normal Cumulative Probability in right-hand tail (For negative z values, areas are found by symmetry)

Area = $P(z \ge z_T)$



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$$\int_{Z_T}^{\infty} \Phi(L_1; \mu = 0, \sigma^2 = 1) dL_1$$

Standard Normal Cumulative Probability in right-hand tail (For negative z values, areas are found by symmetry)

Area	$Area = P(z \ge z_T)$				
.50	50%				
.42	42%				
.35	35%				
.27	27%				
.21	21%				
.16	16%				
.12	12%				
.08	8%	2 Z_{T} 3			
.06	6%	-3 0 3			
.036	3.6%				
.023	2.3%				
.014	1.4%				
800.	.8%	$\Phi(L_1; \mu = 0, \sigma^2 = 1) \alpha$			
.005	.5%				
.003	.3%	Z_T			
.002	.2%				
	Area .50 .42 .35 .27 .21 .16 .12 .08 .06 .036 .023 .014 .008 .005 .003 .003 .002	Area .50 50% .42 42% .35 35% .27 27% .21 21% .16 16% .12 12% .08 8% .06 6% .036 3.6% .023 2.3% .014 1.4% .008 .8% .005 .5% .003 .3% .002 .2%			

We can find the area between any two thresholds

$Area = P(.6 \le z \le 1.8)$



 Z_0 Area to the right.6.27(27 %)1.8.036(3.6 %) -27-3.6 = 23.4 %

Ability to work out the areas under the curve (proportions) enables the reverse operation, e.g. find the z-value to describe proportion of affected individuals in a sample or proportion scoring e.g 0, 1, 2 on item.

From sample counts to z-value

It is possible to find a **z-value** (threshold) so that the area exactly matches the observed proportion of the sample e.g. sample of 1000 individuals, where 80 have met the criteria for a disorder (8%): the z-value is 1.4

Z ₀	Area					
.6	.27	27%				
.8	.21	21%		1		
1	.16	16%		-		
1.2	.12	12%				
1.4	.08	8%				
1.6	.055	6%		1 /		
1.8	.036	3.6%				
2	.023	2.3%		-3	0 1.4	3
2.2	.014	1.4%				
2.4	.008	.8%		/		
2.6	.005	.5%		unaff		aff
2.8	.003	.3%		unan		an
2.9	.002	.2%	Counts:	920		80

Two ordinal traits: Data from twins

> Contingency Table with 4 observed cells:

cell a: pairs concordant for unaffected cell d: pairs concordant for affected cell b/c: pairs discordant for the disorder

Twin1		
Twin2	0	1
0	۵	b
1	C	d

0 = unaffected 1 = affected

Joint Liability Model for twin pairs

- Assumed to follow a bivariate normal distribution, where both traits have a mean of 0 and standard deviation of 1, but the correlation between them is variable.
- The **shape** of a bivariate normal distribution is determined by the **correlation** between the traits



r =.90



Bivariate Normal (R=0.6) partitioned at threshold 1.4 (z-value) on both liabilities









Expected Proportions of the BN, for R=0.6, Th1=1.4, Th2=1.4









How are expected proportions calculated?

By **numerical integration** of the bivariate normal over two dimensions: the liabilities for twin1 and twin2

e.g. the probability that both twins are affected :

$$\int_{T_{1}T_{2}}^{\infty} \Phi(L_{1}, L_{2}; \mu = 0, \Sigma) dL_{1} dL_{2}$$



Ability to work out the expected proportions given any correlation (shape of the BND) and set of thresholds on the liabilities, enables the reverse operation i.e. the sample proportions in the 4 cells of the CT (i.e. number of 00, 01,10 and 11 scoring pairs) are used to estimate the correlation between liabilities and the thresholds

Twin Models

- Estimate correlation in liabilities separately for MZ and DZ pairs from their Count data
- Variance decomposition (A, C, E) can be applied to the *liability* of the trait
- Correlations in liability are determined by path model
- Estimate of the heritability of the *liability*

Summary

- To estimate correlations for ordinal traits (counts) we make assumptions about the joint distribution of the data (Bivariate Normal)
- The relative proportions of observations in the cells of the Contingency Table are translated into proportions under the BN
- The most likely thresholds and correlations are estimated
- Genetic/Environmental variance components are estimated based on these correlations derived from MZ and DZ data

ACE Liability Model



Test of BN assumption

For a 2x2 CT with 1 estimated TH on each liability, the χ^2 statistic is always 0, 3 observed statistics, 3 param, df=0 (it is always possible to find a correlation and 2 TH to perfectly explain the proportions in each cell). No goodness of fit of the normal distribution assumption.

This problem is resolved if the CT is at least 2x3 (i.e. more than 2 categories on at least one liability) A significant χ^2 reflects departure from normality.



Fit function Raw Ordinal Data

- The likelihood for a vector of observed ordinal responses is computed by the expected proportion in the corresponding cell of the MV distribution
- The likelihood of the model is the sum of the likelihoods of all vectors of observation
- This is a value that depends on the number of observations and isn't very interpretable (as with continuous raw data analysis)
- So we compare it with the LL of other models, or a saturated (correlation) model to get a χ^2 model-fit index

(Equations given in Mx manual, pg 89-90)

Raw Ordinal Data

	ordinal	ordinal
Zyg	respons1	respons2
1	0	0
1	0	0
1	0	1
2	1	0
2	0	0
1	1	1
2	•	1
2	0	•
2	0	1

Power issues

 Ordinal data / Liability Threshold Model: less power than analyses on continuous data

Neale, Eaves & Kendler 1994

- Solutions:
- 1. Bigger samples
- 2. Use more categories



Model-fitting to Raw Ordinal Data

Practical

Sample and Measures

- TEDS data collected at age 8
- Parent report
- Childhood Asperger Syndrome Test (CAST) (Scott et al., 2002)
- twin pairs: 1221 MZ 2198 DZ
- Includes children with autism spectrum conditions



The CAST score dichotomized at 98% (i.e. Scores of >16), is the official cut-off point for children at risk for <u>Autism Spectrum Disorder</u>

This resulted in only 16 concordant affected pairs (0 in some groups).

Numbers improved using a cut-off point of 90% (however, clinically less interesting)

Practical Exercise

CAST score dichotomized (0,1) at 90% > threshold (z-value) of around 1.28 Prevalence of boys (14%)

Observed counts: MZM 0 1

0 483 170 435 531 29 441 54 29

DZM

1

File:	cast90m.dat
R Script:	UnivACE_MatrRawOrd.R
Dir:	fruhling/Ordinal Analyses/Binary

Cast90m.dat



```
# Program: UnivACE_MatrRawOrd.R
require(OpenMx)
source("GenEpiHelperFunctions.R")
 #
# Read data from REC ASCI text file (cast90m.dat) with '.' as missing values, space sep
# Pariabels: zyg cast90_tw1 cast90_tw2
# żyg: 1=mz, 2=dz (all males)
al[Vars<- c('zyg', 'cast90_tw1', 'cast90_tw2')
Castdata <- read.table ('cast90m.dat', header=F, sep="", na.strings=".", col.names=allVars)
nv<sub>m</sub><- 1
ntv <- nv^{*}2
summary(Castdata)
str(Castdata)
Vars < -('cast90')
selVars <- c('cast90_tw1', 'cast90_tw2')
mzData <- subset(Castdata, zyg==1, selVars)
dzData <- subset(Castdata, zyg==2, selVars)
 Е
# Print Descriptive Statistics
summary(mzData)
summary(dzData)
table(mzData$cast90_tw1, mzData$cast90_tw2)
table(dzData$cast90_tw1, dzData$cast90_tw2)
 W
```

 \mathbf{O}

Specify and Run Saturated Model (Tetrachoric correlations)

```
twinSatModel <- mxModel("twinSat", mxModel("MZ",
```

Matrix & Algebra for expected means, Thresholds and correlation mxMatrix(type="Zero", nrow=1, ncol=nv, name="M"), mxAlgebra(expression= cbind(M,M), name="expMean"),

mxMatrix(type="Full", nrow=1, ncol=ntv, free=TRUE, values=.8, name="expThreMZ", dimnames=list('th1',selVars)),

mxMatrix(type="Stand", nrow=2, ncol=2, free=T, values=.5, lbound=-.99, ubound=.99, name="expCorMZ"),

mxData(mzData, type="raw"), mxFIMLObjective(covariance="expCorMZ", means="expMean", thresholds="expThreMZ", dimnames=selVars,)), # Specify and Run Saturated Model (Tetrachoric correlations) twinSatModel <- mxModel("twinSat", mxModel("MZ", # Matrix & Algebra for expected means, Thresholds and correlation mxMatrix(type="Zero", nrow=1, ncol=nv, name="M"), $\left(0,0\right)$ mxAlgebra(expression = cbind(M,M), name="expMean"), mxMatrix(type="Full", nrow=1, ncol=ntv, free=TRUE, values=.8, name="expThreMZ", dimnames=list('th1',selVars)), cast90 tw1 cast90 tw2 th1 ThLiab1, ThLiab2

mxMatrix(type="Stand", nrow=2, ncol=2, free=T, values=.5, lbound=-.99, ubound=.99, name="expCorMZ"),





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mxMatrix(type="Stand", nrow=2, ncol=2, free=T, values=.5, lbound=-.99, ubound=.99, name="expCorMZ"),

mxData(mzData, type="raw"), mxFIMLObjective(covariance="expCorMZ", means="expMean", thresholds="expThreMZ", dimnames=selVars,)),

```
mxModel("DZ",
# Matrix & Algebra for expected means, Thresholds and correlation
.
.
```

, twinSatFit <- mxRun(twinSatModel) twinSatSumm <- summary(twinSatFit) twinSatSumm # Specify and Run Saturated SubModel 1 equating Thresholds across Twin 1
and Twin 2 within zyg group

twinSatSub1 <- twinSatModel</pre>

twinSatSub1\$MZ\$expThreMZ <- mxMatrix(type="Full", nrow=1, ncol=2, free=T, 0.8, label="threMZ", name="expThreMZ", dimnames=list('th1',selVars))

twinSatSub1\$DZ\$expThreDZ <- mxMatrix(type="Full", nrow=1, ncol=2, free=T, 0.8, label="threDZ", name="expThreDZ", dimnames=list('th1',selVars))

twinSatSub1Fit <- mxRun(twinSatSub1)
twinSatSub1Summ <- summary(twinSatSub1Fit)
twinSatSub1Summ</pre>

cast90_tw1 cast90_tw2 th1 threMZ, threMZ # Fit ACE Model with RawData and Matrices Input, ONE overall Threshold

univACEOrdModel <- mxModel("univACEOrd", mxModel("ACE", # Matrices a, c, and e to store a, c, and e path coefficients mxMatrix(type="Full", nrow=nv, ncol=nv, free=TRUE, values=.6, label="a11", name="a"), mxMatrix(type="Full", nrow=nv, ncol=nv, free=TRUE, values=.6, label="c11", name="c"), mxMatrix(type="Full", nrow=nv, ncol=nv, free=TRUE, values=.6, label="e11", name="e"), # Matrices A, C, and E compute variance components F C Α mxAlgebra(expression=a %*% t(a), name="A"), mxAlgebra(expression=c %*% t(c), name="C"), mxAlgebra(expression=e %*% t(e), name="E"), # Algebra to compute total variances and SD mxAlgebra(expression=A+C+E, name="V"), mxMatrix(type="Iden", nrow=nv, ncol=nv, name="I"), mxAlgebra(expression=solve(sqrt(I*V)), name="sd"),# Constraint on variance of ordinal variables mxConstraint(alg1="V", "=", alg2="I", name="Var1"), A + C + E = '

Illustration

Run script and check that the values in the Table are correct.

What are the conclusions about the thresholds?

What is the final model in terms of the thresholds?

MODEL	np	-2LL	df	$\Delta \chi^2(\mathbf{d} \mathbf{f})$	sig
1 All TH free ^{&}	6	1599.8	2282	-	-
2 TH tw1=tw2 in MZ and DZ \$	4	1602.9	2284	3.18 (2)	.20 ns
3 One TH for all males	3	1605.6	2285	5.85 (3)	.12 ns

[&] Thresholds: MZM twin 1 =1.14, MZM twin 2 = 1.25 DZM twin 1 = 1.06, DZM twin 2 = 1.06

^{\$} Thresholds: MZM =1.19, DZM = 1.06

Based on these results, the final TH model in the script is: 1TH for males: 1.11

The correlations for this model are: r MZM = 0.87 (.80-.93) r DZM = 0.45 (.29-.59)



- Add the 'CE'sub-model, using the same logic as for the 'AE' sub-model
- Note: In
- **# Print Comparative Fit Statistics**

univOrdACENested <- list(univOrdAEFit, univOrdCEFit, univOrdEFit) tableFitStatistics(univOrdACEFit,univOrdACENested)

DF and Constraints OS 2288

Δ

ACE Model param

NP
BeforeConstraintNumber
Of Constr
1NP
AfterConstrainta, c, e (3)
thresholds (1)12

3

df OS - NP_{AC} = 2288 - 3 = 2285 OpenMx: OS + number of Constr - NP_{BC} = 2289 - 4 = 2285

Model	-2LL	df	np _{BC}	np _{AC}	Model of comp	$\Delta \chi^2 (\Delta df)$	sig
ACE	1605.6	2285	4*	3	-	-	-
CE	1633.6	2286	3	2	ACE	27.9 (1)	p=<.001
AE	1605.7	2286	3	2	ACE	0.02 (1)	p=.89
E	1774	2287	2	1	ACE	168 (2)	p=<.001

* A, C, E + 1 Threshold

Estimates

	h ²	c ²	e ²
ACE	.85	.02	.13
AE	.88		.12

Multiple Thresholds: more than two categories

For multiple threshold models, to ensure t1>t2>t3 etc.....

We use a slightly more complicated model for the thresholds

Threshold Specification



3 Categories > 2 thresholds per liability Matrix T: 2 x 2 T(1,1) T(1,2) threshold 1 for twin1 & twin2 T(2,1) T(2,2) increment



Use multiplication to ensure that second threshold is higher than first

Expected Thresholds: _____ L³

 $\begin{pmatrix} 1 & 0 \\ 1 & 1 \end{pmatrix} * \begin{pmatrix} t_{11} & t_{12} \\ t_{21} & t_{22} \end{pmatrix} = \begin{array}{c} t_{11} & t_{12} \\ t_{11} + t_{21} & t_{12} + t_{22} \end{array}$

 $T_{12} = t_{12}$ $T_{22} = t_{12} + t_{22}$

 $T_{11} = t_{11} \quad T_{21} = t_{11} + t_{21}$

Thresholds twin 1Thresholds twin 2T11T12T21T22

nth <- 2 thRows <- paste("th",1:nth,sep="") # thRows <- c('th1','th2')

number of thresholds

mxMatrix(type="Full", nrow=nth, ncol=ntv, free=TRUE, values=.5, lbound = c(-3, 0.0001, -3, 0.0001), name = "Thmz"),

mxMatrix(type="Lower", nrow \neq nth, ncol=nth, free=FALSE, values=1, name="Inc"), mxAlgebra(expression= Inc %*% Thmz, dimnames=list(thRows,selVars), name="expT/nmz"),

 $t_{11} + t_{21} + t_{12} + t$ explhmz

Note

- This script will work if all variables have all ordered categories in the right order: e.g. 1 2 3 4 or 0 1 2 3
- If that is not true, e.g. you have a variable with possible categories 1-6, but no one has scored 3 and 6
- Or if you have a categorical variable with 4 possible scores 1 4 9 16, openMx will treat it as continuous.
- This can all be done internally